



Influence of Transcranial Direct Current Stimulation Dosage and Associated Therapy on Motor Recovery Post-stroke: A Systematic Review and Meta-Analysis

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Purpose: (1) To determine the impact of transcranial direct current stimulation (tDCS) applied alone or combined with other therapies on the recovery of motor function after stroke and (2) To determine tDCS dosage effect.

Methods: Randomized controlled trials comparing the effects of tDCS with sham, using the Barthel Index (BI), the upper and lower extremity Fugl–Meyer Assessment (FMA), and the Modified Ashworth Scale (MAS), were retrieved from PubMed, Medline (EBSCO), and Cumulative Index to Nursing and Allied Health Literature (CINAHL) from their inception to June 2021. Calculations for each assessment were done for the overall effect and associated therapy accounting for the influence of stroke severity or stimulation parameters.

Results: A total of 31 studies involving metrics of the BI, the upper extremity FMA, the lower extremity FMA, and the MAS were included. tDCS combined with other therapies was beneficial when assessed by the BI (mean difference: 6.8; P < 0.01) and these studies typically had participants in the acute stage. tDCS effects on the upper and lower extremity FMA are unclear and differences between the sham and tDCS groups as well as differences in the associated therapy type combined with tDCS potentially influenced the FMA results. tDCS was not effective compared to sham for the MAS. Stimulation types (e.g., anodal vs. cathodal) did not influence these results and dosage parameters were not associated with the obtained effect sizes. Conventional therapy associated with tDCS typically produced greater effect size than assisted therapy. The influence of stroke severity is unclear.

Conclusion: Potential benefits of tDCS can vary depending on assessment tool used, duration of stroke, and associated therapy. Mechanistic studies

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are needed to understand the potential role of stimulation type and dosage effect after stroke. Future studies should carefully conduct group randomization, control for duration of stroke, and report different motor recovery assessments types.

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Keywords: brain stimulation, tDCS, stroke, Barthel Index, Fugl-Meyer, Ashworth Scale, rehabiliatation

INTRODUCTION

In the United States alone, it has been estimated that every 40 s, someone suffers from a stroke, averaging more than 795,000 incidents of stroke every year with around 185,000 of those stroke occurrences happening in people who have already suffered from a stroke (CDC, 2021). Stroke is also regarded as one of the leading causes for disability, leading to reduced motor function, which limits participation in normal activities of daily life, such as locomotion, dressing, or eating (Kim et al., 2014; Hatem et al., 2016). This reduction in daily activities of life and physical activity due to disability further increases the affected person's risk for further cardiovascular disease, which may lead to a subsequent stroke (Adeyemo et al., 2012). Stroke frequently leads to significant alterations in cortical excitability of the primary motor cortex in the affected and unaffected hemispheres, which lead to the idea that manipulating the cortical excitability may have an influence on stroke recovery (Hummel and Cohen, 2006).

Transcranial direct current stimulation (tDCS) is a noninvasive brain stimulation technique that involves applying a small current to the scalp aiming to modulate cortical excitability (Nitsche and Paulus, 2000; Bikson et al., 2016; Orrù et al., 2019). A common way of applying tDCS is based on the interhemispheric competition model aiming to reduce interhemispheric inhibition and increase excitability of lesioned hemisphere. Typical configurations of tDCS are: (1) the anode electrode placed over the brain area of interest aiming to increase excitation and the cathode electrode placed as a reference (i.e., anodal stimulation); (2) cathode electrode placed over the contralesional hemisphere, aiming to decrease excitability, and the anode electrode placed as reference such as the ipsilesional supraorbital region (i.e., cathodal stimulation); and (3) bihemispheric stimulation aiming to both decrease contralesional and increase ipsilesional excitability. However, applying tDCS in accordance to the interhemispheric competition model is currently under debate considering reports challenging the influence of interhemispheric imbalance on motor recovery (Di Pino et al., 2014; Xu et al., 2019).

Earlier meta-analysis evaluating upper limb motor function suggested that tDCS could be beneficial for individuals with chronic stroke (Butler et al., 2013; Chhatbar et al., 2015). However, a more recent comprehensive systematic review is inconclusive on the effects of tDCS on several aspects of physical function (Elsner et al., 2020), whereas others suggest that tDCS could be beneficial for the upper limb motor function (Van Hoornweder et al., 2021). Several factors are suggested to influence the tDCS results such as tDCS dosage, severity of disease, and type of associated therapy (Chhatbar et al., 2015; Elsner et al., 2020; Van Hoornweder et al., 2021). Thus, further analysis is still necessary on the effects of associated therapy used with tDCS, dosage effect, and severity of disease in multiple domains of motor function recovery. Minimal clinically important differences (MCIDs) were also not compiled in the previous reviews and they are another relevant information. Thus, the purpose of this systematic review is two-fold: (1) to determine the influence of tDCS alone or combined with other therapies on the recovery of motor function after stroke and (2) to determine the influence of therapy type, stimulation configuration (e.g., anodal vs. cathodal), and tDCS dosage (i.e., current, duration, electrode size, session number, and frequency) on the potential benefits of tDCS. We hypothesized that tDCS combined with other therapies would be beneficial for stroke recovery and the results will be dependent on the assessment used, dosage, severity, and the type of combined therapy.

MATERIALS AND METHODS

We followed the recommendations of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) checklist (Page et al., 2021) for study retrieval and subsequent analysis.

Search Strategy

The online databases that were searched are PubMed, Medline (EBSCO), and Cumulative Index to Nursing and Allied Health Literature (CINAHL). In each of these databases, we used a combination of terms "stroke," "tDCS or transcranial direct current stimulation," and "Fugl–Meyer or Ashworth or Barthel" to locate relevant articles. The most recent search of PubMed, Medline (EBSCO), and CINAHL using the combined terms produced 45, 89, and 34 results, respectively.

Inclusion and Exclusion Criteria

Randomized controlled trials assessing the effects of tDCS compared to a sham intervention were included. For the sham, typically tDCS is turned off within the first min to simulate the itching sensation from the beginning of stimulation. The trials contained results for one of or a combination of the Barthel Index (BI), the upper extremity Fugl–Meyer Assessment (FMA), the lower extremity FMA, and the Modified Ashworth

tDCS and Stroke

Scale (MAS) were included in this study. Studies used either anodal, cathodal, or bihemispheric tDCS to rehabilitate patients after suffering from either a hemorrhagic or ischemic stroke. As this review focuses on the effects of tDCS on stroke, any study, which includes the use of brain stimulation aside from tDCS such as, transcranial random noise stimulation, was excluded. Additionally, the use of combined stimulation techniques, such as the use of tDCS in combination with repetitive tDCS, was excluded.

