

COMPARISONS AMONG COMBINATIONS OF PROGENY
TEST DESIGNS AND SELECTION SYSTEMS FOR
MAXIMUM GENETIC GAIN IN LOBLOLLY
PINE (PINUS TAEDA L.)

By

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INTRODUCTION

Commercial forest land in southeastern Oklahoma and southwestern Arkansas is often highly heterogeneous. Progeny tests are designed to evaluate the performance of offspring of selected trees to aid in the choice of parents for the next generation of commercial seedlings. If the progeny tests are to provide a realistic estimate of the performance to be expected of commercial seedlings, they must be established on a variety of commercial site types. Thus, progeny test sites often contain a high degree of variability.

Progeny tests usually serve the dual function of providing information about the parent trees as well as sources from which the next generation of parents are selected. The selection process typically involves both family mean and individual tree performance. On a variable site, both the progeny test design and the method of within-family selection may be very important.

Several approaches have been pursued in this study to compare the statistical efficiencies of various designs, to compare the relative genetic gains obtained by using combinations of several designs and selection methods and to quantify the environmental relationships among the positions of trees in a plantation.

Three separate articles constitute Parts I, II and III. The first, "Statistical Efficiency of Six Progeny Test Designs on Three Loblolly Pine (Pinus taeda L.) Site Types", will be submitted for publication in the Canadian Journal of Forest Research. The second article, "Computer Simulations to Compare Combinations of Progeny Test Designs and Selection Systems for Maximum Genetic Gain in Loblolly Pine (Pinus taeda L.)", will be presented at a meeting of the International Union of Forest Research Organizations in Williamsburg, VA, in October, 1986. It will subsequently be published in the IUFRO proceedings. The third article, "Neighborhood Comparison as a Basis for Selection in Loblolly Pine (Pinus taeda L.)", will be submitted for publication in Theoretical and Applied Genetics.

PART I

STATISTICAL EFFICIENCY OF SIX PROGENY

TEST DESIGNS ON THREE LOBLOLLY PINE

(PINUS TAEDA L.) SITE TYPES

ABSTRACT

The environmental variances associated with family mean estimates for height were compared among six progeny test designs on three loblolly pine (Pinus taeda L.) progeny test sites in southeastern Oklahoma and southwestern Arkansas. Twenty-five- (or 30-) tree block plots, 10-tree row plots, 5-tree noncontiguous plots, 2-tree row plots and single-tree plots were compared. The 5-tree noncontiguous, 2-tree and single-tree plot designs consistently resulted in lower environmental variances associated with family means than the larger plots. The noncontiguous plot design resulted in environmental variances as low as those for the single-tree plot design, except on the most variable site. The environmental variance associated with family means was as small with 30 trees per family arranged in single-tree or 5-tree noncontiguous plots as with 60 trees per family arranged in 5-tree row plots.

INTRODUCTION

Commercial loblolly pine (Pinus taeda L.) sites in southeastern Oklahoma and southwestern Arkansas are usually highly variable. Progeny tests, designed to evaluate the breeding value of selected parent trees, must be established on a range of commercial site types. Thus, progeny tests in this region often contain a high degree of heterogeneity. Variability in a test site may appear as random microsite differences or as patches, that may be due to soil moisture, texture, or depth. Another common site type includes the variability associated with a gradient caused by distance from a water source, by a slope, or other factors. The precision in estimating family means in progeny tests depends upon the test design, and the importance of design efficiency may vary with site type.

Barnes and Schweppenhauser (1979) expressed the variance of a family mean determined from a replicated progeny test as:

$$\text{Var}(\bar{F}) = \sigma_f^2 + \sigma_p^2/r + \sigma_w^2/qr$$

where \bar{F} = family mean, σ_f^2 = family variance component, σ_p^2 = among-plot variance component, σ_w^2 = within-plot variance component, r = number of replicates, and q = number of trees per plot. The part of the variance of a family mean that is largely attributable to environment is $\sigma_p^2/r + \sigma_w^2/qr$.

Thus, the environmental portion of the among-plot and within-plot variance can be substantially altered by changing the relative plot size and number of replicates.

The theoretical value of the replicated single-tree plot design is well known (Libby and Cockerham 1980, Wright 1973). Single-tree plots maximize the number of microsites sampled, and replication size is minimized. Possible confounding caused by environmental covariances among members of a family is eliminated. A 2-tree plot design should be almost as efficient as the single-tree plot design. The number of replicates would be halved, but replicate size would still be small. A noncontiguous plot design, with each "plot" consisting of a number of trees of each family randomly located throughout a replicate, should have most of the advantages of the single-tree plot design (Lambeth et al. 1983). In addition, the probability of any family being absent in a replicate is much lower.

A 4- to 6-tree row plot design has probably been used more frequently than the other designs in past and present progeny tests. Ten-tree row plot designs and designs with longer rows or block plots have also been used extensively in the past. Barnes and Schweppenhauser (1979) indicated that 10-tree row plots were used in Pinus patula Shiede et Deppe progeny tests established between 1967 and 1972 because the authors were concerned about possible non-normality of the genetic component of among-plot effects. They subsequently examined the theoretical relationships

between within- and among-plot variances and concluded that a considerable increase in precision may result from using 3- to 5-tree row plots with more replicates. Evans et al. (1961) used 25-tree plots because they noted that the coefficient of variation became stable at approximately 20 trees. They were also concerned about possible non-normality of within-plot effects with small plots. Franklin (1971) noted, however, that the environmental effect is always sufficiently large relative to the genetic values to validate the assumption of normality. In addition the genetic values themselves are usually normally distributed (Falconer 1981).

The objective of this study was to compare the statistical efficiencies in estimating family means for height using 25- (or 30-) tree block plots, 10-tree row plots, 5-tree row plots, 5-tree noncontiguous plots, 2-tree plots and single-tree plots. The designs were compared on three distinctly different site types in southeastern Oklahoma and southwestern Arkansas.

MATERIALS AND METHODS

Height data from three Weyerhaeuser Company loblolly pine progeny tests were used to estimate environmental values for each site. The Dierks plantation, AOGC7505, is a half-sib progeny test planted in 1975, located near Dierks in southwestern Arkansas. The site is relatively uniform with a slight incline. The test area is rectangular with 12 replicates, each consisting of a 15 X 18 tree block with 54 families. Eight of the replicates consist of 5-tree row plots and the remaining four replicates consist of 5-tree noncontiguous plots. The Cloudy plantation, AOWF6088, is a full-sib progeny test planted in 1980, located near Cloudy in southeastern Oklahoma. The test straddles a ridge oriented in an east-west direction and six replicates are located on each side of the ridge on a steep slope. The replicates consist of 34 families of 5-tree row plots. Filler plots were used on microsites that were judged to be anomalous. The test shape is highly irregular. The Hempwallace plantation, AOWF6085, is a full-sib progeny test planted in 1980, located near Hempwallace in southwestern Arkansas. The site is relatively wet, causing a patchy pattern of variability. The test consists of 40 families arranged in 12 replicates, eight of which have a 5-tree row plot design and the remaining four have a 5-tree

noncontiguous plot arrangement. Filler trees were planted extensively on anomalous microsites, and the shape of the test is highly irregular. Eight-year data from the Dierks plantation and 5-year data from the other two plantations were used. The Dierks site was considered to be the most uniform; the Hempwallace site was judged to be intermediate and the Cloudy site was considered to be the most variable.

Residual values for height were obtained by subtracting the plantation mean and family effect from the height value for each tree in each plantation using the GLM procedure of the Statistical Analysis System (1982a). The residual values consisted of the remaining individual genetic effect and an effect associated with the microsite. Removal of full-sib family effects removes one-half of the additive genetic effect. In the case of half-sib families, one-quarter of the additive genetic effect is removed. It would be expected, however, that the environmental effect would be sufficiently larger than the remaining genetic effect to allow the residual values to primarily reflect the quality of the microsite, as suggested by Franklin (1971). If individual tree heritability is equal to .20, approximately 90% of the variance of the residual values in a full-sib progeny test would be due to nonadditive genetic and environmental causes.

The residual values became the environmental map for each plantation. Each position was given a row and column number and filler trees were indicated. As filler trees

were not measured in the plantations, they appeared as missing data in the mapped residual data sets. Values were approximated for filler trees using a random function of the Statistical Analysis System (1982b) to avoid the inflated apparent mortality. For each filler tree, a value was chosen at random from a normal distribution using the mean and variance of the surrounding 24 trees. Positions empty due to mortality were treated as missing data.

In comparing the six designs, it was important that each design fit in the same area of each plantation in such a way that replicates paralleled gradients to ensure valid comparisons. Because of these constraints, the number of usable positions varied among plantations. The rectangular Dierks plantation provided 2880 usable positions. The Hempwallace plantation had 1900 usable positions and only 1500 positions could be used from the Cloudy plantation. Each of the six designs was superimposed on the map of residual values using a computer model (Loo-Dinkins, Part II). Two variations in the number of families and number of trees within families were tested. Because the number of usable positions within each plantation was constant, increasing the number of families resulted in a proportionate decrease in the number of trees per family. The first set of conditions included 48 families with 60 trees per family on the Dierks plantation, 38 families with 50 trees per family on the Hempwallace plantation, and 30 families with 50 trees per family on the Cloudy plantation.

The numbers of families were increased to 96, 63, and 50 for Dierks, Hempwallace, and Cloudy, respectively, and each included 30 trees per family under the second set of conditions. The block plot design was not tested with the lower number of trees per family.

An analysis of variance was performed on the residual values with each of the superimposed designs to examine the partitioning of the variance components. The model used was:

$$\text{Residual} = \text{Rep} + \text{Plot}(\text{Rep}) + \text{Within Plot}.$$

The part of the total environmental variance that would contribute to the variance of a family mean, $\sigma_p^2/r + \sigma_w^2/qr$, denoted as $E\bar{F}$, was estimated for each site. The within and among plot variances would be subject to change among noncontiguous plot designs on the same environment depending upon the randomization pattern. Thus, for this design, the average values for σ_p^2 and σ_w^2 were estimated from 10 randomly assigned arrangements of noncontiguous plots for each site. Comparisons were made among the designs relative to the performance of the 5-tree row plot design, because this design is more frequently used than the others.

RESULTS

The environmental portion of the variance of a family mean, E_F , was larger for the large plot designs than for the designs with smaller plots (Tables 1, 2, and 3), as expected. Increasing the number of families to double or almost double those under the first set of conditions generally doubled or almost doubled E_F . The increase in the variance as the number of families was increased was greater in the case of 5-tree and 10-tree row plots than for the 5-tree noncontiguous, 2-tree plot, or single-tree plot designs, however.

The 25- (or 30-) tree block plots produced a higher E_F value than did any other design on each site. On the least variable site, the Dierks plantation, E_F was approximately three times larger for the 30-tree block plot design than for the 5-tree row plot design. E_F was five and 10 times greater for the 25-tree block plot design than for the 5-tree row plot design on the second most uniform site, the Hempwallace plantation, and on the most variable site, the Cloudy plantation, respectively. The block plot designs also performed substantially worse than the 10-tree row plot design on each site. The 10-tree row plot design consistently had the second highest E_F value when the number of families was at the lower number. The 10-tree row plot

design consistently had the highest $E\bar{F}$ value when the higher number of families was tested, as block plots were only tested with the lower number of families. $E\bar{F}$ values resulting from 10-tree row plot designs were consistently approximately 1.4 times greater than the $E\bar{F}$ values for the 5-tree row plot design, regardless of site and number of families tested.

The 5-tree row plot design consistently had smaller $E\bar{F}$ than block plot or 10-tree row plot designs, but larger $E\bar{F}$ values than the 5-tree noncontiguous, 2-tree, or single-tree plot designs. The superimposition of the 5-tree row plot design on each of the plantations resulted in plots being placed on the same five positions as the original five-tree row plots. Thus, error in estimating family effects would result in higher among plot variance and lower within plot variance for the five-tree row plot design than could be attributed to the environment. To test this, mean correlation coefficients were estimated between residual values for each position within a moving five by five-tree area (Loo-Dinkins, Part III). Correlations between residual values within the original five-tree plot areas were no larger, on the average, than correlations between residuals from parallel plots. Thus, possible confounding effects in $E\bar{F}$ estimates for five-tree plots are minimal.

The 5-tree row plot $E\bar{F}$ ratio for low to high number of families (high to low number of tree per families) was greatest on the most variable site. On the most uniform

site, the Dierks plantation, the numbers of trees per family tested were 60 and 30 (2:1) and the E_F ratio was 1:1.997. The E_F ratio increased to 1:2.022 on the second most uniform site, the Hempwallace plantation, while the ratio of high to low number of trees per family decreased to 5:3. On the most variable site, the Cloudy plantation, the ratio of high to low number of trees per family tested was also 5:3, but the E_F ratio increased to 1:2.184. The opposite trend was apparent for the five-tree noncontiguous plots, with ratios of E_F values from high to low numbers of trees per family of 1:1.970, 1:1.714, and 1:1.667 for the most uniform to the most variable sites, respectively. Thus, relative to the 5-tree row plot design, the 5-tree noncontiguous plot design improved with increased numbers of trees per family on the sites with greatest variability.

The 5-tree noncontiguous plot design resulted in E_F values that were always as small as or smaller than those for 5-tree row plots. The noncontiguous plot design also compared favorably with the 2-tree plot design, but except for the Hempwallace site with the low number of families, the single-tree plot design resulted in lower E_F values. The difference was slight on the most uniform site (Table 1).

The 2-tree plot design always performed better than the 5-tree row plot design on each site regardless of number of families tested, but generally slightly worse than the 5-tree noncontiguous plot design. The 2-tree plot design did

not improve as much relative to the 5-tree row plot design when the number of families was increased as did the noncontiguous plot and single-tree plot designs.

The single-tree plot design generally yielded the lowest $E\bar{F}$ values in spite of missing plots due to a 5 to 15% mortality rate in the plantations which resulted in the absence of some families from some replicates. The among replicate difference was greatest when the higher number of families was tested; thus more variance was removed by the replicate effect for the higher number of families than for the lower number with single-tree plots. The single-tree plot $E\bar{F}$ values for the high number of families were similar or a little higher than those for the 5-tree row plot design with low number of families.

DISCUSSION

Statistical theory predicts that either increasing plot size or reducing replicate number increases the variance of a family mean (Cochran and Cox 1957). Also the site type would be expected to influence the magnitude of differences resulting from the various designs. Results of this study indicate that the site type may have a strong impact on the relative efficiencies of various plot sizes and replicate number.

Large block plots are sometimes used to estimate the yield performance of a family by harvesting the entire block. This design should only be used on uniform sites if a realistic estimate of family performance is to be obtained. The performance of the block plot design decreased drastically relative to that of the 5-tree row plot design on the more variable sites.

The difference in E_F values between the 5-tree and 10-tree row plot designs (approximately 1:1.4) did not change substantially with site or when the number of families was increased at the expense of the number of trees per family. In most cases, the magnitude of the individual variance components, σ_r^2 , σ_p^2 , and σ_w^2 did not greatly change between the 5-tree and 10-tree row plot designs. The difference in E_F values between the two designs appeared to result

primarily from the reduction by half of the divisor for σ_p^2 in $E\bar{F}$ values for the 10-tree row plot design. Each of the other designs responded differently on each site relative to the 5-tree row plot design, implying that other factors were more important than the magnitude of the divisor in those cases. On the most uniform site, more variance was removed by the replication with 10-tree row plots than with 5-tree row plots. On the more variable sites, this tended to be reversed. In general, the replicates containing 2- and single-tree plots removed more variance than in the case of larger plot designs, but replication appears to be much less important on the two more uniform sites than on the most variable site.