Data Extraction

One of the authors (A-MC) extracted the following variables from each study: (1) mean score and SD before and after treatment for each assessment used to measure functional recovery alone (i.e., the BI, the FMA, and the MAS). The first data point after treatment was used in case multiple follow-ups were reported; (2) number of intervention sessions in which tDCS was applied; (3) application time of tDCS during each session; (4) total time of tDCS application during all the testing sessions; (5) current; (6) electrode size; (7) current density of tDCS; (8) charge of tDCS; (9) charge density of tDCS; (10) total charge of tDCS; (11) total charge density of tDCS; (12) placement of the tDCS electrodes (e.g., ipsilesional hemisphere or contralesional supraorbital region); (13) stimulation type (e.g., anodal, cathodal, or bihemispheric); (14) type of stroke, whether ischemic or hemorrhagic; (15) time after occurrence of stroke before intervention; and (16) type of therapy used. Data from figures were extracted using WebPlotDigitizer (version 4) (Kim et al., 2010; Bolognini et al., 2011, 2020; Khedr et al., 2013; Fusco et al., 2014; Beaulieu et al., 2019; Bornheim et al., 2020; Prathum et al., 2021). Studies that presented data as mean \pm SE were manually converted to mean \pm SD. Authors were contacted via email for studies whose data could not be extracted from visual inspection (Rossi et al., 2013; Ilic et al., 2016; Andrade et al., 2017; Cheng et al., 2021) and a response was not obtained.

Dosage Calculations

Information regarding the characteristics of the studies, such as number of tDCS sessions and stroke duration, can be found in **Supplementary Table 1** and information regarding the calculation of tDCS total charge density, such as electrode size and current, can be found in **Supplementary Table 2**. We used previous reported equations (Chhatbar et al., 2015) to calculate the dosage effect.

- Current density (mA/cm²) = Current (mA) ÷ Electrode size (cm²)
- Charge = Current (mA) \times tDCS duration (min) \div 60
- Charge density (mAh/cm²) = Charge (mAh) ÷ Electrode size (cm²)
- Total charge (mAh) = Charge (mAh) × Number of sessions (cm²)
- Total charge density (mAh/cm²) = Charge density (mAh/cm²) × tDCS sessions

Two additional equations were used to calculate total tDCS application time and number of sessions per week:

- Total tDCS time (min) = Number of sessions × Session time (min)
- Sessions per week = Number of sessions ÷ Intervention period (weeks).

Risk of Bias Assessment

The Cochrane Risk of Bias 2 tool (August 2019 version) was used. This tool assesses selection, reporting, performance, detection, and attrition biases in a 5-domain list containing multiple items with risk of bias being declared as high, low, or unclear. Two authors (JS and ADC) separately assessed each article's risk of bias using this tool and discussed the results. Any conflicting results were further discussed in extensive detail with a third investigator to come to a consensus.

Statistical Analysis

Forest plots were generated using Review Manager (RevMan version 5.4.1). Meta-analyses were performed for each of the four assessments for motor function recovery comparing the post-intervention data between the groups. Additionally, because minimal differences in the baseline motor function between the tDCS and sham groups were previously suggested to influence the results of tDCS intervention (Chhatbar et al., 2015), we also performed meta-analysis with the change score (mean difference between baseline and post-intervention) and pooled SD for each group. The influence of stroke severity and therapy type was investigated with tests for subgroup differences and significance was set at P = 0.10. For each comparison, mean differences between the groups were calculated with 95% CIs. For each analysis, a fixed effects model was used if the results were homogenous (P > 0.10) and a random effects model was used if heterogeneity was present. Sensitivity analysis was performed by determining the influence of each study from the model. Funnel plot's visual inspection indicated publication bias was unlikely.

Statistical Package for the Social Sciences (SPSS) (version 27) was used to assess the association between Hedge's *g* effect size (Hedges, 1981) and dosage (i.e., current density, charge, charge density, total charge, total charge density, and total tDCS time). Hedge's *g* effect size was calculated for each study using scores from pre- to post-intervention (i.e., post-intervention values – preintervention/pooled SD) of the tDCS groups for each assessment. Spearman's rho was used for the association analysis because of the lack of normal distribution in several variables according to the Shapiro–Wilk test. For all the analyses, the significance was set at P < 0.05.

Minimal clinically important difference was determined by the difference in pre- and post-intervention scores reaching a minimal value. To indicate a MCID in the tDCS or sham group, an increase must be shown in the score of the BI by at least 1.85 points (Hsieh et al., 2007), the upper extremity FMA by at least 5.25 points (Page et al., 2012), and the lower extremity FMA by at least 6 points (Pandian et al., 2016). To indicate a MCID for the MAS, however, a reduction of at least 0.48 points must be shown (Chen et al., 2019).

RESULTS

After excluding duplicates and performing a manual search on the reference list of the retrieved manuscripts, a total of 31 individual manuscripts were included in this review (Supplementary Figure 1). Five manuscripts (Kim et al., 2010; Hesse et al., 2011; Khedr et al., 2013; Rocha et al., 2016; Yi et al., 2016) recruited independent groups of participants to assess the effects of stimulation type (e.g., anodal vs. cathodal) in comparison to sham and in this case, careful consideration was taken to not include the sham group twice in the total sample size. The number of participant assessed with the BI, the upper extremity FMA, the lower extremity FMA, and the MAS tested was 515, 913, 179, and 172, respectively. Despite all the studies being randomized controlled trials, out of the 31 studies that were included in this review, one study reported significant differences between the tDCS and sham groups at baseline (Pinto et al., 2021) and some studies made no mention of baseline differences either in their discussion or with a statistical analysis (Allman et al., 2016; Achacheluee et al., 2018; Oveisgharan et al., 2018). The rest of the included studies reported that the sham and tDCS groups were similar at baseline.

Studies were excluded for several reasons such as: (1) Pilot studies that were later published (Hesse et al., 2007; Mazzoleni et al., 2017); (2) tDCS in combination with other forms of electrical stimulation, such as repetitive tDCS or neuromuscular electrical stimulation; (3) Studies with electrode placements to rehabilitate cognition instead of motor function; (4) Not reporting data on stroke individuals; (5) Studies that failed to indicate post-intervention descriptive or numerical results; and (6) No use of the FMA, the BI, or the MAS.

Study Characteristics

The majority of studies used different participants for the sham and tDCS intervention groups (i.e., between-group design) and only one study used a within-group design with washout period of 72 h (Achacheluee et al., 2018; **Supplementary Table 1**). Some manuscripts reported more than one assessment. Ten of them investigated the BI (Kim et al., 2010; Hesse et al., 2011; Khedr et al., 2013; Fusco et al., 2014; Yi et al., 2016; Koo et al., 2018; Bolognini et al., 2020; Bornheim et al., 2020; Yao et al., 2020; Pinto et al., 2010; Bolognini et al., 2011, 2020; Hesse et al., 2011; Nair et al., 2011; Khedr et al., 2013; Rossi et al., 2013;

		tDCS			Sham			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
.1.2 Moderate Dependency	61-90								and the second second
Bornheim et al. 2019	85.6	19.89	25	73.96	19.89	25	6.3%	11.64 [0.61, 22.67]	100
Kim et al. 2010 (Anodal)	79.07	12.59	6	73.79	12.95	7	3.9%	5.28 [-8.63, 19.19]	
Kim et al. 2010 (Cathodal) Subtotal (95% CI)	92.25	12.42	5 36	73.79	12.95	7 39	3.6% 13.9%	18.46 [3.95, 32.97]	•
deterogeneity Chi ² = 1.65 df	= 2 (P =	0 44) 1	= 0%					•	6
est for overall effect: Z = 3.07	r (P = 0.0)	002)	- 070						
.1.3 Severe Dependency 21	-60								-
usco et al. 2014	68.56	31.82	5	59.09	32.2	6	0.5%	9.47 [-28.50, 47.44]	3
lesse et al. 2011 (Anodal)	53.6	14.5	28	56.3	15.5	28	12.4%	-2.70 [-10.56, 5.16]	
lesse et al. 2011 (Cathodal)	59.2	12.4	29	56.3	15.5	28	14.3%	2.90 [-4.40, 10.20]	
hedr et al. 2013 (Anodal)	51.84	9.3	14	41.14	9.51	13	15.1%	10.70 [3.60, 17.80]	
hedr et al. 2013 (Cathodal)	52.24	9.96	13	41.14	9.51	13	13.6%	11.10 [3.61, 18.59]	
oo et al. 2018	65.25	13.02	12	55.17	17.96	12	4.8%	10.08 [-2.47, 22.63]	
into et al. 2021	74.8	22.7	31	62.1	28.6	29	4.4%	12.70 [-0.42, 25.82]	
ao et al. 2020	72	17.1	20	64.5	12.4	20	8.9%	7.50 [-1.76, 16.76]	
i et al. 2016 (Cathodal) ubtotal (95% CI)	46	20.5	10 162	36.8	13.3	10 159	3.3% 77.5%	9.20 [-5.95, 24.35] 6.83 [3.69, 9.97]	•
leterogeneity: Chi ² = 10.31, d est for overall effect: Z = 4.26	f = 8 (P = 6 (P < 0.0	= 0.24); 0001)	l ² = 229	6					
.1.4 Total Dependency 0-20									
olognini et al. 2020	34.66	17.05	16	40.07	17.05	16	5.5%	-5.41 [-17.22, 6.40]	
'i et al. 2016 (Anodal) subtotal (95% CI)	42.1	21.3	10 26	36.8	13.3	10 26	3.2% 8.6%	5.30 [-10.26, 20.86] -1.49 [-10.91, 7.92]	•
Heterogeneity: Chi ² = 1.15, df Fest for overall effect: Z = 0.31	= 1 (P = I (P = 0.7	0.28); lª 76)	= 13%						
Total (95% CI)			224			224	100.0%	6.77 [4.01, 9.54]	★ …
leterogeneity: Chi ² = 17.71, d	f = 13 (P	= 0.17)	; = 27	%					
est for overall effect: Z = 4.80	(P < 0.0)	00001)							-50 -25 0 25 5
est for subgroup differences.	$Chi^2 = 4$	60. df =	2 (P =	0.10)	= 56 5	%			Favouis sildini Favouis IDCS

FIGURE 1 | Effects of transcranial direct current stimulation (tDCS) on stroke recovery as assessed by the post-intervention data of the tDCS and sham groups for the Barthel Index.

Fusco et al., 2014; Viana et al., 2014; Ang et al., 2015; Triccas et al., 2015; Allman et al., 2016; Kim et al., 2016; Rocha et al., 2016; Straudi et al., 2016; Yi et al., 2016; Mazzoleni et al., 2017; Achacheluee et al., 2018; Oveisgharan et al., 2018; Beaulieu et al., 2019; Edwards et al., 2019; Jin et al., 2019; Alisar et al., 2020; Bornheim et al., 2020; Liao et al., 2020; Yao et al., 2020; Pinto et al., 2021; Prathum et al., 2021), five included the lower extremity FMA (Chang et al., 2015; Seo et al., 2017; Bornheim et al., 2020; Pinto et al., 2021; Prathum et al., 2021), and nine included the MAS (Hesse et al., 2011; Viana et al., 2014; Andrade et al., 2017; Mazzoleni et al., 2017; Beaulieu et al., 2019; Supplementary Table 1). Characteristics of dosage used (i.e., number of sessions, time, current, etc.) are shown in Supplementary Table 2. In summary, the number of test sessions, which the intervention was given to subjects, varied from as little as a single session to as many as 40. Application time per session of tDCS ranged from 10 to 30 min, with the majority applying tDCS for 20 min per session. The majority of studies (~48%) investigated tDCS effects on rehabilitation in patients with predominantly ischemic stroke and approximately 48% of studies investigated both the ischemic and hemorrhagic stroke (Supplementary Table 1). Three studies gave insufficient data or a range, so total charge density was not calculated (Supplementary Table 2). Six studies reported tDCS without the addition of other therapies, 11 studies used conventional therapy, 11 studies used assisted therapy, 1 study used a combination of both the conventional and robotic therapy, and three studies used other miscellaneous therapy types (Supplementary Table 1).