Five-tree row plots are extensively used in many tree improvement programs. They are relatively easy to lay out, map, and thin by family (Lambeth et al. 1983). They are also small enough to avoid the more serious problems of the larger plots. Results of this study indicate that the convenience of 5-tree row plot designs may not outweigh the increased efficiency of the 5-tree noncontiguous, 2-tree, or single-tree plot designs. Lambeth et al. (1983) reported that 60 individuals arranged in 5-tree noncontiguous plots would provide an estimate of a family mean as efficiently as 85 individuals arranged in 5-tree row plots. The ratio appears to be site specific. On the most uniform site, 30 trees per family arranged in 5-tree noncontiguous plots produced an $E\bar{F}$ value very similar to that of 60 trees per

family arranged in 5-tree row plots. This implies that, on a uniform site, 30 trees per family arranged in a 5-tree noncontiguous plot design may provide as accurate an estimate of family means as twice that number of trees per family in 5-tree row plots. On the less uniform sites, the noncontiguous plot designs with 30 trees per family performed at a level between 5- and 10-tree row plots with 50 trees per family. The difference may be explained by the observation that the ratio of within- to among-plot variance for the 5-tree row plot was much greater on the more variable sites than on the most uniform site. The noncontiguous plot design performed very well compared with the 5-tree row plot design on the most uniform site because the among-plot variance is much lower for the noncontiguous plot design. The ratio of within- to among-plot variance was always very high for noncontiguous plots. Because the divisor for within-plot variance is five times the divisor for among-plot variance in the expression for $E\bar{F}$, the within- to among-plot variance ratio is the primary difference between the efficiencies of row and noncontiguous plot designs. As the ratio of within- to among-plot variance increased from 3:1 to 41:1 for the 5-tree row plots from the most uniform site to the most variable site, the difference between the 5-tree row plot and 5-tree noncontiguous plot designs decreased.

The noncontiguous plot design generally produced lower $E\bar{F}$ values than the 2-tree plot design except on the most

variable site. The $E\bar{F}$ values for noncontiguous and single-tree plot designs were similar on the two most uniform sites but the single-tree plot design was considerably better on the most variable site. The primary difference between the noncontiguous and single-tree plot $E\bar{F}$ values is in the amount of variance removed by the replication. On the two most uniform sites, the relative sizes of replicates were not reflected in the size of σ_r^2 . Replication was much more important on the most variable site, as would be expected.

The single-tree plots generally out-performed the other designs, although on the most uniform sites, the noncontiguous plot design was a close second. The difference between single-tree plots and 5-tree row plots was considerably smaller on the most variable site than on the other two sites due to orientation of the replicates. The replicates paralleled the gradient and the family plots extended across the gradient. Thus, most of the within-replicate variance was contained within plots rather than among plots. The 5-tree row within-plot variance component would have the same divisor as the single-tree plot among-plot variance. Thus, the smaller the among- to within- plot variance ratio, the closer 5-tree row plots will be in performance to single-tree plots, regardless of the size of the replicate variance component.

CONCLUSIONS

A 5-tree noncontiguous, 2-tree, or single-tree plot design should always produce more accurate estimates of family means than 5-tree row plots. The difference between the 5-tree row plot design and the more efficient designs is greatest on sites with a random pattern of variability where a significant part of the within-replicate variability is due to-plot-to plot variability. On a site with a strong gradient, if the replication is done efficiently, most of the variance will be contained within family plots, allowing the 5-tree row plot design to compare more favorably with the more efficient designs. Accuracy of family mean estimates, on a relatively uniform site, may be as high with only 30 trees per family arranged in noncontiguous or single-tree plots as with 60 trees per family arranged in 5-tree row plots. This would allow twice as many families to be tested within the same area without loss of precision, allowing greater genetic gain per unit cost in progeny testing.

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Table 1. Partitioning of the environmental portion of the variance of a family mean for height in loblolly pine, $E_{\bar{F}}$, for the most uniform site, the Dierks plantation

Design	No. reps	σ_r^2	σ_p^2	σ_w^2	$E_{\bar{F}}$	^a Rel value
<u>48 family plots/rep</u>						
30-tree blocks	2	.1200	.1283	.8333	.0887	2.99
10-tree rows	6	.1312	.1690	.7410	.0405	1.36
5-tree rows	12	.1247	.2191	.6858	.0297	1.00
5-tree nc plots	12	.1284	.0050	.9009	.0154	0.52
2-tree plots	30	.1378	.1326	.7527	.0166	0.56
Single-tree plots	60	.1596	.8619		.0144	0.48
<u>96 family plots/rep</u>						
10-tree rows	3	.1589	.1725	.7410	.0822	1.39
5-tree rows	6	.1378	.2185	.6858	.0593	1.00
5-tree nc plots	6	.1398	.0000	.9135	.0304	0.51
2-tree plots	15	.1734	.2237	.6291	.0359	0.61
Single-tree plots	30	.1844	.8405		.0280	0.47

^a $E_{\bar{F}} = \sigma_p^2/r + \sigma_w^2/rq$, where σ_p^2 is the among plot variance, r is the number of replicates, σ_w^2 is the within plot variance, and q is the number of individuals in a plot. ^b σ_r^2 is the among replicate variance.

^cRelative value refers to the size of $E_{\bar{F}}$ for each design relative to $E_{\bar{F}}$ for the 5-tree row plot design.

Table 2. Partitioning of the environmental portion of the variance of a family mean for height in loblolly pine, $\mu E\bar{F}$, for the second most uniform site, the Hempwallace plantation.

Design	No. reps	$b\sigma_r^2$	σ_p^2	σ_w^2	$E\bar{F}$	cRel value
<u>38 family plots/rep</u>						
25-tree blocks	2	.0692	.0415	.1533	.0238	5.12
10-tree rows	5	.0643	.0176	.1600	.0067	1.44
5-tree rows	10	.0655	.0156	.1544	.0046	1.00
5-tree nc plots	10	.0659	-.0034	.1730	.0035	0.74
2-tree plots	25	.0648	.1160	.1641	.0037	0.81
Single-tree plots	50	.0627	.1675		.0036	0.78
<u>63 family plots/rep</u>						
10-tree rows	3	.0652	.0247	.1602	.0135	1.45
5-tree rows	6	.0581	.0250	.1548	.0093	1.00
5-tree nc plots	6	.0584	.0003	.1792	.0060	0.64
2-tree plots	15	.0651	.0292	.1601	.0066	0.70
Single-tree plots	30	.0641	.1662		.0055	0.59

$\mu E\bar{F} = \sigma_p^2/r + \sigma_w^2/rq$, where σ_p^2 is the among plot variance, r is the number of replicates, σ_w^2 is the within plot variance, and q is the number of individuals in a plot. $b\sigma_r^2$ is the among replicate variance.

cRelative value refers to the size of $E\bar{F}$ for each design relative to $E\bar{F}$ for the five-tree row plot design.

Table 3. Partitioning of the environmental portion of the variance of a family mean for height in loblolly pine, $E\bar{F}$, for the most variable site, the Cloudy plantation.

Design	No. reps	σ_r^2	σ_p^2	σ_w^2	$E\bar{F}$	^a Rel. value
<u>30 family plots/rep</u>						
30-tree blocks	2	.0408	.0916	.2071	.0499	10.21
10-tree rows	5	.0814	.0122	.2403	.0072	1.47
5-tree rows	10	.1045	.0053	.2183	.0049	1.00
5-tree nc plots	10	.1045	.0044	.2191	.0048	0.98
2-tree plots	25	.1551	.0238	.1449	.0038	0.79
Single-tree plots	50	.1386	.1821		.0036	0.74
<u>50 family plots/rep</u>						
10-tree rows	3	.0852	.0207	.2403	.0149	1.39
5-tree rows	6	.0944	.0207	.2183	.0107	1.00
5-tree nc plots	6	.0948	-.0005	.2413	.0080	0.75
2-tree plots	15	.1267	.0433	.1563	.0081	0.75
Single-tree plots	30	.1524	.1703		.0057	0.53

^a $E\bar{F} = \sigma_p^2/r + \sigma_w^2/rq$ where σ_p^2 is the among plot variance, r is the number of replicates, σ_w^2 is the within plot variance, and q is the number of individuals in a plot.

^b σ_r^2 is the among replicate variance.

^cRelative value refers to the size of $E\bar{F}$ for each design relative to $E\bar{F}$ for the five-tree row plot design.

PART II

COMPUTER SIMULATIONS TO COMPARE COMBINATIONS

OF PROGENY TEST DESIGNS AND SELECTION

SYSTEMS FOR MAXIMUM GENETIC GAIN

IN LOBLOLLY PINE (PINUS

TAEDA L.)

ABSTRACT

Six progeny test designs in combination with four within family selection systems were compared on three loblolly pine (Pinus taeda L.) progeny test sites in southeastern Oklahoma and southwestern Arkansas. Large block plots, 5- and 10-tree row plots, 5-tree noncontiguous plots, 2-tree plots, and single-tree plots were tested in combination with selection based on deviations from replicate means, deviations from neighborhood means, a combination of replicate and neighborhood deviations, and deviations from plot means. Large block plots performed consistently poorly compared with the other designs, and selection based on deviations from a plot mean was always poor compared with the other selection systems. The single-tree plot design in combination with replicate or neighborhood deviation selection performed best on all sites. The 5-tree noncontiguous and 2-tree plot designs generally performed better than the larger row or block plots, particularly when the number of families tested was increased at the expense of number of trees per plot.

INTRODUCTION

Loblolly pine (Pinus taeda L.) is the most important commercial forest species in southeastern Oklahoma and southern Arkansas. The species is often planted on sites that are highly variable. This variability may be due to such factors as slope, aspect, soil moisture conditions, soil texture and soil fertility. Commercial forest sites typically include much more heterogeneity than do agricultural fields. The heterogeneity may appear as random microsite variability; it may be characterized by larger patches that grade into one another; or it may be in the form of a definite gradient.

The primary purpose of progeny testing is to evaluate family performance in order to rank potential parents of the next generation of commercial seedlings. Progeny tests must sample a range of commercial planting sites in order to predict the commercial performance of a family. As the size and purpose of progeny tests does not allow avoidance of much of the variability, the test sites may be highly variable. A secondary but equally important function of many progeny tests is to provide a base for selection of the next generation of parents. Thus, in addition to accurate estimation of family means, which depends primarily on the test design, an optimal within-family

selection system is important.

PROGENY TEST DESIGNS

The variance associated with a family mean estimate is largely determined by among-plot and within-plot variance (Barnes and Schweppenhauser 1979). The among-plot variance can be substantially altered by changing the relative plot size and number of replicates. Given a fixed plantation size and number of families tested, size of plots will be inversely related to the number of replicates. In addition, the number of trees per family will vary inversely with the number of families that can be tested. On a variable site, the relative size and arrangement of plots largely determines the statistical efficiency of the test (Libby and Cockerham 1980, Loo-Dinkins, Part I).

Statistical theory indicates that the smaller the plot size, the more efficient it should be (Libby and Cockerham 1980). Single tree plots maximize the number of microsites sampled and minimize replicate size (Wright 1973). Possible confounding of environmental and genetic covariances among members of a genetic unit is eliminated. In practice, use of single tree plots may lead to many missing cells due to the combined effects of thinning and unplanned mortality. A 2-tree plot design would also ensure thorough microsite sampling by each family and replicates would be small, although twice as large as replicates of single-tree plots.

A noncontiguous plot design, with each plot consisting

of a number of trees randomly located throughout a replicate, would have most of the advantages of the single-tree plot design. In addition, the probability of a given family being absent from one or more replicates is much lower (Lambeth et al. 1983, Libby and Cockerham 1980). Lambeth et al. (1983) compared a test design consisting of 5-tree noncontiguous plots with a 5-tree row plot design and found that among plot variance was greater for row plots than for noncontiguous plots. Thus the number of replicates required to obtain a given level of precision is lower for noncontiguous plots than for row plots.

Four- to 6-tree row plots constitute the most extensively used design in today's forest tree improvement programs. Field layout and silvicultural thinning by family are relatively easy. Plots are large enough that analysis on a plot means basis usually would not result in missing cells. Larger row plots and block plots have been used frequently in the past in progeny test designs. Early tree-improvement workers were concerned about possible non-normality of among- or within-plot effects with small plot size (Evans et al. 1961, Barnes and Schweppenhauser 1980). Franklin (1971) noted that the environmental effect is always sufficiently large relative to the genetic values to validate the assumption of normality.

Loo-Dinkins (Part I) compared the theoretical efficiencies of a number of progeny test designs on three site types. She found that the relative efficiency of

designs varied with the site type, but that single-tree, two-tree and noncontiguous plot designs always resulted in lower variance associated with family means than larger row or block plots. Her results indicated that, on a relatively uniform site, 30 trees per family, arranged in noncontiguous or single-tree plots, may provide as accurate an estimate of family means as would 60 trees per family arranged in 5-tree row plots.

WITHIN FAMILY SELECTION

Expressions for expected genetic gain have been derived and are well known for the traditional selection systems (Shellbourne 1969, Namkoong 1979). Some combination of among-and within-family selection is usually utilized when selecting in a progeny test. The optimal combination of among- and within-family selection depends primarily upon the heritability of the trait under improvement (Falconer 1981). In practice, the relative degree of within- and among-family selection is also influenced by inbreeding considerations. Frequently, the best families are identified and one or two individuals within each family are selected. The optimal method of selection of individuals within families may vary with design and site type.

Current within-family selection methods include use of deviations from plot means and deviations from replicate means in the initial stages, to identify trees worthy of a field assessment. The Western Gulf Forest Tree Improvement

Cooperative selects trees from progeny tests using deviations from plot means among other considerations. The North Carolina State Forest Tree Improvement Cooperative includes deviations from a replicate mean as part of the selection criteria within progeny tests.

The optimal size and shape of a neighborhood comparison area may fall between a plot and a replicate. Gardner (1961) described a method to reduce the environmental variance within the selection area for corn. He superimposed a grid on the test area and selected a set number of individuals from within each square. A similar procedure that would be applicable to selection of forest trees in a progeny test would be to superimpose a grid on the plantation to create neighborhoods. Within-family selections would then be made on the basis of deviations from neighborhood means. The neighborhoods should be small enough that on the average the environmental correlations between positions would be greater than zero. The size and shape of neighborhoods would vary with site type (Loo-Dinkins, Part III).

The theory of progeny test design efficiency and expected gains for various selection systems have been examined, but not in combination. Combinations of progeny test designs and within-family selection systems can be compared on various site types, using a simulation model. A simulation model imposes limitations on the kinds of selection systems that can be tested; and results provide

guidelines, only as good as the assumptions used in building the model. Modeling may be the only way to compare different designs on one site, however, and the technique may provide new insights.

The objective of the study was to compare the influence of various progeny test designs in combination with several within-family selection methods on the genetic gain in 5- or 8-year height in loblolly pine. The efficiency of the various combinations was compared using an environmental data base generated from three progeny test sites with superimposed genetic populations generated by a simulation model.