Risk of Bias

According to the Cochrane Risk of Bias 2 tool (**Supplementary Figures 2, 3**), common bias was related to inadequate description of allocation concealment, blinding of researchers, assessors, or therapists.

Overall Effects of Transcranial Direct Current Stimulation on Motor Recovery for Each Assessment and Influence of Stroke Severity or Stimulation Type Barthel Index

Descriptive

Out of the 10 manuscripts reporting the BI, only three manuscripts concluded that there was a significant difference between the tDCS group and the sham group. All the 10 studies reported reaching MCID (**Supplementary Table 1**) in both the sham and tDCS groups (**Supplementary Table 1**).

Meta-Analysis

Post-intervention data showed that tDCS was beneficial when assessed by the BI [mean difference: 6.77 (Confidence Interval (CI): 4.01, 9.54); P < 0.01] (**Figure 1**). Similarly, meta-analysis using the change scores showed a positive effect of tDCS [mean difference: 6.13 (CI: 2.56, 9.69); P < 0.01]. Sensitive analysis because of large SD in one study (Fusco et al., 2014) revealed that it had no influence on the above results.

Influence of Stroke Severity

The baseline BI score of each study was used to classify stroke severity and divided in the following: total dependency (0–20), severe dependency (21–60), and moderate dependency (61–90) (Collin et al., 1988). Subgroup analysis shows that the severity of stroke had a trend to influence the results presented for the post-intervention data meta-analysis (P = 0.10), so that studies in the moderate and severe dependency categories had greater effect size than the studies in the total dependency category. The influence of stroke severity should be cautiously interpreted considering the majority of studies are in the severe dependency category (**Figure 1**) and change score meta-analysis had no influence of stroke severity (P = 0.47).

Influence of Stimulation Type

Studies were divided in the following subgroups: cathodal, anodal, and bihemispheric to investigate the influence of tDCS montage on the above meta-analysis. Type of stimulation had no influence on the meta-analysis results showed above (test for subgroup differences: post-intervention data: P = 0.52 and change score: P = 0.25). The cathodal group had 166 participants (sham: 84 vs. tDCS: 82), the anodal group had 190 participants (sham: 95 vs. tDCS: 95), and the bihemispheric group had 92 participants (sham: 45 vs. tDCS: 47).

Upper Extremity Fugl–Meyer Assessment *Descriptive*

76 and 44% of the studies observed the tDCS and sham groups, respectively, reached MCID (**Supplementary Table 1**). The majority of the studies indicate both the sham and tDCS groups reached MCID and out of the 25 studies, only 8 studies reported that MCID was reached in the tDCS group, but not in the sham group.

Meta-Analysis

Post-intervention data showed that tDCS was not superior to sham (**Figure 2**). However, the meta-analysis of change scores showed a positive effect of tDCS [mean difference: 1.68 (CI: 0.25, 3.11); P = 0.02] (**Figure 3**). Sensitive analysis because of large SD in one study (Fusco et al., 2014) showed that it had no influence on the above results.

Influence of Stroke Severity

Subgroup analysis by severity of stroke, calculated with the upper extremity FMA score at the beginning of the intervention, as previously done (Van Hoornweder et al., 2021), did not influence the meta-analysis results (post-intervention data: P = 0.36 and change score: P = 0.16) (**Figures 2, 3**). However, the majority of studies are from the severe and moderate categories (9 and 12 studies, respectively) compared with only 4 studies in the mild impairment category.

Influence of Stimulation Type

Test of subgroups indicate that the meta-analysis showed above was not influenced by the type of stimulation (postintervention data: P = 0.32 and change score: P = 0.74). For this comparison, there were 408 participants in the anodal stimulation (sham: 203 vs. tDCS: 205), 148 participants in

Study or Subaroup	Moan	SDC3	Total	Moan	en	Total	Moight	N Eixed 05% Cl	N/ Eixed 05% Cl
A 4 Mild Impairment 43 66	Wean	30	Total	Wean	30	Total	weight	IV, FIXEU, 95% CI	IV, FIXEU, 55% CI
2. 1. 1 Mild Impail Ment 43-00	00.07	0.00	7	04.07		-	05.00	0.0011.55.0.55	-
Beaulieu et al. 2019	62.37	2.69	1	64.37	2.14	1	25.9%	-2.00 [-4.55, 0.55]	
Jin et al. 2019	49.1	15.61	10	50.5	17.3	10	0.8%	-1.40 [-15.84, 13.04]	
Oveisgharan et al. 2018	59.6	16.76	10	63.1	6.64	10	1.3%	-3.50 [-14.67, 7.67]	
Rocha et al. 2015 (Anodal)	55.7	4.3	7	54.85	6.7	7	4.8%	0.85 [-5.05, 6.75]	
Rocha et al. 2015 (Cathodal) Subtotal (95% Cl)	58.9	3.7	7 41	54.85	6.7	7 41	5.2% 38.2%	4.05 [-1.62, 9.72] - 0.85 [-2.95, 1.25]	•
Heterogeneity: Chi² = 4.19, df = Fest for overall effect: Z = 0.79 (4 (P = 0 P = 0.43	.38); I² =)	: 5%						
2.1.2 Moderate Impairment 29	-42								
chacheluee et al. 2018	38.33	5.65	15	38.46	5.81	15	10.0%	-0.13[-4.23, 3.97]	
lisaretal 2019	35	20.73	16	28.12	23.5	16	0.7%	6 88 1-8 47 22 23	
allman et al 2016	50 36	11.6	11	45.53	14 62	13	1.5%	4 83 [-5 67 15 33]	
and et al. 2015	36.2	3	10	35.4	4	q	16.4%	0.80 [-2.41.4.01]	
Rombeim et al 2019	50.08	20.64	25	41 5	20.64	25	1 3%	8 58 1-2 86 20 02	
(im et al. 2010 (Anodal)	45.57	12 50	6	49 31	12.04	7	0.9%	-3 74 [-17 65 10 17]	
(im et al. 2010 (Cathodal)	53.04	12.00	5	10 31	12.00	7	0.9%	3 73 [10 78 18 24]	
ian et al. 2010 (Cathodal)	41.67	11 41	12	34.5	10.46	8	1.8%	7 17 12 54 16 88	
indenhera et al 2010	41.07	12.2	10	34.5	11.40	10	1.6%	2 90 1 7 76 12 26	
andenberg et al. 2010	45.0	10 36	20	19 91	15 39	10	1.0%	-3.64 [-1.4.69, 7.31]	
lair at al. 2015	40.2	12.0	20	40.04	10.00	15	1.4%	1 40 [10 60 12 40]	
Prothum of al. 2011	12 19	15.25	12	14 92	16 25	12	1 1 1 %	-1 44 [-13 64 10 76]	
/iana et al. 2014 Subtotal (95% CI)	50.6	13.4	10	46.9	12.4	10	1.3%	3.70 [-7.62, 15.02]	
Jotorogonoity Chiz - 6 20 df-	12 /P -	0 001-12	- 0%			150	55.570	1.25 [-0.1 0, 0.05]	The second se
Fest for overall effect: Z = 1.24 (P = 0.22	0.90), r)	- 0 %						
1 3 Severe Impairment 0-28									
alegnini et al 2011	21 51	12.05	7	20.24	12.05	7	0.00	247144 40 45 741	
Subgrini et al. 2011	31.51	12.90	24	29.34	12.95	25	0.9%	2.17 [-11.40, 10.74]	
uwarus et al. 2019	32	10.0	34	33.4	19.2	30	2.1%	-1.40 [-10.37, 7.37]	10 La
usco et al. 2014	33.99	29.20	20	24.00	22.83	20	0.2%	9.43 [-22.07, 40.93]	The second se
lesse et al. 2011 (Anodal)	18.3	13.1	28	17.5	13.3	28	3.5%	0.80 [-0.11, 7.71]	
linto et al. 2021	18.9	10.5	29	17.5	13.3	28	4.5%	7 60 [0 70 45 00]	Contraction of the second
nito et al. 2021	24.1	10.2	31	10.5	14.5	29	2.4%	2.00 [-0.70, 15.90]	
troudiet el 2012	20.05	10.1	25	9.6	10.3	25	0.0%	-2.10[-1.10, 3.00]	
vience et al. 2016	28.05	16.90	12	20.04	16.12	11	0.0%	1.41 [-12.94, 15.76]	
nicas et al. 2015	33.04	10.25	12	44.82	10.29	11	0.9%	-11.18 [-24.49, 2.13]	and the second second
subtotal (95% Cl)	34.4	17.8	203	32.4	16.4	200	22.0%	0.48 [-2.29, 3.25]	•
łeterogeneity: Chi² = 7.30, df = 'est for overall effect: Z = 0.34 (9 (P = 0 P = 0.73	.61); l² =)	:0%						
otal (95% CI)			403			399	100.0%	0.30 [-1.00, 1.59]	+
leterogeneity: Chi ² = 19.86, df est for overall effect: Z = 0.45 (= 27 (P = P = 0.65	= 0.84);)	l² = 0%		0.000				-20 -10 0 10 20 Favours sham Favours tDCS
Fest for subgroup differences: (Chi ² = 2.	07. df =	2 (P =	0.36), I²	= 3.3%				

FIGURE 2 | Effects of tDCS on stroke recovery as assessed by the post-intervention data of the tDCS and sham groups for the upper extremity Fu Assessment.

the cathodal stimulation (sham: 75 vs. tDCS: 73), and 246 participants in the bihemispheric stimulation (sham: 121 vs. tDCS: 125).

Lower Extremity Fugl-Meyer Assessment

Descriptive

Out of the five studies that reported data from this assessment, three studies concluded that the use of tDCS was able to significantly improve recovery compared to sham. However, only one study reached MCID by an increase in score of at least six points in the tDCS group. No study reported MCID in the sham group (**Supplementary Table 1**).

Meta-Analysis

Post-intervention data showed that tDCS had a positive effect on motor recovery compared with sham [mean difference: 2.19 (CI: 1.07, 3.30); P < 0.01] (**Figure 4**). One study had a heavy weight on this analysis due to small SD reported (Chang et al., 2015) and removing this study from calculations maintained the positive effect of tDCS compared with sham intervention [mean difference: 2.51 (CI: 0.07, 4.94); P = 0.04]. However, meta-analysis of change scores showed no effect of tDCS [mean difference: -0.26 (CI: -1.82, 1.31); P = 0.75].

Influence of Stroke Severity

The lower extremity FMA score at the beginning of the intervention had no influence on the meta-analysis results (post-intervention data: P = 0.43 and change score: P = 0.54). However, there was only one study in the severe category (score < 21) and four studies in the moderate category (score > 21) (Kwong and Ng, 2019; **Figure 4**). The study with heavier weight in the meta-analysis is in the moderate impairment category.