MATERIALS AND METHODS

Environmental data were obtained from three Weyerhaeuser Company loblolly pine progeny tests occupying distinctly different site types in southeastern Oklahoma and southwestern Arkansas. The shape and layout of replicates for each plantation is shown in Fig. 1. Dierks plantation, AOGC7505, is a half-sib progeny test, planted in 1975, located near Dierks, Arkansas. The site is relatively uniform with apparently random microsite variability. The test area is rectangular with 12 replicates each consisting of a 15 X 18 tree block with 54 families. Eight of the replicates consist of 5-tree row plots. The remaining four replicates contain 5-tree noncontiguous plots. The Cloudy plantation, AOWF8088, is a full-sib progeny test planted in 1980, located near Cloudy, Oklahoma. The test contains a strong gradient as it straddles a ridge with an east-west orientation. Six replicates consisting of 34 5-tree family row plots are located on each side of the ridge. Filler plots were used on microsites which were judged to be anomalous. The test shape is highly irregular (Fig. 1). The Hempwallace plantation, AOWF8085, is a full-sib progeny test planted in 1980, located near Hempwallace, Arkansas. The site is relatively wet, causing a pattern of patchy variability.

The test consists of 40 families arranged in 12 replicates; eight of which have a 5-tree row plot arrangement. The remaining four replicates have a 5-tree noncontiguous plot arrangement. Filler trees were planted extensively on anomalous microsites and the shape of the test is highly irregular. Eight-year height data were used from the Dierks plantation, and 5-year height data were used from the other two plantations. The survival in the three plantations ranged from 85 to 95%.

Residual values for height were obtained by subtracting the family effect and overall plantation mean from the height of each tree using the GLM procedure of the Statistical Analysis System (1982a). The resultant deviations with a mean of zero included a portion of the within family genetic effects in addition to the true environmental effects. The environmental effects were assumed to be sufficiently larger than the remaining genetic effects to allow the residual value to primarily reflect a measure of the microsite value, as suggested by Franklin (1971). For example, if individual tree heritability is .20, approximately 90% of the variance of the residual values in a full-sib progeny test would be due to nonadditive genetic and environmental causes.

The environmental residual data sets were mapped using a row and column coordinate system. The original 5-tree plots in each plantation fell within columns. As filler trees were not measured in plantations, they appeared as

missing values in the mapped residual data sets. Thus they were indistinguishable from trees missing because of mortality. To avoid the problem of an apparent inflated mortality rate, environmental values were approximated for the filler positions. The mean and variance of neighboring 24 trees, including two trees on each side of the filler position, were used to generate approximate environmental values for the filler positions with a random normal function of the Statistical Analysis System (1982b) A grid pattern was superimposed on each data set to establish 5 X 5-tree neighborhoods for neighborhood selection.

Six planting designs were tested on each environmental data set with a relatively low number of families (30 to 48) and a high number of trees per family (50 or 60). Block plots, 5 X 6 trees on the Dierks plantation, and 5 X 5 trees on the other two plantations were arranged in two replicates. The block plots could not be tested when the number of families was increased (50, 63, or 96) and number of trees per family was reduced to 30. The remaining five designs were tested under both sets of conditions. Ten-tree row plots, 5-tree row plots, 5-tree noncontiguous plots, 2-tree row plots, and single-tree plots were arranged in the appropriate number of replicates for each data set. The within-family selection systems included selection based on deviations from a replicate mean, deviations from a plot mean, deviations from a neighborhood mean, and a combination of replicate and neighborhood mean deviations.

The combination selection method involved choosing the best ten trees from each selected family on the basis of deviations from the replicate mean and the best two trees from within the ten were then chosen on the basis of neighborhood deviations.

A model was constructed using SAS Macro Language (1982b) to generate genetic populations with an among- and within-family variance structure to correspond to a heritability of .20 or .30, and to superimpose each genetic population, arranged in each of the designs on the environmental data sets. These heritabilities were chosen because they are within the range of reported heritability values for height in loblolly pine (Zobel and Talbert 1984). The genetic structure simulated half-sib families with $3/4$ of the additive genetic variance occurring within families for the Dierks plantation. The genetic structure simulated full-sib families for the other two plantations. The original plantation mean was added to the additive genetic and environmental values for each design to result in simulated height values corresponding to each position. The ten best families were chosen based on ranking of family means and the within family selection systems were applied to choose two phenotypes from each of the chosen families.

The model was iterated 100 times for each set of conditions. The two heritability levels (.20 and .30) were modeled at a high and low number of families on each of the

three data sets. The test design-selection system combinations were compared by examining the mean simulated genetic value of the selected phenotypes relative to the maximum possible mean genetic value. The maximum possible genetic value was calculated as the mean genetic value of the actual two best genotypes from the 10 families with the highest family genetic values. The relative performance of a design-selection system combination was then expressed as the percentage of the maximum that was attained. The mean percentages were calculated over 100 iterations. The relative proportions of the genetic value of selected trees due to family versus individual tree genetic effects were determined under some sets of conditions.

RESULTS

The differences among progeny test design-selection system combinations, expressed as the mean percentage of the maximum genetic value that could be attained, are less than might have been expected based on relative statistical efficiencies of the designs (Loo-Dinkins, Part I, Lambeth et al. 1983) (Tables 1-12). The percentage of the maximum possible mean genetic value of the 20 selected phenotypes was consistently between 50 and 75%, regardless of test design, selection system, site type, or heritability. The standard errors associated with the mean percentage for 100 iterations were consistently small, ranging from 0.50 to 0.95. Thus, small differences, on the order of 1.6 to 2.5 percentage points, between percentages of the maximum possible mean genetic value could be declared statistically significant. Differences among designs, among selection systems across designs, and among design-selection system combinations were found to be highly significant using an analysis of variance. The differences between design-selection system combinations decreased when heritability was increased from .20 to .30. The percentage of the maximum genetic value increased by five to 10 points with increased heritabilities and the increase was greatest for designs with lowest percentages of maximum genetic value.

On the two most variable sites, Cloudy and Hempwallace, the percentage of the maximum genetic value increased by eight to 10 points with the 25-tree block plot design but only by five to six points for the single-tree plot design when heritability was increased from .20 to .30.

The large block plot designs consistently performed poorly in comparison with the other designs regardless of the selection system. Differences among the other designs were slight in most cases but a trend for small plots to perform better than larger plots was apparent under each set of conditions for each site. The 5-tree row plot design was almost always slightly better than the 10-tree row plot design and the noncontiguous, single-tree and 2-tree plot designs almost always performed better than the 5-tree row plot design. The trend was strongest when the lower number of trees per family with the higher number of families was tested.

Selection based on deviations from a plot mean consistently resulted in lower percentages of the maximum genetic value than any of the other selection systems. This method performed particularly poorly in combination with the 2-tree plot design. Selection based on deviations from plot means performed best relative to the other selection systems when the design consisted of large block or row plots. As plot size decreased the difference between percentages resulting from plot deviation selection and percentages resulting from the other selection methods

increased. Selection based on deviations from a replicate mean generally improved as plot size, and consequently, replicate size decreased. Selection based on deviations from a neighborhood mean was always better than plot mean deviation selection and generally slightly better than selection based on deviations from a replicate mean when the plot size was greater than five trees.

When the number of families was increased at the expense of the number of trees per family, differences between the designs were greater. Generally the row plot designs performed slightly worse and the 5-tree noncontiguous, single-tree and 2-tree plot designs resulted in slightly higher percentages of the maximum genetic value than were attained with the smaller number of families.

The average proportion of the genetic value of selected phenotypes that was due to the individual genetic value relative to the family genetic value varied from 0.40 to 0.60 (Tables 13 and 14). The proportions for the Dierks site were higher than for the other two plantations because the genetic structure in the Dierks plantation is half-sib, thus three-quarters of the genetic variance is within family. The other two plantations have a full-sib family structure, thus one-half of the genetic variance would be within family.

In general, the designs showing the highest proportions attributable to individual genetic values are those which attained lowest percentages of the maximum genetic value.

Within designs, however, the selection systems that resulted in the highest individual proportions of the genetic value also had highest percentages of the maximum genetic value. The proportion of the mean genetic value of selections that was due to the individual genetic value increased by approximately 0.10 when the number of families in the Cloudy plantation decreased from 50 to 30, and number of trees per family was increased from 30 to 50 (Tables 13 and 14). The number of families chosen remained constant at 10, allowing the selection intensity within families to increase from $1/15$ to $1/25$, while the among family selection intensity decreased from $1/5$ to $1/3$ when the number of families was reduced. The overall intensity did not change but the composition of the mean genetic value shifted substantially in the individual direction. This shift was accompanied by a decrease in the percent of the maximum genetic value attained by the most efficient designs. Increasing the heritability from .20 to .30 resulted in a slight increase in the proportion of the genetic value that was due to individual values for the Cloudy site. This was not examined on the other sites.

DISCUSSION

Statistical theory indicates that differences among the designs, regardless of the selection system, could be expected to be greater than those found in this study (Lambeth et al. 1983, Loo-Dinkins, Part I). The designs showed greater differences when number of families was increased and number of trees within family was decreased. It appears that given 50 or 60 trees per family and 30 to 50 families, the design is much less important, on the sites tested, than when the number of trees per family is reduced to 30. In particular, the difference between 5- and 10-tree row plot designs is significant only when 30 trees per family were tested with a high number of families.

The percentage of the maximum genetic value that was attained with 5-tree noncontiguous, single, and 2-tree plot designs generally increased when the number of families was increased and number of trees per family decreased. The percentage attained using 5- and 10-tree row plot designs generally remained approximately the same or decreased slightly when the number of families was increased and number of trees per family decreased. This trend was accompanied by a general decrease in the proportion of the mean genetic value that reflected the individual value,

when the higher number of families was tested. Thus, at the heritability values tested, the increased number of families with fewer trees per family, allowing higher among-family selection intensity, led to greater genetic gains when the designs that allowed most accurate estimates of family genetic values were used. The higher among family selection intensity did not result in higher gains, however, when the less efficient designs were used. This result implies that if the more efficient designs are used, a greater number of families may be tested in the same space, resulting in increased genetic gain as well as increased economic efficiency.

Selection based on deviations from a plot mean consistently resulted in lower percentage of the maximum genetic value than for each of the other selection systems. As the number of trees per plot increased, the percentage of the maximum genetic value attained by using plot deviation selection increased to a plot size of 10 trees, then decreased with 25- (or 30-) tree block plots (Fig. 2). The selection systems based on deviations from replicate or neighborhood means yielded lower percentages of the maximum genetic value as plot size increased. Replicate deviation selection appeared slightly better with single and 2-tree plots, and neighborhood deviation selection performed better, on the average, for the block plot designs. Five- and 10-tree row plot designs did not differ for selection based on deviations from replicate and

neighborhood means when the relatively large number of trees per family was tested. Differences between the 5- and 10-tree row plot design percentages were apparent with the replicate and neighborhood deviation selection systems when 30 trees per family were tested (Fig. 2). There does not appear to be a difference between the 5- and 10-tree row plot designs, however, when the selection method was based upon deviations from plot means at the higher number of families. These results imply that with the larger number of families, the improved selection accuracy that could be expected of the plot deviation method when the plot size increased was offset by the negative influence of the larger replicate size. The size of the 10-tree row plot replicates approximately doubled when the number of families was increased, in each case, resulting in less efficient estimates of family value relative to the 5-tree row plot design.

The percent of the maximum genetic value, averaged over the three sites for the 5-tree noncontiguous design, was very close to the 2-tree plot results when the replicate and neighborhood deviation selection systems were used. The mean percentage value for noncontiguous plots was similar to that for 5-tree row plots with the relatively low number of families, and better than any of the other designs tested with the high number of families, when selection was based on deviations from plot means.

It was expected that the progeny test design-within

family selection system combination would be most important on the most variable site. The Cloudy site appeared to have the greatest amount of variability, with the plantation slanting down-slope on either side of a ridge. Differences between 10-tree row plot, 5-tree row plot and 5-tree noncontiguous plot designs were less than for the other designs with 30 trees per family, however. Selections at the Cloudy site had the lowest mean proportion of individual vs. family genetic value (Table 13). Thus, accurate estimates of family means appeared to be more important in determining the mean genetic value of selected trees at the Cloudy site than at the other sites. Replicates of 5- and 10-tree row plots on the Cloudy site were arranged so that the family plots extended across the gradient. As most of the variation in the site was due to the gradient (Loo-Dinkins, Part III), each family row plot sampled most of the variation within a particular replicate, allowing accurate estimates of family genetic value. Thus 5- or 10-tree row plot designs may be very efficient on sites in which a strong, identifiable gradient is the main source of variation, if the plots extend across the gradient.

CONCLUSIONS

Differences in efficiency of estimation of family genetic value and selection accuracy among designs which include 10 trees or less in replicated family plots do not appear to be important if 50 or 60 trees per family are included. When the number of trees per family is reduced to 30 and the number of families tested is increased, the single-tree, 2-tree, and 5-tree noncontiguous plot designs are significantly better than the larger row plot designs. The single-tree plot design consistently performed the best under these conditions in spite of a 5-15% level of mortality resulting in some missing plots.

The optimum neighborhood size for selection may be greater than the 25-tree neighborhoods that were tested, but is smaller than the size of a replicate with 5-tree plots and 30 or more families. Selection based on deviations from a replicate mean is as good as selection based on deviations from a 25-tree neighborhood mean when replicates contained family plots with 10 trees or less.

Selection based on either deviations from a replicate mean or deviations from neighborhood means produces higher genetic gains than selection based on deviations from the plot mean, especially when plot size is five trees or less.

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Table 1. Mean percent of maximum genetic value for six designs and four selection systems on the Dierks plantation with 60 trees per family and $h^2 = .2$.

Design	Selection System			
	Rep Dev	Comb	N'hood	Plot
30-tree block plots	53.9	55.8	56.2	52.5
10-tree row plots	59.6	60.2	59.7	56.6
5-tree row plots	59.7	60.3	59.5	53.3
5-tree noncontiguous plot	59.6	61.3	60.6	53.9
2-tree plots	61.2	60.6	60.5	47.2
single tree plots	62.0	60.2	60.1	-

Standard errors ranged from 0.70 to 0.95.

Table 2. Mean percent of maximum genetic value for six designs and four selection systems on the Dierks plantation with 60 trees per family and $h^2 = .3$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
30-tree block plots	62.0	65.3	64.1	61.0
10-tree row plots	65.9	66.3	65.6	62.6
5-tree row plots	66.5	66.6	66.3	59.9
5-tree noncontiguous plots	67.6	69.6	69.3	60.6
2-tree plots	68.3	67.7	67.3	53.3
single tree plots	69.8	68.7	68.5	-

Standard errors ranged from 0.50 to 0.83.

Table 3. Mean percent of maximum genetic value for five designs and four selection systems on the Dierks plantation with 30 trees per family and $h^2 = .2$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
10-tree row plots	56.9	55.3	56.8	53.8
5-tree row plots	58.2	57.8	58.3	53.9
5-tree noncontiguous plots	61.5	60.8	61.8	57.5
2-tree plots	62.3	61.0	62.3	51.4
single tree plots	63.1	62.1	63.3	-

Standard errors ranged from 0.63 to 0.89.