Ctuchy or Culturoup	Mean	IDCS	Total	Maan	Sham	Tatal	Moinht	Mean Difference	Mean Difference
6.4.4 Mild Impairment 43.66	wear	50	Total	mean	50	Total	weight	IV, FIXEU, 95% CI	IV, FIXEU, 95% CI
Deculiau et al. 2010	E 75	2.00	7	2.46	254	7	10.50	2 20 (0 60 7 40)	10.00
Beaulieu et al. 2019	5.75	3.88	10	2.40	3.54	10	13.5%	3.29 [-0.60, 7.18]	18 19 19 19 19 19 19 19 19 19 19 19 19 19
Jin et al. 2019	3.3	15.79	10	1.2	11.12	10	0.9%	2.10 [-12.61, 16.81]	
Ovelsgnaran et al. 2018	14.3	36	10	9.9	11.92	10	0.4%	4.40 [-19.10, 27.90]	1 mar 1 m
Rocha et al. 2015 (Anodal)	11.1	4.2	1	3.85	7.88	1	4.7%	7.25 [0.64, 13.86]	100 C
Rocha et al. 2015 (Cathodal) Subtotal (95% Cl)	7.3	3.96	41	3.85	7.88	41	4.8% 24.3%	3.45 [-3.08, 9.98] 4.05 [1.15, 6.96]	•
Heterogeneity: Chi ² = 1.15, df =	4 (P = 0	.89); l² =	:0%						
Test for overall effect: Z = 2.74	(P = 0.00	6)							
6.1.2 Moderate Impairment 29	-42								
Achacheluee et al. 2018	0.07	5.65	15	0.26	5.75	15	12.3%	-0.19 [-4.27, 3.89]	
Alisar et al. 2019	9.12	19.26	16	1.74	23.55	16	0.9%	7.38 [-7.53, 22.29]	
Allman et al. 2016	11.46	13.73	11	9.11	16.06	13	1.4%	2.35 [-9.57, 14.27]	
Ang et al. 2015	0.9	5.91	10	2.8	6.39	9	6.6%	-1.90 [-7.45, 3.65]	
Bornheim et al. 2019	11.62	14.66	25	4.7	14.82	25	3.1%	6.92 [-1.25, 15.09]	
Kim et al. 2010 (Anodal)	14.57	11.92	6	8.31	11.7	7	1.2%	6.26 [-6.63, 19.15]	
Kim et al. 2010 (Cathodal)	13.84	16.05	5	8.31	11.7	7	0.7%	5.53 [-10.99, 22.05]	
Liao et al. 2020	5.34	9.82	12	3.75	9.83	8	2.6%	1.59 [-7.20, 10.38]	
Lindenberg et al. 2010	5.6	12.81	10	1.2	11.65	10	1.8%	4.40 [-6.33, 15.13]	
Mazzoleni et al. 2019	12	18.86	20	15.73	15.43	19	1.8%	-3.73 [-14.52, 7.06]	
Nair et al. 2011	4.1	12.17	7	1.7	10	7	1.5%	2.40 [-9.27, 14.07]	
Prathum et al. 2021	7.09	11.27	12	3.1	11.27	12	2.5%	3.99 [-5.03, 13.01]	
Viana et al. 2014	9.3	14.87	10	7.5	15.11	10	1.2%	1.80 [-11.34, 14.94]	
Subtotal (95% CI)			159			158	37.7%	1.31 [-1.02, 3.64]	*
Heterogeneity: Chi ² = 6.64, df =	12 (P =	0.88); l ²	= 0%						
Test for overall effect: Z = 1.11	(P = 0.27)							
6.1.3 Severe Impairment 0-28									
Bolognini et al. 2011	615	11 97	7	2.23	14 53	7	1 1 %	3 92 [-10 03 17 87]	
Edwards et al. 2019	6.3	17 59	34	81	17.81	35	2.9%	-1.80 [-10 15 6 55]	
Fusco et al. 2014	3.97	28.79	5	3.97	22.22	ĥ	0.2%	0.00 [-30.87, 30 87]	
Hesse et al. 2011 (Anodal)	10.5	9.64	28	9.3	9.91	28	7.8%	1.20 [-3.92, 6.32]	
Hesse et al. 2011 (Cathodal)	11	78	29	93	9.91	28	9.5%	1.70 [-2.94 6.34]	
Pinto et al. 2021	52	17.91	31	4.6	121	29	3.5%	0 60 1-7 09 8 291	
Rossietal 2012	3.4	8 4 5	25	0	914	25	8.6%	-1 60 [-6 48 3 28]	
Straudi et al. 2016	3.97	17.82	12	519	14 75	11	1 2%	-1 22 [-14 55 12 11]	
Triccas et al. 2015	8 73	1613	12	7 73	14 99	11	1 396	1.00 [-11.72 13.72]	
Yan et al. 2020	101	17 21	20	64	15.81	20	2.0%	3701-654 1394	
Subtotal (95% CI)	10.1	11.21	203	0.4	10.01	200	37.9%	0.52 [-1.80. 2.84]	•
Heterogeneity: Chi ² = 2.01 df =	9 (P = 0	99): I ² =	: 0%						T
Test for overall effect: Z = 0.44	(P = 0.66)	2.10						
Total (95% CI)			403			399	100.0%	1.68 [0.25, 3.11]	•
Haterogeneity Chi2-12 41 df	- 27 (P -	- 0 001-	12 - 0%						
Tast for overall effect: 7 = 2.20	- 27 (F -	- 0.99),	- 0%						-20 -10 Ó 10 20
	r = 110/2	1							

FIGURE 3 | Effects of tDCS on stroke recovery as assessed by the change scores (mean difference between baseline and post-intervention) of the tDCS and sham groups for the upper extremity Fugl-Meyer Assessment.

Influence of Stimulation Type

Subgroup test indicates that the meta-analysis presented above was not influenced by stimulation type (post-intervention data: P = 0.95 and change score: P = 0.56). For this analysis, there were 95 participants in the anodal subgroup (sham: 47 vs. tDCS: 48) and 84 participants in the bihemispheric subgroup (sham: 41 vs. tDCS: 43). No study using the lower extremity FMA used cathodal stimulation.

Modified Ashworth Scale

Descriptive

Out of the three studies that reported this assessment, none of the study had found MCID in either the tDCS or sham groups (**Supplementary Table 1**).

Meta-Analysis

Post-intervention data showed no effect of tDCS on the MAS (**Figure 5**). Likewise, the meta-analysis of change scores showed no effect in favor of tDCS [mean difference: -0.25 (CI: -0.76, 0.27); P = 0.35].

Influence of Stroke Severity

Subgroup analysis by severity was not performed for the MAS, as they reported similar baseline average values.

Influence of Stimulation Type

Test for subgroups indicate that there was no influence of type of stimulation on the meta-analysis described above (post-intervention data: P = 0.77 and change score: P = 0.47). Only one study used cathodal stimulation and no study used bihemispheric

Ctuche or Cubaroun	Maan	CD	Total	Maan	CD	Total	Moinht	N/ Fixed OFW CI	B/ Fixed OFM CI
Study of Sungroup	mean	50	Total	wean	50	Total	weight	IV, Fixed, 95% CI	IV, Fixed, 95% CI
3.1.1 Moderate Impair	ment > 2	21							
Bornheim et al. 2019	29.51	20.64	25	24.43	20.64	25	0.9%	5.08 [-6.36, 16.52]	
Chang et al. 2015	33.5	0.7	12	31.4	2.1	12	79.1%	2.10 [0.85, 3.35]	
Prathum et al. 2021	24.38	15.25	12	15.25	1.72	12	1.6%	9.13 [0.45, 17.81]	
Seo et al. 2017	24.1	3.4	11	21.4	4.7	10	9.9%	2.70 [-0.84, 6.24]	+
Subtotal (95% CI)			60			59	91.6%	2.32 [1.16, 3.49]	•
Heterogeneity: Chi ² = 2 Test for overall effect: 2	2.75, df = Z = 3.91 (3 (F = 0 P < 0.00).43),1)01)	- 0 %					
Heterogeneity: Chi ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme	2.75, df = Z = 3.91 (ent < 21	9 (F = 0 F < 0.00).43),1)01)	- 0 %	200		12/12/24		
Heterogeneity: Ch ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme Pinto et al. 2021 Subtotal (05% CD)	2.75, df = Z = 3.91 (ent < 21 17.9	3 (F = C P < 0.0(6.6	31 31	17.2	8.4	29	8.4%	0.70 [-3.14, 4.54]	
Heterogeneity: Chi ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme Pinto et al. 2021 Subtotal (95% CI)	2.75, df = 2 = 3.91 (ent < 21 17.9	3 (F = 0 F < 0.00 6.6	31 31 31	17.2	8.4	29 29	8.4% 8.4%	0.70 [-3.14, 4.54] 0.70 [-3.14, 4.54]	-
Heterogeneity: Chi ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme Pinto et al. 2021 Subtotal (95% CI) Heterogeneity: Not app	275, df = 2 = 3.91 () ent < 21 17.9 blicable	3 (F = 0 F < 0.00	31 31	17.2	8.4	29 29	8.4% 8.4%	0.70 [-3.14, 4.54] 0.70 [-3.14, 4.54]	-
Heterogeneity: Chi ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme Pinto et al. 2021 Subtotal (95% CI) Heterogeneity: Not app Test for overall effect: 2	275, df = 2 = 3.91 () ent < 21 17.9 blicable 2 = 0.36 ()	8 (P = 0.00 6.6 P = 0.72	31 31 31 31	17.2	8.4	29 29	8.4% 8.4%	0.70 [-3.14, 4.54] 0.70 [-3.14, 4.54]	-
Heterogeneity: Chi ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme Pinto et al. 2021 Subtotal (95% CI) Heterogeneity: Not app Test for overall effect: 2 Total (95% CI)	275, ar= 2 = 3.91 () ent < 21 17.9 blicable 2 = 0.36 ()	8 (P = 0.00 6.6 P = 0.72	31 31 31 2) 91	17.2	8.4	29 29 88	8.4% 8.4% 100.0%	0.70 [-3.14, 4.54] 0.70 [-3.14, 4.54] 2.19 [1.07, 3.30]	•
Heterogeneity: Chi ² = 2 Test for overall effect: 2 3.1.2 Severe Impairme Pinto et al. 2021 Subtotal (95% CI) Heterogeneity: Not app Test for overall effect: 2 Total (95% CI) Heterogeneity: Chi ² = 3	275, df = 2 = 3.91 () ent < 21 17.9 blicable 2 = 0.36 () 3.38, df =	8.6 P < 0.00 6.6 P = 0.72	31 31 31 2) 91 1.50): ²	- 0 % 17.2 = 0 %	8.4	29 29 88	8.4% 8.4% 100.0%	0.70 [-3.14, 4.54] 0.70 [-3.14, 4.54] 2.19 [1.07, 3.30]	

FIGURE 4 | Effects of tDCS on stroke recovery as assessed by the post-intervention data of the tDCS and sham groups for the lower extremity Fugl–Meyer Assessment.



stimulation. The subgroup for the anodal stimulation involved 115 individuals (sham: 57 vs. tDCS: 58) and the subgroup for the cathodal stimulation involved 57 individuals (sham: 28 vs. tDCS: 29).

Transcranial Direct Current Stimulation as a Stand-Alone Therapy

Subgroup meta-analysis was conducted on studies reporting the effects of tDCS without other intervention. Studies available assessed the effects of tDCS on the BI and the upper extremity FMA, but not the lower extremity FMA or the MAS. tDCS was not different than sham when used as a stand-alone therapy for the post-intervention data or change scores for either the BI or the upper extremity FMA (**Supplementary Table 3**).

Influence of Stroke Severity

For the BI, two studies had individuals in the severe dependency category (Fusco et al., 2014; Koo et al., 2018), but only one study found that tDCS was effective compared with sham (Koo et al., 2018). Also, for the BI, one study had individuals in the total dependency category (Bolognini et al., 2020) and tDCS was not beneficial compared with sham. For the upper extremity FMA, two studies had individuals in the severe category (Rossi et al., 2013; Fusco et al., 2014)

and one study had individuals in the moderate category (Achacheluee et al., 2018) and all of them showed that tDCS was not beneficial compared with sham. However, one study recruiting individuals in the mild category showed that tDCS was beneficial compared with sham (Oveisgharan et al., 2018).

Influence of Stimulation Type

For the BI, out of the three studies included in this review, one study used cathodal, one study used anodal, and one study used bihemispheric (**Supplementary Table 1**). Only one study using anodal stimulation showed a statistical difference between the sham and tDCS groups (Koo et al., 2018). Out of the four studies assessing the upper extremity FMA, two used anodal, one study used cathodal, and one study used bihemispheric stimulations. From these studies, the only one showing statistical significant differences between the tDCS and sham groups used bihemispheric stimulation (Oveisgharan et al., 2018).

Influence of the Type of Therapy Associated With Transcranial Direct Current Stimulation

Three subgroups of therapy types were previously suggested (Elsner et al., 2020) and used in this review: (1) conventional

(e.g., physical therapy or occupational therapy); (2) assisted (e.g., mirror, virtual reality, robot-assisted, or brain-computer interface-assisted motor imagery); and (3) miscellaneous (e.g., constraint-induced movement therapy). In these comparisons one study was excluded because it used a combination of two of the subgroups (Pinto et al., 2021).

Barthel Index

Meta-analysis of the post-intervention data shows that studies using conventional therapy associated with tDCS had improvement in the BI score, but not the studies using assisted therapy (subgroup difference: P < 0.01) (**Figure 6**). Meta-analysis with change score data agrees with the post-intervention data (P < 0.01) (**Supplementary Figure 4**).