Table 4. Mean percent of maximum genetic value for five designs and four selection systems on the Dierks plantation with 30 trees per family and $h^2 = .3$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
10-tree row plots	67.0	66.0	65.6	64.1
5-tree row plots	68.0	69.5	69.5	63.7
5-tree noncontiguous plots	67.6	69.5	69.4	63.5
2-tree plots	70.0	70.7	70.6	60.6
single tree plots	69.4	70.5	70.9	-

Standard errors ranged from 0.68 to 0.79.

Table 5. Mean percent of maximum genetic value for six designs and four selection systems on the Cloudy plantation with 50 trees per family and $h^2 = .2$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
25-tree blocks	55.4	55.1	55.4	54.8
10-tree row plots	63.4	62.0	63.1	58.3
5-tree row plots	62.7	63.6	62.7	57.6
5-tree noncontiguous plots	62.2	63.1	62.3	55.8
2-tree plots	64.6	66.1	64.3	54.2
Single-tree plots	65.0	65.9	64.5	-

Standard errors ranged from 0.60 to 0.94.

Table 6. Mean percent of maximum genetic value for six designs and four selection systems on the Cloudy plantation with 50 trees per family and $h^2 = .3$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
25-tree block plots	65.4	64.3	64.9	63.3
10-tree row plots	69.5	67.8	69.4	63.4
5-tree row plots	69.7	70.6	69.6	62.8
5-tree noncontiguous plots	70.9	71.4	70.7	63.1
2-tree plots	69.8	72.0	69.5	59.9
single tree plots	71.2	71.8	70.5	-

Standard errors ranged from 0.50 to 0.78.

Table 7. Mean percent of maximum genetic value for five designs and four selection systems on the Cloudy plantation with 30 trees per family and $h^2 = .2$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
10-tree row plots	65.5	65.6	65.7	62.8
5-tree row plots	65.6	65.5	65.4	61.3
5-tree noncontiguous plots	67.6	67.0	67.3	60.5
2-tree plots	67.5	67.4	66.6	58.1
single tree plots	70.0	68.5	68.3	-

Standard errors ranged from 0.61 to 0.79.

Table 8. Mean percent of maximum genetic value for five designs and four selection systems on the Cloudy plantation with 30 trees per family and $h^2 = .3$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
10-tree row plots	70.8	71.1	71.1	67.2
5-tree row plots	72.7	71.7	71.7	66.8
5-tree noncontiguous plots	73.9	73.2	73.2	66.9
2-tree plots	73.2	73.4	73.2	63.1
single tree plots	74.7	73.1	73.0	-

Standard errors ranged from 0.51 to 0.64.

Table 9. Mean percent of maximum genetic value for six designs and four selection systems on the Hempwallace plantation with 50 trees per family and $h^2 = .2$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
25-tree block plots	54.8	55.9	56.1	56.1
10-tree row plots	61.4	62.0	62.1	59.1
5-tree row plots	61.0	61.9	61.8	58.0
5-tree noncontiguous plots	62.5	63.1	63.2	58.0
2-tree plots	62.2	62.9	62.7	52.0
single tree plots	62.3	62.9	63.1	-

Standard errors ranged from 0.63 to 0.81.

Table 10. Mean percent of maximum genetic value for six designs and four selection systems on the Hempwallace plantations with 50 trees per family and $h^2 = .3$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
25-tree block plots	63.4	64.6	64.7	64.6
10-tree row plots	67.5	67.9	68.1	64.7
5-tree row plots	66.7	67.3	67.1	62.7
5-tree noncontiguous plots	67.0	67.6	67.4	61.4
2-tree plots	66.6	66.6	66.6	56.5
single tree plots	67.6	68.0	68.0	-

Standard errors ranged from 0.54 to 0.80.

Table 11. Mean percent of maximum genetic value for five designs and four selection systems on the Hempwallace plantation with 30 trees per family and $h^2 = .2$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
10-tree row plots	61.0	60.5	61.1	59.4
5-tree row plots	64.4	62.8	64.5	61.6
5-tree noncontiguous plots	64.9	64.5	64.9	61.1
2-tree plots	65.0	64.6	64.8	56.1
single tree plots	65.6	64.8	65.5	-

Standard errors ranged from 0.62 to 0.79.

Table 12. Mean percent of maximum genetic value for five designs and four selection systems on the Hempwallace plantation with 50 trees per family and $h^2 = .3$.

Design	Selection System			
	Rep dev	Comb	N'hood	Plot dev
10-tree row plots	67.9	69.1	69.1	67.9
5-tree row plots	69.8	70.6	70.9	67.2
5-tree noncontiguous plots	69.9	71.1	71.1	66.3
2-tree plots	73.1	72.9	72.8	62.9
single tree plots	71.5	72.1	72.1	-

Standard errors ranged from 0.53 to 0.67.

Table 13. Proportion of the mean genetic value of the twenty selected phenotypes that was due to the individual genetic value relative to the family genetic value for each of the three sites with $h^2=.2$ and 30 trees per family.

Design	Dierks			Cloudy			Hempwallace		
	Rep	Nay	Plot	Rep	Nay	Plot	Rep	Nay	Plot
10-tree row plots	0.61	0.62	0.60	0.40	0.41	0.39	0.43	0.43	0.41
5-tree row plots	0.58	0.59	0.55	0.39	0.41	0.38	0.43	0.42	0.38
noncontiguous plots	0.54	0.55	0.52	0.38	0.39	0.35	0.42	0.42	0.35
2-tree plots	0.55	0.56	0.46	0.39	0.39	0.30	0.43	0.42	0.33
single-tree plots	0.55	0.55	-	0.38	0.39	-	0.43	0.42	-

Table 14. Proportion of the mean genetic value of the twenty selected phenotypes that was due to the individual genetic value relative to the family genetic value for the Cloudy plantation with $h^2=.2$ and $h^2=.3$; 50 trees per family and 30 families.

Design	$h^2=.2$			$h^2=.3$		
	Rep	Nay	Plot	Rep	Nay	Plot
25-tree block plots	0.59	0.59	0.58	0.60	0.60	0.60
10-tree row plots	0.49	0.50	0.46	0.52	0.53	0.48
5-tree row plots	0.50	0.49	0.45	0.53	0.52	0.47
noncontiguous plots	0.49	0.49	0.43	0.53	0.53	0.44
2-tree plots	0.51	0.50	0.40	0.54	0.52	0.44
single-tree plots	0.51	0.50	-	0.54	0.53	-

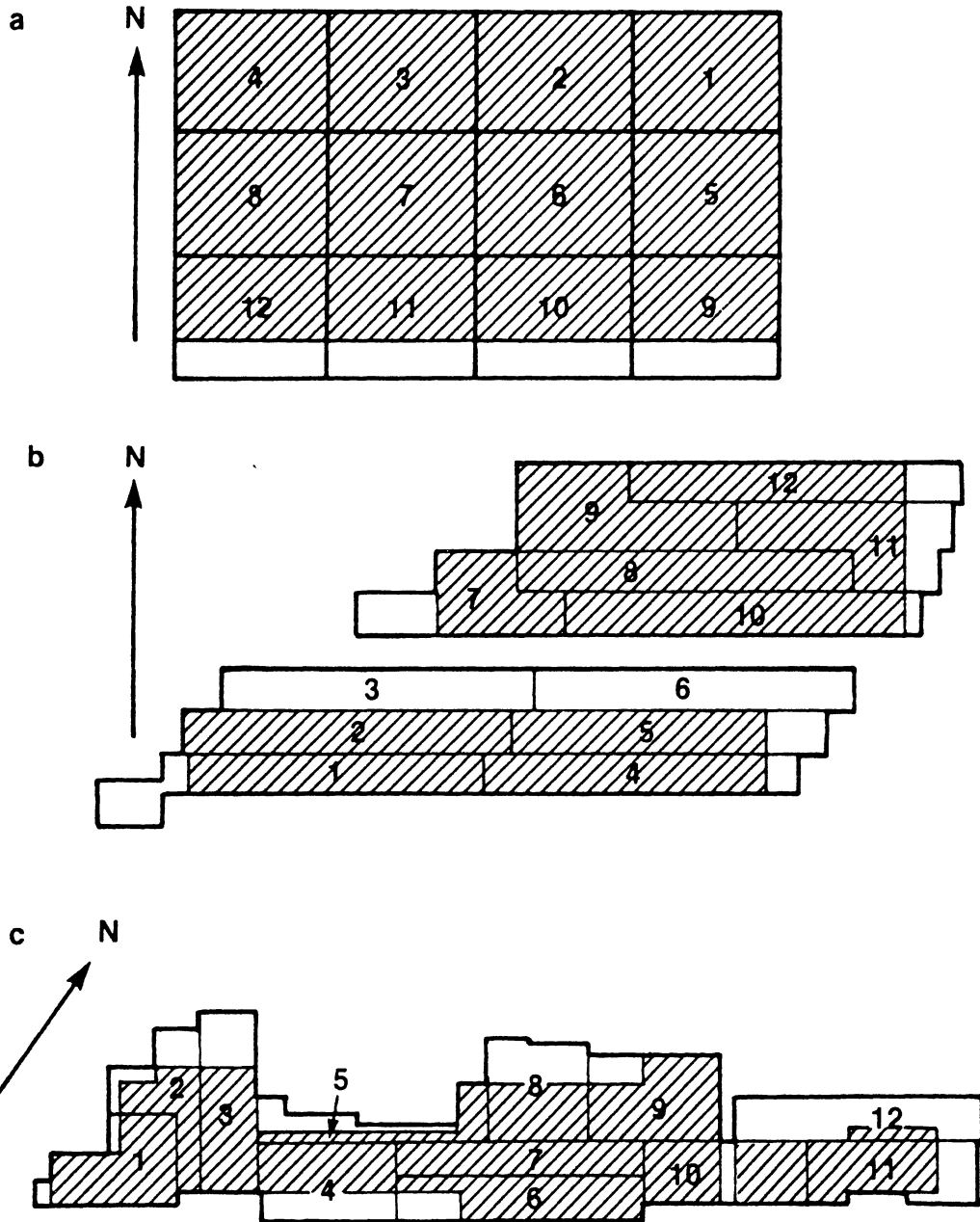


Fig. 1. Plantations used in the simulation model; AOCG7505, located near Dierks, Arkansas (a), AOWF8088, located near Cloudy, Oklahoma (b), and AOWF8085, located near Hempwallace, Arkansas (c).

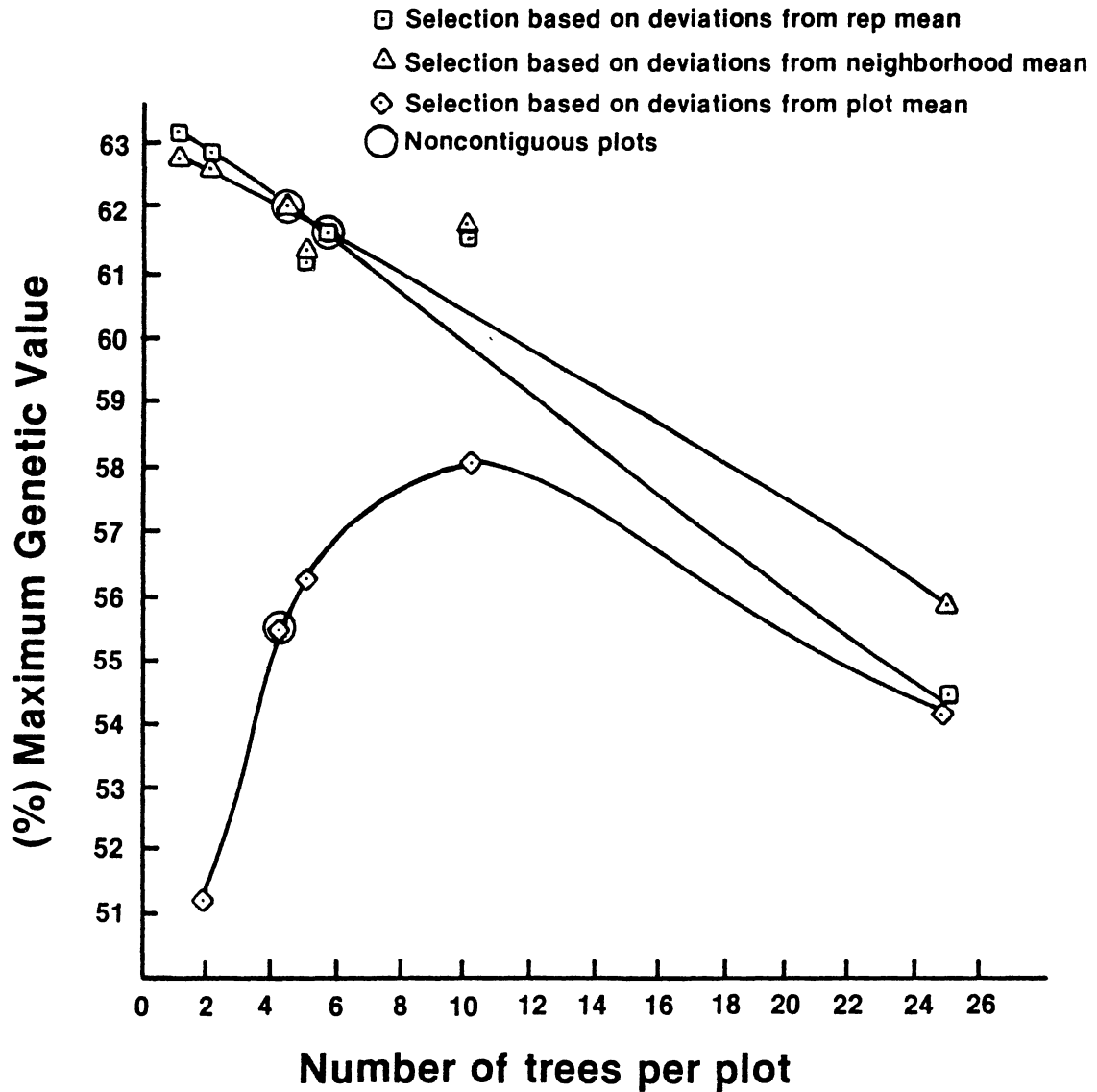


Fig. 2. Mean percentage of maximum genetic values plotted over number of trees per plot for each of three selection systems. Values were averaged over the three sites with a heritability of .20 and 50 or 60 trees per family.

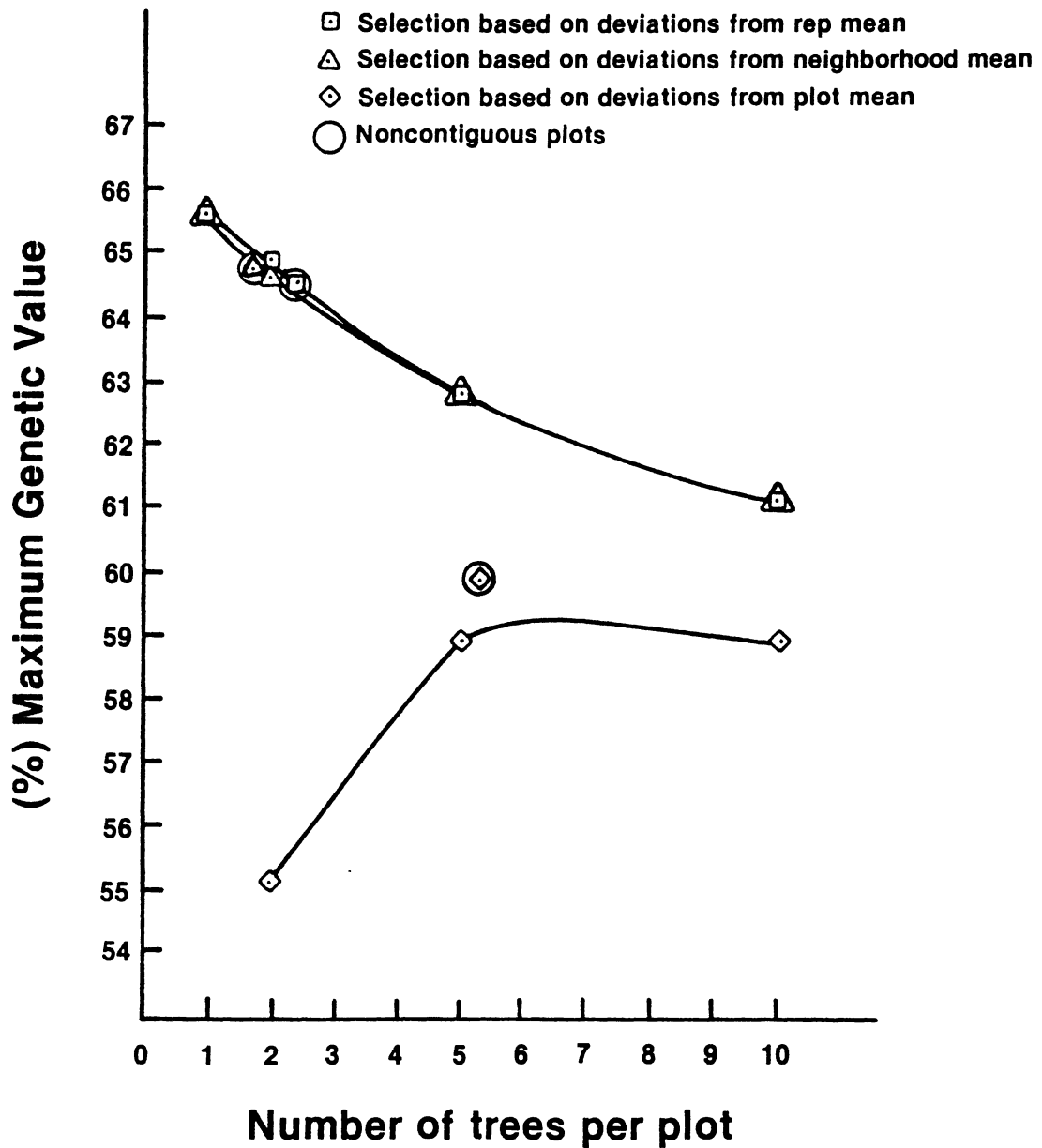


Fig. 3. Mean percentage of maximum genetic values plotted over number of trees per plot for each of three selection systems. Values were averaged over the three sites with a heritability of .20 and 30 trees per family.

PART III

NEIGHBORHOOD COMPARISON AS A BASIS FOR

SELECTION IN LOBLOLLY PINE

(PINUS TAEDA L.)

SUMMARY

Optimizing within-family selection in a progeny test may considerably improve genetic gains. A comparison of potential select trees with unrelated neighbors which occupy an area within the range of significant positive environmental correlations is suggested as an optimal selection method for individual trees within families. A selection index is described which includes neighborhood data in addition to family and individual tree information. Construction of semi-variograms for three loblolly pine (Pinus taeda L.) progeny test sites in southeastern Oklahoma and southwestern Arkansas indicates that the size and shape of the optimum neighborhood area for selection is site specific, depending upon spatial dependency relationships within the site.

Key words: Neighborhood selection, semi-variogram, selection index

INTRODUCTION

Selection of parents for testing and breeding of forest trees increasingly is being carried out in progeny tests. Trees in such tests have at least one known parent, are even-aged, and have already benefited from some degree of genetic improvement. Thus, selection of superior families and of superior individuals within families is an important aspect of progeny testing. The primary purpose of most progeny tests, to evaluate seed orchard trees as potential parents of commercial seed, necessitates locating the tests on sites representative of the range of intended commercial site types. Commercial loblolly pine (Pinus taeda L.) sites in southeastern Oklahoma and southwestern Arkansas often contain a high degree of variability due to factors such as changes in slope and aspect, soil texture, soil moisture, and fertility. Consequently, progeny tests are often located on highly heterogeneous sites.

Some combination of family and within-family information is almost always used in selection of forest trees from progeny tests. The progeny test design determines the efficiency of estimation of family means (Barnes and Schweppenhauser 1979). Even on a highly variable site, families can be accurately ranked provided

an efficient test design is used (Loo-Dinkins, Part I). Heritabilities of growth traits in forest trees are generally low to moderate (Zobel and Talbert 1984); thus, given the same selection intensity, selection at the family level is typically more effective than within-family selection (Falconer 1981). Because of large family size, within-family selection intensity is often considerably higher than the intensity of among family selection. This would allow within-family selections to contribute significantly to genetic gain if an accurate selection method were employed. Thus the selection system that will most accurately identify the best individuals within families is of great importance in a breeding program.

The optimum within-family selection method may depend upon the size and shape of the area in which comparisons are made relative to the type of variation present in the site. Commonly, trees are compared with family plot or replicate means. Particularly on a site with a gradient, an effective family plot should sample the range of microsites within a replicate, but the environmental variability within a comparison area for individual tree selection should be minimized. Because of this conflict and the usual small size of family plots, within-family selection based on deviations from plot means may be inaccurate, particularly in variable environments (Loo-Dinkins, Part II). Progeny tests usually include large numbers of families (often 50 to 75). The most common test

designs consist of replicates of 4- to 6-tree family row plots; thus, replicate size may be large when compared to the area in which positions may be environmentally correlated. Neighborhoods, based on the size of areas where all trees are significantly, positively correlated with each other due to environmental conditions, may constitute the optimal comparison area size.

Gardner (1961) developed a selection method for corn (Zea mays L.) which reduced the environmental variance within a selection area by superimposing a grid on a test field and selecting a set percentage of plants from within each stratum. This allowed greater separation of genetic and environmental effects, thus more effective selection. In loblolly pine progeny tests, the number of selections within a neighborhood would not be expected to be constant. Typically, some percentage of the best families and only one or two individuals within those families are selected to minimize inbreeding. Selection using deviations from a neighborhood mean or making use of environmental correlations in an index may result in gains greater than those attained by other commonly employed selection methods due to better estimates of individual genetic values.

Optimal neighborhoods for selection would be expected to be site specific. The nature of the site variability may be characterized using regionalized variable theory (Burgess and Webster 1980, Vieira et al. 1981). Burgess and Webster (1980) described a function called a semi-variogram

which provides a measure of the average similarity between points separated by a given distance. Semi-variograms are used to describe the spatial dependence relationships within a sampled region. A measure of the average similarity between environmental values associated with each position in a plantation may be plotted over the distance between sets of positions to determine the average size and shape of a region with spatial dependency among environmental values. Application of this theory to the nature of site variability within specific progeny tests could be very useful in determining the optimal size and shape of neighborhoods for selection.

The objective of this study was to characterize the environmental tree to tree variation within three different site types in southeastern Oklahoma and southwestern Arkansas and to develop a selection index which includes environmental correlations as a source of information.

MATERIALS AND METHODS

Three Weyerhaeuser Company loblolly pine progeny tests provided an environmental data base for site characterization. The Dierks plantation, AOGC7505, is a half-sib progeny test located near Dierks, Arkansas. The site was considered the most uniform of the three with no apparent gradient or patchiness. The test is rectangular with 12 replicates, each consisting of a 15 X 18 tree block with 54 families. Eight of the replicates consist of 5-tree row plots. The remaining four replicates contain 5-tree noncontiguous plots. The Cloudy plantation, AOWF8088, is a full-sib progeny test located near Cloudy, in southeastern Oklahoma. This plantation was considered to be the most variable of the three sites. The site straddles a ridge oriented in an east-west direction and six replicates of 34 5-tree row family plots slope downward on each side of the ridge. The Hempwallace plantation, AOWF8085, is a full-sib progeny test located near Hempwallace in southwestern Arkansas. The plantation is on a wet site, causing a pattern of patchy variability. The test consists of 40 families arranged in 12 replicates. Eight of the replicates have a 5-tree row plot arrangement, and the remaining four consist of 5-tree noncontiguous

plots.

Residual values for height were obtained for each position in each plantation by subtracting the plantation mean and the family effect from the height value of each tree. Eight-year data from the Dierks plantation and five-year data from the other two plantations were used. A component of the genetic effect remained in the residual values. However, the environmental effect was assumed to be sufficiently larger than the remaining genetic effect to allow the residual values to primarily reflect the microsite values, as suggested by Franklin (1971). In the case of full-sib families, subtraction of the family effect removes one-half of the additive genetic effect. For example, if individual tree heritability is equal to 0.2, approximately 90% of the variance of the residual values would be due to nonadditive genetic and environmental causes. The residual values became the environmental map for each plantation, and each position was given a row and column number.

Each position, excluding the outer two rows and columns in each plantation, was considered to be the center of a 25-tree square neighborhood. Positions were numbered consecutively from the upper left-hand corner of the neighborhood (Fig. 1) so that position 13 was always the center of the moving neighborhood. The average correlation between each pair of positions within the 25-tree moving neighborhood was estimated for each plantation. This

provided estimates of the environmental correlations between each pair of microsites within a 5-tree proximity in any direction.

A selection index was derived that made use of environmental correlations among neighborhood trees as well as individual tree and family mean data. All trees within a neighborhood were assumed to be unrelated; thus, the index would be appropriate only for single-tree or noncontiguous plot designs. The selection index including neighborhood information was compared with a selection index constructed with individual and family information alone, using a simulation model. Neighborhood size was nine trees. Values were generated to correspond to a heritability of 0.20 to 0.40 and a phenotypic, environmental correlation among neighbors of 0.15 to 0.35. Both indices were applied to the populations with a selection intensity of 0.05.

To further characterize the site variability in each plantation, a semi-variogram was constructed to examine the relationship between spatial dependency and distance between trees. The semi-variance function is:

$$\Gamma(h) = (1/2 m) E \{ [z(x_1) - z(x_1-h)]^2 \}$$

where $\Gamma(h)$ is the semi-variance between pairs of positions, m is the number of pairs of positions included in the summation, $z(x_1)$ is the height residual at position i , and $z(x_1-h)$ is the height residual value at a distance h from position i . The semi-variance reflects the degree to which

two points at a given distance apart are correlated. The relationship between a semi-variance and the correlation between two points may be expressed as:

$$\Gamma(h) = \sigma_p^2 (1 - r_{ij})$$

where σ_p^2 is the phenotypic variance and r_{ij} is the correlation between two points (i and j). A $\Gamma(h)$ value was estimated for pairs of trees at distances of one to nine positions apart within rows and columns for the Hempwallace and Cloudy plantations and at distances of one to 14 positions apart for the Dierks plantation. $\Gamma(h)$ values were plotted over distance both in the direction of rows and in the direction of columns to construct a semi-variogram.

RESULTS AND DISCUSSION

The variance of residual values for height included a portion of the genetic variance; one-half of the additive genetic variance in the case of the full-sib plantations, and three-quarters of the additive genetic variance in the half-sib plantation. The family plots fell within columns. Thus, comparisons of correlation estimates across rows vs. across columns (Table 1) provide an indication of whether error in estimating family effects influenced the correlations among positions. Correlations within family plots (position 1 with positions 6, 11, 16, and 21) are slightly higher than correlations across family plots (position 1 with positions 2, 3, 4, and 5) on the Dierks plantation. The opposite trend is apparent in each of the other plantations. Thus, the assumption that residual values primarily reflect microsite quality appears valid.

The Dierks plantation was considered to be the most uniform with existing variability appearing random. Correlations between positions were lower in general ($r = 0.11$ to 0.35) at this site than at either of the other sites ($r = 0.22$ to 0.45 at Hempwallace and $r = 0.19$ to 0.63 at Cloudy). The Hempwallace plantation appears to have patchy variability, and the size of the patches is probably

larger than the 25-tree neighborhoods. The correlation estimates do not differ greatly between rows and columns, and remain relatively high at a distance of four positions apart. The gradient effect is apparent in the correlation estimates for the Cloudy site. The correlations are high, $r = 0.54$ to 0.63 , between positions 1 to 5 within rows; perpendicular to the gradient. The magnitude of correlations declines rapidly with distance parallel to the gradient, however, $r = 0.55$ to 0.20 , for positions 1 with 2 and 1 with 5, respectively.

Given the magnitude of the correlation coefficients between trees, particularly on the more variable sites, environmental covariances should be useful in the selection process. Increased gains should result from comparison of potential select trees in a progeny test with unrelated neighbors located in positions expected to be significantly positively correlated with the position occupied by the potential selected tree. Selection based on deviations from a neighborhood mean was included in a simulation study to compare various progeny test designs in combination with within family selection methods (Loo-Dinkins, Part II) She found that neighborhood selection was always better than selection based on deviations from a plot mean and was as good as selection based on deviations from replicate means. With large replicate size, selection on the basis of deviations from a neighborhood mean generally performed better than selection based on deviations from replicate

means.

Selection Index

One approach to using neighborhood information is by inclusion of neighborhood covariances as an additional source of information in a selection index. An index was constructed using individual tree value, family mean, and neighborhood mean. The form of the index is:

$$I = P_1b_1 + P_2b_2 + P_3b_3,$$

where I is the index value, P_1 is the individual tree value, P_2 is the family mean, P_3 is the neighborhood mean, and b_1 , b_2 , and b_3 are the weighting coefficients. The coefficients are found by solving the set of equations:

$$b_1P_{11} + b_2P_{12} + b_3P_{13} = A_{11}$$

$$b_1P_{21} + b_2P_{22} + b_3P_{23} = A_{12}$$

$$b_1P_{31} + b_2P_{32} + b_3P_{33} = A_{13}$$

where P_{11} is the phenotypic variance, P_{12} is the phenotypic covariance between individual and family mean, P_{13} is the phenotypic correlation between the individual and the neighborhood mean, $P_{21} = P_{12}$, P_{22} is the variance of a family mean, P_{23} is the covariance between the family and neighborhood means, $P_{31} = P_{13}$, $P_{32} = P_{23}$, P_{33} is the variance of a neighborhood mean, A_{11} is the covariance between the individual's phenotype and its breeding value, A_{12} is the covariance between the individual's breeding value and the family mean, and A_{13} is the covariance between the individual breeding value and the neighborhood mean.

The variance of a family mean can be expressed as:

$$([1+t_f(N-1)]/N)\sigma_p^2.$$

where t_f is the phenotypic correlation among family members, and N is the number of individuals within a family. The additive genetic component of the variance of a family mean is:

$$([1+r(N-1)]/N)\sigma_A^2$$

where r is the correlation among individuals within a family, and σ_A^2 is the additive genetic variance (Falconer 1981). The variance of a neighborhood mean is:

$$([1+t_n(n-1)]/n)\sigma_p^2$$

where n is the number of trees within a neighborhood, and t_n is the phenotypic, environmental correlation among neighbors. Thus:

$$P_{12} = ([1+t_f(N-1)]/N)\sigma_p^2$$

$$P_{13} = ([1+t_n(n-1)]/n)\sigma_p^2$$

$$P_{22} = P_{12}$$

$$P_{23} = ([1+t_n(n-1)+t_f(N-1)]/Nn)\sigma_p^2$$

$$P_{33} = P_{13}$$

$$A_{11} = \sigma_A^2$$

$$A_{12} = ([1+r(N-1)]/N)\sigma_A^2$$

$$A_{13} = (1/n)\sigma_A^2$$

Letting: $[1+t_f(N-1)]/N = F,$

$$[1+t_n(n-1)]/n = H,$$

$$[1+t_n(n-1) + t_f(N-1)]/Nn = C, \text{ and}$$

$$[1+r(N-1)]/N = R,$$

$$b_1 = h^2([FH-C^2-RH(F-C)+(1/N)F(C-H)] / [FH(1+2C-F-H)-C^2])$$

$$b_2 = h^2([-H(F-C)+RH(1-H)+(1/N)(FH-C)] / [FH(1+2C-F-H)-C^2])$$

$$b_3 = h^2([F(C-H)+R(FH-C)+(1/N)F(1-F)] / [FH(1+2C-F-H)-C^2])$$

The coefficients were programmed with various combinations of t_n , t_f , N , n , σ_p^2 , and σ_A^2 to examine the relative magnitudes of the coefficients under different conditions. When trees within neighborhoods are positively correlated, b_3 will always be negative; and if t_n is approximately equal to the heritability (h^2), the magnitude of b_1 is approximately the same as $-b_3$. In comparisons between genotypes selected using an index with and without the neighborhood information, on simulated data with varying h^2 and t_n values, the selection index using neighborhood information was consistently better by approximately 4% when t_n was approximately equal to 0.30 (Table 2). If the average breeding value for height of selected trees using an index which includes only family and individual tree information, is 2.5 feet, representing one-sixth of the total height (15 feet), an increase of 4% due to including neighborhood information, would increase the average genetic value to 2.6. This represents an increase in total height of 1.1%. The heritability was initially varied from 0.20 to 0.40 but it affected both indices in the same way for a given t_n value.

The most serious drawback to adding new parameters to

an index is the additional variance associated with I when more parameters must be estimated. Thus, to be useful, variance introduced by estimation of t_n must be less than the gain that is added due to its inclusion in the index.

Characterization of Site Variability

Results of semi-variogram plots indicate that the size and shape of the region of spatial dependency varies with site (Fig. 2, 3, and 4). The larger the semi-variance the smaller the spatial dependency. The semi-variance would be expected to increase with distance to a value approaching the total environmental variance in the site, then remain constant. The point where a curve becomes level in a semi-variogram is termed the "range". The total environmental variances in the three plantations were 1.04, 0.24, and 0.33 at Dierks, Hempwallace, and Cloudy, respectively. The difference in age of the trees accounts for the much higher environmental variance at the Dierks site. Neither the Dierks nor the Hempwallace semi-variances reached the magnitude of the total environmental variances of the plantations although the variance within columns approached the total variance at a distance of 14 positions in the Dierks site and at nine positions in the Hempwallace site. The semi-variance within columns at a distance of nine positions exceeded the total variance at the Cloudy site. Within rows, however, the semi-variance remained far below the total variance at a distance of nine positions.

The optimum size and shape of a neighborhood for

selection would include positions within a region showing spatial dependency. Thus the semi-variance between any two positions in a neighborhood should be less than the total environmental variance in the plantation. If the semi-variogram plot has a similar shape and semi-variances are similar in magnitude in both directions, circular neighborhoods with a maximum diameter equal to the range of the semi-variogram would be optimal. If square neighborhoods are used, the maximum distance between diagonal corners should equal the range of the semi-variogram. The Hempwallace semi-variogram plot shows similar curves for each direction although rows appear to have slightly smaller semi-variances among positions than columns (Fig. 2). A circular or square neighborhood with a maximum diameter equal to a distance of about eight or nine positions should be optimal.

If semi-variances are consistently smaller in one direction, neighborhoods should be oblong or rectangular with the long axis in the direction with smallest semi-variances. The maximum length of the short axis should not exceed the range of the semi-variogram with largest semi-variances. Optimum neighborhoods in the Dierks plantation would be oriented with the long axis within rows, not exceeding a distance of about 11 positions, and the short axis, in the direction of columns, would not exceed a distance of approximately six positions (Fig. 1).

Optimum neighborhoods indicated by the semi-variogram

plot for the Cloudy site would include a maximum distance of three positions within columns (Fig. 3). The semi-variogram for the within rows direction is almost flat and the semi-variance for an average distance of nine positions apart within rows is less than the average semi-variance between any two positions within columns. Thus a long, narrow neighborhood should be optimal.

CONCLUSIONS

A measure of environmental correlations or spatial dependency within a neighborhood should allow more accurate within family selection on a variable site. The inclusion of neighborhood information in a selection index can increase efficiency relative to an index including only family mean and individual values by approximately 4% when the phenotypic correlation within neighborhoods is approximately 0.30. Among the three study sites, the overall mean correlation within a 25-tree square neighborhood was approximately 0.30. Thus, on the average, with square neighborhoods, the use of neighborhood information in an index would increase the genetic gain slightly. The Cloudy site had considerably higher correlations than the other two sites. Thus, the neighborhood approach may be more valuable in a highly variable test site.

The size and shape of optimal neighborhoods for selection is site specific depending upon the spatial dependence relationships within the site. A semi-variogram is a useful tool in characterizing the nature of site variability in a plantation. It could be used easily to determine the size and shape of an area with sufficiently

high tree-to-tree environmental correlations to merit a neighborhood selection approach.

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Table 1. Average Pearson correlation estimates between pairs of residual values for tree height in moving 25-tree square neighborhoods for each of three plantations.

*Positions correlated	plantation		
	Dierks	Hempwallace	Cloudy
1-2	0.25	0.45	0.63
1-3	0.21	0.40	0.57
1-4	0.20	0.36	0.55
1-5	0.18	0.35	0.54
1-6	0.35	0.38	0.55
1-7	0.22	0.32	0.51
1-8	0.22	0.33	0.47
1-9	0.16	0.35	0.48
1-10	0.15	0.34	0.50
1-11	0.27	0.34	0.37
1-12	0.19	0.36	0.42
1-13	0.21	0.33	0.40
1-14	0.17	0.34	0.37
1-15	0.18	0.30	0.39
1-16	0.25	0.30	0.22
1-17	0.17	0.29	0.22
1-18	0.15	0.27	0.27
1-19	0.13	0.27	0.23
1-20	0.13	0.30	0.23
1-21	0.23	0.22	0.20
1-22	0.18	0.26	0.18
1-23	0.12	0.30	0.19
1-24	0.11	0.29	0.21
1-25	0.11	0.25	0.19

*All correlations are statistically significant at $P=0.01$.

Table 2. Ratio of simulated genetic values selected with a neighborhood index to values selected using only family and individual information. Five levels of neighborhood phenotypic correlations were tested.

Ratio	<u>Neighborhood intraclass correlation</u>				
	<u>0.15</u>	<u>0.20</u>	<u>0.25</u>	<u>0.30</u>	<u>0.35</u>
<u>Index1/Index2</u>	<u>1.01</u>	<u>1.01</u>	<u>1.02</u>	<u>1.04</u>	<u>1.04</u>

* Results are the mean of eighty iterations.

		← columns →			
	1	6	11	16	21
↑	2	7	12	17	22
rows	3	8	13	18	23
↓	4	9	14	19	24
	5	10	15	20	25

Fig. 1. The numbering of positions within the moving neighborhood.

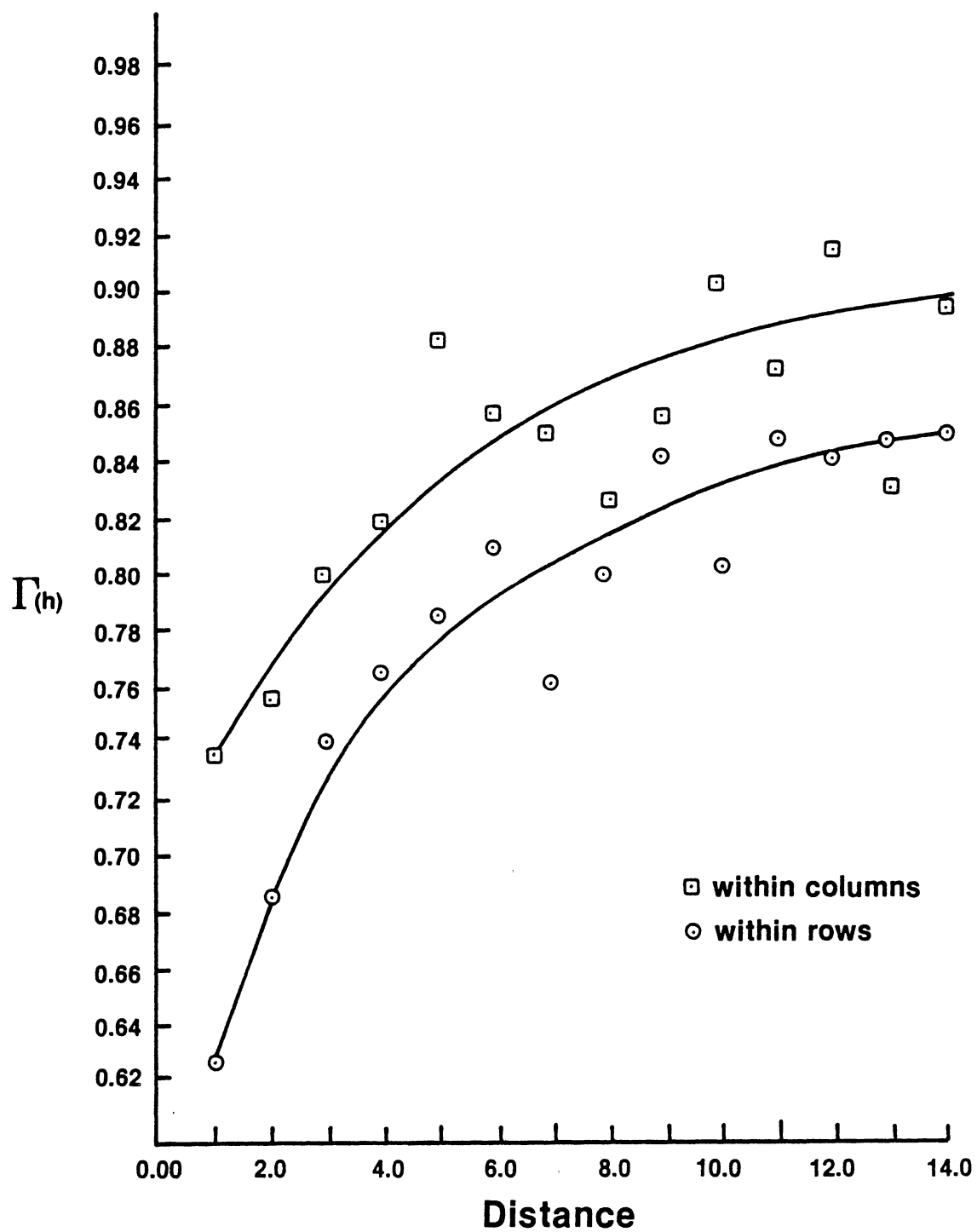


Fig. 2. Semi-variogram for the Dierks plantation in two directions.

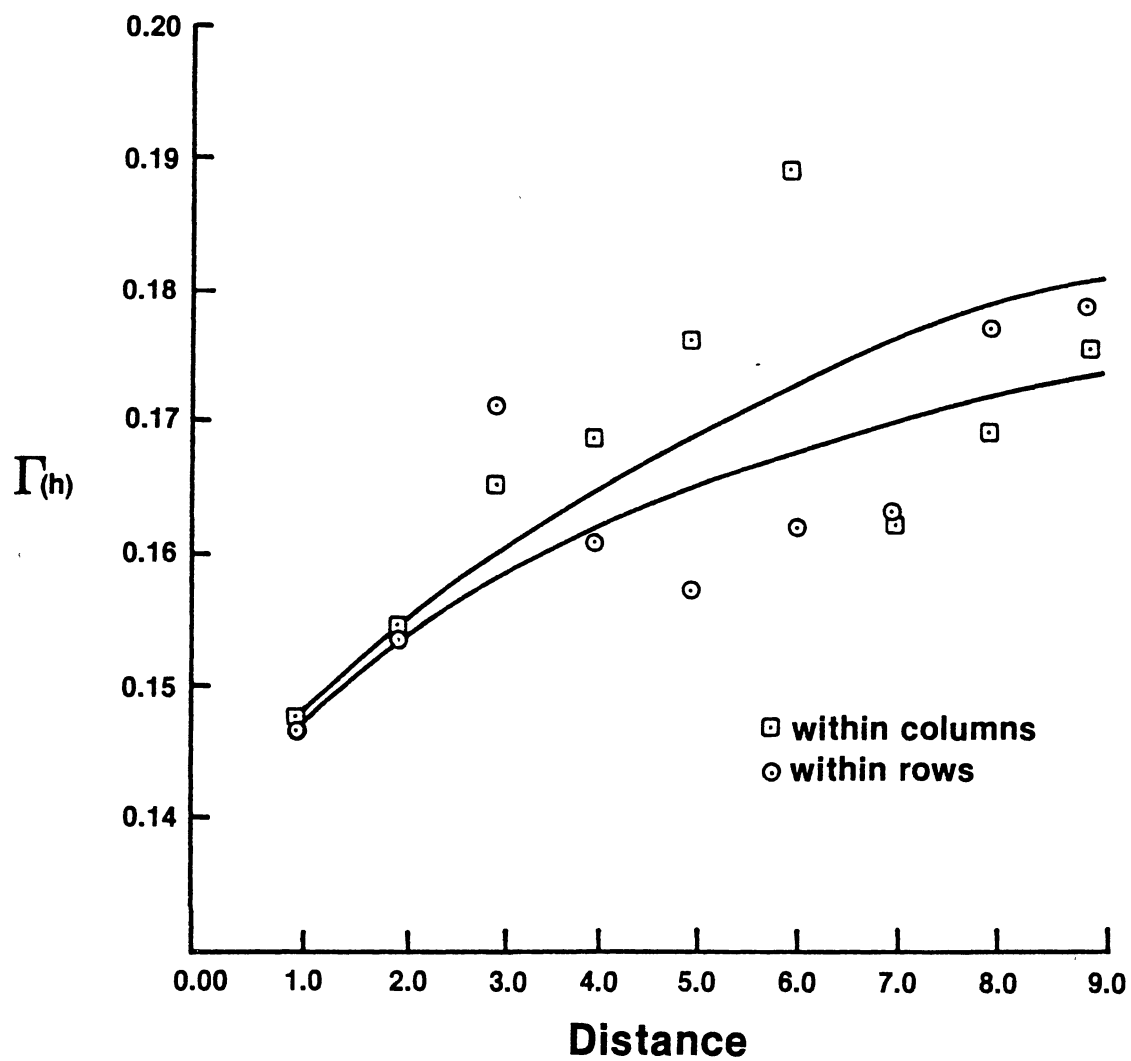


Fig. 3. Semi-variogram for the Hempwallace plantation in two directions.

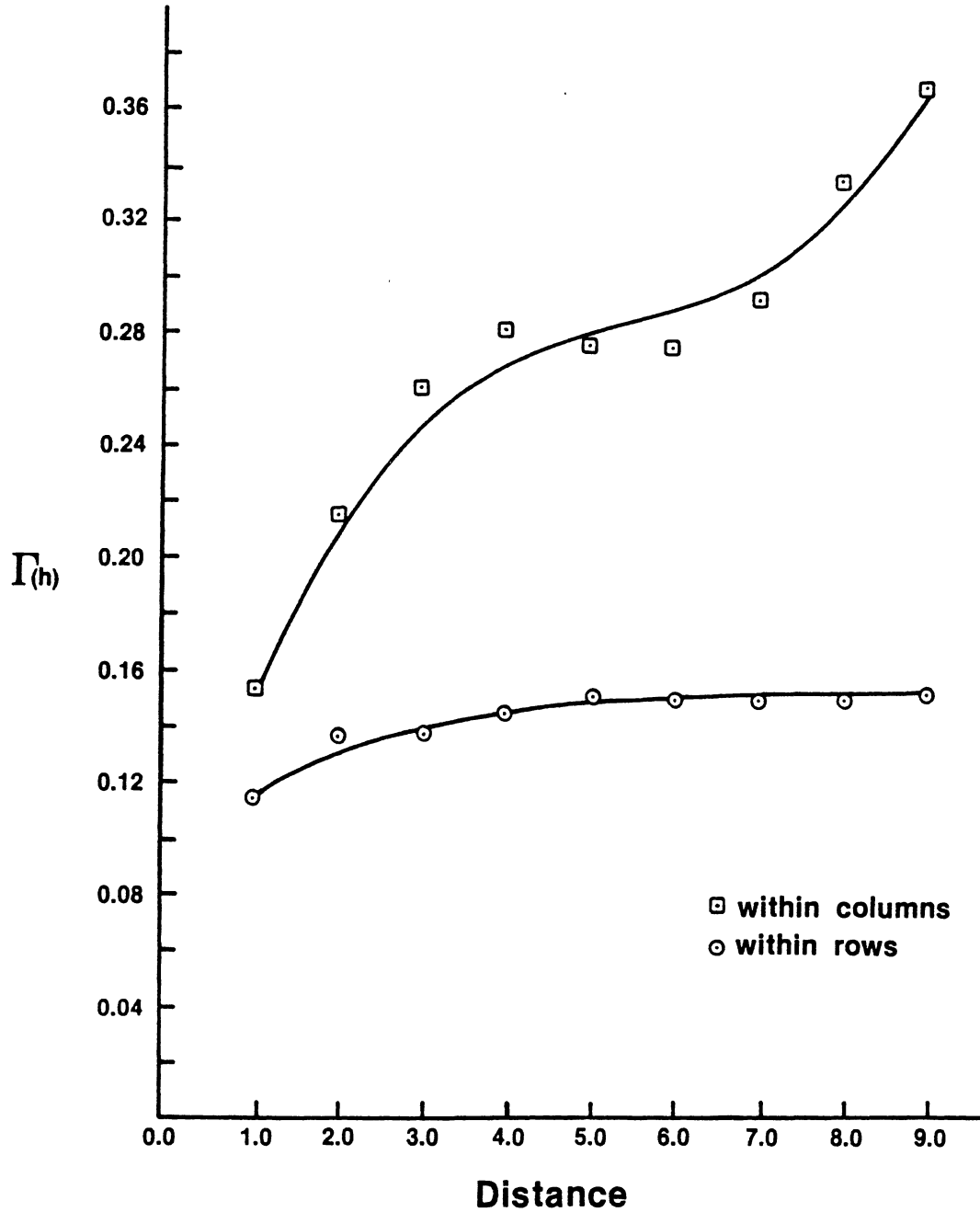


Fig. 4. Semi-variogram for the Cloudy plantation in two directions.

APPENDIX A
 COMPUTER MODEL USED IN DESIGN AND
 SELECTION SYSTEM SIMULATIONS
 (SAS)

```

1      OPTIONS DQUOTE;
2      %MACRO GENET(N, NFAM1, NFAM2, NTREE1, NTREE2, TPOS);
3      %DO I=1 %TO &N;
4      %* N=THE NUMBER OF ITERATIONS;
5      DATA GEN2;
6      *CREATING THE GENETIC POPULATION FOR THE FIRST DESIGN;
7      DO FAM=1 TO &NFAM2;
8      GF=RANNOR(O)*.2537;
9      DO TREE=1 TO &NTREE2;
10     GT=RANNOR(O)*.4394;
11     GEN=GF+GT;
12     OUTPUT;
13     END;
14     END;
15     RUN;
16     PROC SORT DATA=GEN2 OUT=STANDARD;
17     BY DESCENDING GF DESCENDING GT;
18     RUN;
19     DATA ONE;
20     DO F=1 TO 10;
21     DO T=1 TO &NTREE2;
22     OUTPUT; END;
23     END;
24     RUN;
25     DATA STANDARD; MERGE ONE STANDARD;
26     RUN;
27     DATA STANDARD; SET STANDARD;
28     IF T< 3 AND .<F<11;
29     RUN;
30     PROC MEANS NOPRINT DATA=STANDARD;
31     OUTPUT OUT=STANDARD MEAN=BESTGEN;
32     VAR GEN GF GT;
33     RUN;
34     PROC PRINT;
35     TITLE "BEST GENOTYPES FOR ITERATION &I";
36     RUN;
37     %MACRO DESIGN(NDES, NFAM1, NFAM2, TPOS);
38     %DO DES=3 %TO &NDES;
39     %*NDES=THE NUMBER OF DESIGNS TO BE TESTED;
40     DATA FRAME;
41     %IF &DES>1 %THEN %GOTO ST2;
42     %ELSE %IF &DES=1 %THEN %DO;
43     %LET NFAM=&NFAM1;
44     %LET NTREE=&NTREE1;
45     %LET NSEL=%EVAL(&NTREE.*10);
46     %END;
47     %GOTO START;
48     %ST2: %IF &DES>1 %THEN %DO;
49     %LET NFAM=&NFAM2;

```

```

50      %LET NTREE=&NTREE2;
51      %LET NSEL=%EVAL(&NTREE.*10);
52      %END;
53      %START;
54      %IF &DES>1 %THEN %GOTO L0;
55      %ELSE %IF &DES=1 %THEN %DO;
56      %*DESIGN 1=6*10 BLOCK PLOTS;
57          P=60;
58          %LET Q=60;
59          C=40;
60          S1=10;
61          S2=6;
62      %END;
63      %GOTO L7;
64      %L0: %IF &DES>2 %THEN %GOTO L1;
65      %IF &DES=2 %THEN %DO;
66      %*DESIGN 2 = 5*6 BLOCK PLOTS;
67          P=30;
68          %LET Q=30;
69          C=40;
70          S1=5;
71          S2=6;
72      %END;
73      %GOTO L7;
74      %L1: %IF &DES>3 %THEN %GOTO L3;
75      %IF &DES=3 %THEN %DO;
76      %*DESIGN 3 = 5-TREE NONCONTIGUOUS PLOTS;
77          P=5;
78          %LET Q=5;
79          C=40;
80          S1=20;
81          S2=24;
82      %END;
83      %GOTO L7;
84      %L3: %IF &DES>4 %THEN %GOTO L4;
85      %IF &DES=4 %THEN %DO;
86      %*DESIGN 4 = 5 TREE ROW PLOTS;
87          P=5;
88          %LET Q=5;
89          C=40;
90          S1=20;
91          S2=24;
92      %END;
93      %GOTO L7;
94      %L4: %IF &DES>5 %THEN %GOTO L5;
95      %IF &DES=5 %THEN %DO;
96      %*DESIGN 5 = 10 TREE ROW PLOTS;
97          P=10;
98          %LET Q=10;
99          C=40;
100         S1=40;
101         S2=24;
102     %END;
103     %GOTO L7;
104     %L5: %IF &DES>6 %THEN %GOTO L6;
105     %IF &DES=6 %THEN %DO;
106     %* DESIGN 6 = 2 TREE PLOTS;
107         P=2;
108         %LET Q=2;
109         C=40;
110         S1=8;
111         S2=24;
112     %END;
113     %GOTO L7;
114     %*DESIGN 7 = SINGLE TREE PLOTS;
115     %L6: P=1;
116     %LET Q=1;
117     C=40;
118     S1=8;
119     S2=12;
120     %L7:
121     %LET NPOS=%EVAL(&Q.*&NFAM.);
122     %LET NREP=%EVAL(&TPOS./&NPOS.);

```

```

123     QN=&NPOS;
124     NR=&NREP;
125     PN=P*&NFAM;
126     B=1;
127     J=1;
128     K=0;
129     M1=1;
130     POS=0;
131     REP=0;
132     %IF &DES=2 %THEN %DO;
133         M2=S1;
134         COS=C/S1;
135         CN=&NPOS./COS;
136         DO REP=1 TO &NREP;
137         DO V1=B TO CN BY P;
138         DO BL=1 TO COS;
139         U=RANUNI(O)*100;
140         DO V2=J TO S2;
141         DO V3=M1 TO M2;
142         POS=V3;
143         OUTPUT;
144         END;
145         M1=M1+C;
146         M2=M2+C;
147         END;
148         M1=M1+S1-(C*S2);
149         M2=M1-1+S1;
150         END;
151         M1=M1+C*(S2-1);
152         M2=M1+S1-1;
153         END;
154         END;
155     %END;
156     %ELSE %IF &DES=3 %THEN %DO;
157         K=P;
158         M2=S1;
159         COS=C/S1;
160         RR=NR/COS;
161         DO R=1 TO RR;
162         DO V=1 TO COS;
163         REP=REP+1;
164         DO V1=1 TO S2;
165         DO V2=M1 TO M2;
166         U=RANUNI(O)*100;
167         POS=V2;
168         OUTPUT;END;
169         M1=M1+C;
170         M2=M1+S1-1;
171         END;
172         M1=M1+S1-(C*S2);
173         M2=M1+S1-1;
174         END;
175         M1=M1+C*(S2-1);
176         M2=M1+S1-1;
177         END;
178     %END;
179     %ELSE %IF &DES>3 %THEN %DO;
180         K=P;
181         M2=S1;
182         COS=C/S1;
183         RR=NR/COS;
184         DO R=1 TO RR;
185         DO V=1 TO COS;
186         REP=REP+1;
187         DO V1=1 TO S2;
188         DO V2=M1 TO M2 BY P;
189         U=RANUNI(O);
190         DO V3=J TO K;
191         POS=V3;
192         OUTPUT; END;
193         J=J+P;
194         K=K+P;
195         END;

```

```

196      M1=M1+C;
197      M2=M1+S1-1;
198      J=M1;
199      K=J+P-1;
200      END;
201      M1=M1+S1-(C*S2);
202      M2=M1+S1-1;
203      J=M1;
204      K=J+P-1;
205      END;
206      M1=M1+C*(S2-1);
207      M2=M1+S1-1;
208      J=M1;
209      K=J+P-1;
210      END;
211      %END;
212      KEEP POS REP U;
213      RUN;
214      PROC SORT OUT=FRAME;
215      BY REP U;
216      RUN;
217      DATA IDEN;
218      NF=&NFAM2;
219      DO X=1 TO NF;
220      A=1;
221      DO Y=1 TO &Q;
222      ID=0;
223      DO S=A TO &NREP;
224      ID=ID+1;
225      OUTPUT;
226      END;
227      END;
228      END;
229      KEEP ID;
230      RUN;
231      %MACRO IDENT;
232      DATA NEWGEN;
233      MERGE IDEN GEN2;
234      RUN;
235      PROC SORT OUT=PHEN;
236      BY ID FAM;
237      RUN;
238      DATA PHEN;
239      MERGE PHEN FRAME;
240      PROC SORT OUT=PHEN;
241      BY POS;
242      RUN;
243      DATA PHEN;
244      MERGE ORIGIN PHEN;
245      MU=7.5;
246      HT=MU+GEN+HTRES;
247      RUN;
248      PROC CORR DATA=PHEN;
249      VAR GEN HT;
250      TITLE "CORRELATION FOR ITERATION &I AND DESIGN &DES H2=.2";
251      RUN;
252      PROC MEANS NOPRINT DATA=PHEN;
253      OUTPUT OUT=TEMP MEAN=HTMEAN;
254      VAR HT;
255      RUN;
256      DATA TEMP; SET TEMP;
257      DO T= 1 TO &TPOS;
258      HTMEAN=HTMEAN;
259      OUTPUT; END;
260      RUN;
261      DATA PHEN; MERGE PHEN TEMP;
262      KEEP REP FAM GEN GF GT HTRES HT HTMEAN POS ROW COL NAY;
263      RUN;
264      PROC PRINT DATA=PHEN;
265      TITLE "DESIGN &DES DATA SET";
266      RUN;
267      PROC SORT DATA=PHEN;
268      BY REP;

```

```

269     PROC MEANS NOPRINT DATA=PHEN;
270     OUTPUT OUT=TEMP MEAN=REPMEAN STD=REPSTD;
271     BY REP;
272     VAR HT;
273     RUN;
274     DATA TEMP; SET TEMP;
275     DO T=1 TO &NPOS;
276     REPMEAN=REPMEAN;
277     REPSTD=REPSTD;
278     STANDARD=REPMEAN+1.5*REPSTD;
279     OUTPUT;
280     END;
281     RUN;
282     DATA PHEN; MERGE PHEN TEMP;
283     FIRSTSEL=HT-STANDARD;
284     REPDEV=HT-REPMEAN;
285     KEEP REP FAM GEN GF GT HT HTMEAN NAY REPMEAN REPDEV STANDARD FIRSTSEL;
286     RUN;
287     PROC SORT DATA=PHEN OUT=PHEN;
288     BY FAM;
289     RUN;
290     PROC MEANS NOPRINT DATA=PHEN;
291     OUTPUT OUT=NCSEL MEAN=NCMN FAMMEAN;
292     BY FAM;
293     VAR FIRSTSEL HT;
294     RUN;
295     DATA NCSEL; SET NCSEL;
296     DO TRE=1 TO &NTREE;
297     NCMN=NCMN;
298     FAMMEAN=FAMMEAN;
299     OUTPUT; END;
300     RUN;
301     DATA PHEN; MERGE PHEN NCSEL;
302     RUN;
303     PROC SORT DATA=PHEN OUT=NCSEL;
304     BY DESCENDING NCMN DESCENDING FIRSTSEL;
305     RUN;
306     DATA NCSEL; SET NCSEL(OBS=&NSEL);
307     RUN;
308     DATA TEMP;
309     DO F=1 TO 10;
310     DO T=1 TO &NTREE; OUTPUT;
311     END; END; RUN;
312     DATA NCSEL; MERGE NCSEL TEMP;
313     RUN;
314     DATA NCSEL; SET NCSEL;
315     IF T<3 AND FIRSTSEL>0;
316     RUN;
317     PROC MEANS NOPRINT DATA=NCSEL;
318     OUTPUT OUT=NCSEL MEAN=NCSELMN GFMN GTMN;
319     VAR GEN GF GT;
320     RUN;
321     DATA COMB; MERGE NCSEL STANDARD;
322     PERCENT=(NCSELMN/BESTGEN)*100;
323     RUN;
324     PROC PRINT DATA=COMB;
325     TITLE 'SELECTIONS USING MODIFIED NC STATE METHOD';
326     RUN;
327     DATA GENCOR; SET PHEN;
328     IF TRE=1;
329     RUN;
330     PROC CORR;
331     VAR GF FAMMEAN;
332     TITLE 'CORRELATION BETWEEN FAMMEAN AND GF';
333     RUN;
334     PROC SORT DATA=PHEN OUT=PHEN;
335     BY NAY;
336     RUN;
337     PROC MEANS NOPRINT DATA=PHEN;
338     OUTPUT OUT=TEMP MEAN=NAYMN;
339     BY NAY;

```

```

340     VAR HT;
341     RUN;
342     DATA TEMP; SET TEMP;
343     IF NAY<113 THEN DO A=1 TO 25;
344     NAYMN=NAYMN;
345     OUTPUT; END;
346     ELSE IF NAY>112 THEN DO A=1 TO 20;
347     NAYMN=NAYMN;
348     OUTPUT; END;
349     RUN;
350     DATA PHEN; MERGE PHEN TEMP;
351     NAYDEV=HT-NAYMN;
352     RUN;
353     PROC SORT DATA=PHEN OUT=PHEN;
354     BY DESCENDING FAMMEAN;
355     RUN;
356     DATA PHEN; SET PHEN(OBS=&NSEL);
357     RUN;
358     %IF &DES=7 %THEN %GOTO T2;
359     %ELSE %IF &DES<7 %THEN %DO;
360     PROC SORT DATA=PHEN OUT=TEMP;
361     BY REP FAM;
362     RUN;
363     PROC MEANS NOPRINT DATA=TEMP;
364     OUTPUT OUT=TEMP MEAN=PLOTMEAN;
365     BY REP FAM;
366     VAR HT;
367     RUN;
368     DATA TEMP; SET TEMP;
369     P=%EVAL(&NTREE./&NREP.);
370     DO T=1 TO P;
371     OUTPUT; END;
372     RUN;
373     DATA PHEN; MERGE TEMP PHEN;
374     PLOTDEV=HT-PLOTMEAN;
375     RUN;
376     %END;
377     %T2:
378     PROC SORT DATA=PHEN;
379     BY FAM DESCENDING REPDEV;
380     RUN;
381     DATA MASSEL;
382     DO F=1 TO 10;
383     DO T=1 TO &NTREE;
384     OUTPUT; END;
385     END;
386     RUN;
387     DATA REPNAY; MERGE PHEN MASSEL;
388     IF T>10 THEN DELETE;
389     RUN;
390     PROC SORT DATA=REPNAY OUT=REPNAY;
391     BY FAM DESCENDING NAYDEV;
392     RUN;
393     DATA TEMP;
394     DO F=1 TO 10;
395     DO T=1 TO 10;
396     OUTPUT; END; END;
397     RUN;
398     DATA REPNAY; MERGE REPNAY TEMP;
399     IF T>2 THEN DELETE;
400     RUN;
401     PROC MEANS NOPRINT DATA=REPNAY;
402     OUTPUT OUT=REPNAY MEAN=COMBMN GFMN GTMN;
403     VAR GEN GF GT;
404     RUN;
405     DATA COMB; MERGE REPNAY STANDARD;
406     PERCENT=(COMBMN/BESTGEN)*100;
407     RUN;
408     PROC PRINT DATA=COMB;
409     TITLE 'MEAN GENETIC VALUE OF COMBINED REP AND NABORHOOD SELECTIONS';
410     RUN;
411     DATA MASSEL; MERGE PHEN MASSEL;
412     IF T>2 THEN DELETE;

```

```

413      RUN;
414      PROC MEANS NOPRINT DATA=MASSEL;
415      OUTPUT OUT=MASSEL MEAN=GENMASS GFMN GTMN;
416      VAR GEN GF GT;
417      RUN;
418      DATA COMB; MERGE MASSEL STANDARD;
419      PERCENT=(GENMASS/BESTGEN)*100;
420      RUN;
421      PROC PRINT DATA=COMB;
422      TITLE 'MEAN GENETIC VALUE OF REP DEVIATION SELECTIONS';
423      RUN;
424      DATA PHEN; SET PHEN;
425      PROC SORT DATA=PHEN OUT=PHEN;
426      BY FAM DESCENDING NAYDEV;
427      RUN;
428      DATA NAYSEL;
429      DO F=1 TO 10;
430      DO N=1 TO &NTREE;
431      OUTPUT; END;
432      END;
433      RUN;
434      DATA NAYSEL; MERGE NAYSEL PHEN;
435      IF N>2 THEN DELETE;
436      RUN;
437      PROC MEANS NOPRINT DATA=NAYSEL;
438      OUTPUT OUT=NAYSEL MEAN=NAYGEN GFMN GTMN;
439      VAR GEN GF GT;
440      RUN;
441      DATA COMB; MERGE NAYSEL STANDARD;
442      PERCENT=(NAYGEN/BESTGEN)*100;
443      RUN;
444      PROC PRINT DATA=COMB;
445      TITLE 'MEAN GENETIC VALUE OF NEIGHBORHOOD SELECTIONS';
446      RUN;
447      %IF &DES=7 %THEN %GOTO T4;
448      %ELSE %IF &DES<7 %THEN %DO;
449      DATA PHEN; SET PHEN;
450      PROC SORT DATA=PHEN OUT=PHEN;
451      BY FAM DESCENDING PLOTDEV;
452      RUN;
453      DATA PLOTSEL;
454      DO F=1 TO 10;
455      DO W=1 TO &NTREE;
456      OUTPUT; END;
457      END;
458      RUN;
459      DATA PLOTSEL;
460      MERGE PLOTSEL PHEN;
461      IF W>2 THEN DELETE;
462      RUN;
463      PROC MEANS NOPRINT DATA=PLOTSEL;
464      OUTPUT OUT=PLOTSEL MEAN=PLOTGEN GFMN GTMN;
465      VAR GEN GF GT;
466      RUN;
467      DATA COMB; MERGE PLOTSEL STANDARD;
468      PERCENT=(PLOTGEN/BESTGEN)*100;
469      RUN;
470      PROC PRINT DATA=COMB;
471      TITLE 'MEAN GENETIC VALUE OF PLOT DEVIATION SELECTIONS';
472      RUN;
473      %END;
474      %T4;
475      %IF &DES=3 OR &DES=6 OR &DES=7 %THEN %DO;
476      DATA NAYINDEX; SET PHEN;
477      NF=&NFAM;
478      H2=.4;
479      TN=.25;
480      NNAY=25;
481      CORR=.25;
482      TF=H2*CORR;
483      F=(1+TF*(NF-1))/NF;
484      H=(1+TN*(NNAY-1))/NNAY;
485      C=(1+TF*(NF-1)+CORR*(NNAY-1))/(NF*NNAY);

```

```

486      R=(1+CORR*(NF-1))/NF;
487      D=1/NNAY;
488      DET=F*H*(1+2*C-F-H)-C*C;
489      B1=H2*(F*H-C*C-R*H*(F-C)+F*D*(C-H))/DET;
490      B2=H2*(R*H*(1-H)-H*(F-C)+D*(F*H-C))/DET;
491      B3=H2*(F*(C-H)+R*(F*H-C)+D*F*(1-F))/DET;
492      INDEX=HT*B1+FAMMEAN*B2+NAYMN*B3;
493      RUN;
494      PROC SORT DATA=NAYINDEX OUT=NAYINDEX;
495      BY FAM DESCENDING INDEX;
496      RUN;
497      DATA INDEX;
498      DO F=1 TO 10;
499      DO T=1 TO &NTREE;
500      OUTPUT; END; END;
501      RUN;
502      DATA NAYINDEX; MERGE NAYINDEX INDEX;
503      IF T>2 THEN DELETE;
504      RUN;
505      PROC MEANS NOPRINT DATA=NAYINDEX;
506      OUTPUT OUT=NAYINDEX MEAN=INDSEL2 GFMN GTMN;
507      VAR GEN GF GT;
508      RUN;
509      DATA COMB; MERGE NAYINDEX STANDARD;
510      PERCENT=(INDSEL2/BESTGEN)*100;
511      RUN;
512      PROC PRINT DATA=COMB;
513      TITLE 'MEAN GENETIC VALUE OF SELECTIONS USING NEIGHBORHOOD INDEX';
514      RUN;
515      %END;
516      %MEND IDENT;
517      %IDENT
518      %END;
519      %MEND DESIGN;
520      %DESIGN(7, 1, 96, 2880)
521      %END;
522      %MEND GENET;
523      %MACRO BIG;
524      DATA ORIGIN;
525      INFILE HD;
526      INPUT COL 1-3 ROW 5-6 HTRES 8-14 4 NAY 16-18 POS 20-24;
527      %GENET(1, 1, 96, 1, 30, 2880)
528      %MEND BIG;

```


APPENDIX B
COMPUTER MODEL USED IN SELECTION
INDEX SIMULATIONS (SAS)

```

1      %MACRO NABOR(NRUN);
2      %LET TOTVAR=10;
3      %LET H2=.10;
4      %LET NTREE=45;
5      %LET NNAY=9;
6      %LET TN=.10;
7      %LET COR=.5;
8      %DO SIM=1 %TO 4;
9      %DO HERIT=1 %TO &NRUN;
10     %LET H2=&H2+.05;
11     %DO NAYCOR=1 %TO &NRUN;
12     %LET TN=&TN+.05;
13
14     DATA ONE;
15     SEED=0;
16     NF=&NTREE;
17     PHENVAR=&TOTVAR;
18     HRTBLY=&H2;
19     CORR=&COR;
20     FAMVAR=HRTBLY*PHENVAR;
21     FAMSTD=SQRT(FAMVAR);
22     MU=15;
23     DO FAM=1 TO 20;
24     GF=FAMSTD*RANNOR(SEED);
25     DO TREE=1 TO NF;
26     GT=FAMSTD*RANNOR(SEED);
27     GEN=GF+GT;
28     U=RANUNI(SEED);
29     OUTPUT; END;
30     END;
31     PROC SORT DATA=ONE OUT=THREE;
32     BY U;
33
34     DATA TWO;
35     SEED=0;
36     HRTBLY=&H2;
37     NAYSIZE=&NNAY;
38     NAYCOR=&TN;
39     PHENVAR=&TOTVAR;
40     ENVAR=PHENVAR-(HRTBLY*PHENVAR);
41     AMGVAR=NAYCOR*ENVAR;
42     AMGNAY=SQRT(AMGVAR);
43     WNAV=SQRT(ENVAR-AMGVAR);
44     DO NAYBER=1 TO 100;
45     NV=AMGNAY*RANNOR(SEED);
46     DO IND=1 TO NAYSIZE;
47     EN=WNAV*RANNOR(SEED);
48     OUTPUT; END;
49     END;
50
51     DATA FOUR; MERGE THREE TWO;
52     PHEN=MU+GF+GT+NV+EN;
53     PROC MEANS NOPRINT DATA=FOUR;

```

```

54      OUTPUT OUT=FIVE MEAN=NAYMN;
55      BY NAYBER; VAR PHEN;
56      DATA SIX; SET FIVE;
57      DO N=1 TO 9;
58          NHMN=NAYMN;
59      OUTPUT; END;
60      DATA SEVEN; MERGE FOUR SIX;
61      PROC SORT DATA=SEVEN OUT=EIGHT;
62      BY FAM;
63      PROC MEANS NOPRINT DATA=EIGHT;
64      OUTPUT OUT=NINE MEAN=FMN;
65      BY FAM; VAR PHEN;
66      DATA TEN; SET NINE;
67      DO T=1 TO 45;
68          FAMMN=FMN;
69      OUTPUT; END;
70
71      DATA ELEVEN; MERGE EIGHT TEN;
72      TF=HRTBLY*CORR;
73      F=(1+TF*(NF-1))/NF;
74      H=(1+NAYCOR*(NAYSIZ-1))/NAYSIZ;
75      C=(1+TF*(NF-1)+NAYCOR*(NAYSIZ-1))/(NF*NAYSIZ);
76      R=(1+CORR*(NF-1))/NF;
77      D=1/NAYSIZ;
78      C1=HRTBLY*(1-R)/(1-F);
79      C2=HRTBLY*(R-F)/(F*(1-F));
80      DET=F*H*(1+2*C-F-H)-C*C;
81      B1=HRTBLY*(F*H-C*C-R*H*(F-C)+F*D*(C-H))/DET;
82      B2=HRTBLY*(R*H*(1-H)-H*(F-C)+D*(F*H-C))/DET;
83      B3=HRTBLY*(F*(C-H)+R*(F*H-C)+D*F*(1-F))/DET;
84      INDEX1=PHEN*C1+FAMMN*C2;
85      INDEX2=PHEN*B1+FAMMN*B2+NHMN*B3;
86      DIFF1=(MU+GF+GT)-INDEX1;
87      DIFF2=(MU+GF+GT)-INDEX2;
88      PROC CORR DATA=ELEVEN;
89      VAR INDEX1 GEN;
90      PROC CORR DATA=ELEVEN;
91      VAR INDEX2 GEN;
92      PROC SORT DATA=ELEVEN OUT=TWELVE;
93      BY DESCENDING INDEX1;
94      PROC SORT DATA=ELEVEN OUT=THTEEN;
95      BY DESCENDING INDEX2;
96      DATA FRTEEN;
97      DO SEL=1 TO 45;
98      OUTPUT; END;
99      DATA FIFTEEN; MERGE TWELVE FRTEEN;
100     IF SEL<1 THEN DELETE;
101     KEEP HRTBLY FAM GEN NAYBER NV EN PHEN FAMMN NAYMN C1 C2
102     B1 B2 B3 INDEX1 INDEX2 DIFF1 DIFF2;
103     PROC PRINT DATA=FIFTEEN;
104     TITLE SELECTIONS USING INDEX1;
105     PROC MEANS NOPRINT DATA=FIFTEEN;
106     OUTPUT OUT=SIXTEEN MEAN=GENMN PHENMN;
107     VAR GEN PHEN;
108     DATA SVNTEEN; MERGE THTEEN FRTEEN;
109     IF SEL<1 THEN DELETE;
110     KEEP HRTBLY FAM GEN NAYBER NV EN PHEN FAMMN NAYMN
111     C1 C2 B1 B2 B3 INDEX1 INDEX2 DIFF1 DIFF2;
112     PROC PRINT DATA=SVNTEEN;
113     TITLE SELECTIONS USING INDEX2;
114     PROC MEANS NOPRINT DATA=SVNTEEN;
115     OUTPUT OUT=EITEEN MEAN=GENMN PHENMN;
116     VAR GEN PHEN;
117     PROC PRINT DATA=SIXTEEN;
118     TITLE MEAN VALUE OF SELECTED TREES: FAMILY AND INDIV;
119     PROC PRINT DATA=EITEEN;
120     TITLE MEAN VALUE OF SELECTED TREES WITH NEIBORHOODS;
121     PROC MEANS DATA=FOUR;
122     OUTPUT OUT=NINETEEN MEAN=GENMN PHENMN;
123     VAR GEN PHEN;
124     PROC PRINT DATA=NINETEEN;
125     TITLE OVERALL MEANS FOR RUN=&HERIT,&NAYCOR;

```

```
126      %END;  
127      %LET TN= .1;  
128      %END;  
129      %LET H2= .15;  
130      %END;  
131      %MEND NABOR;  
131
```

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VITA

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Candidate for the Degree of

Doctor of Philosophy

Thesis: COMPARISONS AMONG COMBINATIONS OF PROGENY TEST
DESIGNS AND SELECTION SYSTEMS FOR MAXIMUM GENETIC
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Minor Field: Statistics

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