Upper Extremity Fugl-Meyer Assessment

Meta-analysis of the post-intervention data shows that studies using conventional therapy associated with tDCS had similar results compared with studies using assisted or miscellaneous therapies (4.2 vs. 0.7 vs. 2.5; test for subgroup differences: P = 0.46) (**Supplementary Figure 5**). However, meta-analysis with change score data indicates that studies using conventional therapy were superior to assisted or miscellaneous therapies (3.9 vs. 0.4 vs. 5.19, respectively; test for subgroup differences: P = 0.07) (**Supplementary Figure 6**). The miscellaneous groups have fewer subjects compared with the other therapy types, but sensitive analysis shows that the above results are maintained (change score: P = 0.06 and post-intervention data: P = 0.27).

Lower Extremity Fugl-Meyer Assessment

Meta-analysis of both the post-intervention data and change scores shows that studies using conventional therapy associated with tDCS had similar results compared with assisted therapy combined with tDCS (test for subgroup differences: post-intervention data: P = 0.75; change score: P = 0.18). These results, however, may be a consequence of the lower overall number of studies and the majority using conventional therapy compared with assisted therapy (3 vs. 1) (**Supplementary Figure 7**). Moreover, the only study using assisted therapy found that tDCS combined with a robotic-assisted therapy was not beneficial as compared with sham (Seo et al., 2017). No study was included in the miscellaneous therapy subgroup. One study was excluded from this calculation because it used a combination of assisted and conventional therapies (Pinto et al., 2021).

Modified Ashworth Scale

Meta-analysis could not be conducted on the effects of therapy type using the MAS, as all the studies that provided pre- and postintervention data belonged to the same subgroup (i.e., all used assisted therapy).

Dose Response

We investigated if the Hedge's *g* effect size of tDCS intervention was influenced by: (1) number of sessions, (2) sessions per week, (3) session time, (4) total tDCS application time, (5) current, (6) electrode size, (7) current density, (8) charge, (9) charge density, (10) total charge, and (11) total charge density.

Barthel Index

Overall, there was no association between any of the dosage metrics with Hedge's *g* effect size (**Figure** 7) (All P > 0.05, **Supplementary Table 4**). These results were not influenced by the stimulation type.

Upper Extremity Fugl–Meyer Assessment

When all the available studies were included in the calculations, there was no influence of any metrics of dosage on effect size (**Figure 8**) (All P > 0.05, **Supplementary Table 4**). Of note, session time had a trend of negative association with effect size (r = -0.38; P = 0.05). The session time ranged from 9 to 40 min (**Supplementary Table 2**) and the negative trend was not maintained by removing a study with large effect size and short session time (Rocha et al., 2016). These results were not influenced by the stimulation type (i.e., cathodal vs. anodal vs. bihemispheric).

Modified Ashworth Scale and Lower Extremity Fugl–Meyer Assessment

For both the assessments, there was no association between any of the dosage metrics with Hedge's g effect size (all P > 0.05) (**Supplementary Figures 8, 9**). However, cautious interpretation is required due to the small number of studies reporting these assessments.

DISCUSSION

Transcranial direct current stimulation alone was ineffective to improve motor recovery after stroke. However, tDCS applied in combination with other therapies was somewhat beneficial for motor recovery of stroke survivors and the improvements could be dependent on the assessment used and associated therapy. Specifically: (1) tDCS applied in combination with other therapies was beneficial when assessed by the BI but not by the MAS, and the effects on the upper or lower extremity FMA are unclear (2) conventional therapy combined with tDCS had a greater impact on motor function relative to assisted therapy combined with tDCS when assessed by the BI for either postscores (~13 vs. ~2, respectively, Figure 6) or changes scores (~11 vs. 2, respectively, Supplementary Figure 4). Additionally, for the upper extremity FMA, tDCS combined with conventional and miscellaneous therapies had greater benefits than assisted therapy (3.9 vs. 5.9 vs. 0.44, respectively) when change scores, which are more likely to detect small alterations in motor recovery, were used for meta-analysis. Type of stimulation (i.e., anodal vs. cathodal vs. bihemispheric) had no influence on the motor recovery, which agrees with previous reviews (Elsner et al., 2020; Van Hoornweder et al., 2021). Contrary to our hypothesis, tDCS dosage has minimal influence on the recovery of motor function after stroke. Another new aspect in this review is the evaluation of MCID. Conversely to statistical differences between the tDCS and sham groups, MCID analysis adds to the interpretation by showing that MCID is frequent in both the sham and tDCS groups, particularly when using the BI and the FMA, but not the MAS assessment. The improvement in

		IDCS			Sham			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
13.1.1 Conventional									
Bornheim et al. 2019	85.6	19.89	25	73.96	19.89	25	9.5%	11.64 [0.61, 22.67]	
Khedr et al. 2013 (Anodal)	60.29	9.3	14	41.14	9.51	13	13.3%	19.15 [12.05, 26.25]	
Khedr et al. 2013 (Cathodal)	52.24	9.96	13	41.14	9.51	13	12.9%	11.10 [3.61, 18.59]	
Kim et al. 2010 (Anodal)	79.07	12.59	6	73.79	12.95	7	7.4%	5.28 [-8.63, 19.19]	
Kim et al. 2010 (Cathodal)	92.25	12.42	5	73.79	12.95	7	7.1%	18.46 [3.95, 32.97]	
Yi et al. 2016 (Anodal)	42.1	21.3	10	36.8	13.3	10	6.5%	5.30 [-10.26, 20.86]	
Yi et al. 2016 (Cathodal)	46	20.5	10	36.8	13.3	10	6.7%	9.20 [-5.95, 24.35]	
Subtotal (95% CI)	(and a set	-	83			85	63.3%	13.19 [9.21, 17.18]	
Heterogeneity: Tau ² = 0.42; Chi ²	^e = 6.08,	df = 6 (P = 0.4	1); l² = 1	%				
Test for overall effect: Z = 6.49 (I	^o < 0.00	001)							
13.1.2 Assisted									
Hesse et al. 2011 (Anodal)	53.6	14.5	28	56.3	15.5	28	12.5%	-2.70 [-10.56, 5.16]	
Hesse et al. 2011 (Cathodal)	59.2	12.4	29	56.3	15.5	28	13.1%	2.90 [-4.40, 10.20]	
Yao et al. 2020	72	17.1	20	64.5	12.4	20	11.1%	7.50 [-1.76, 16.76]	
Subtotal (95% CI)			77			76	36.7%	2.21 [-3.29, 7.71]	
Heterogeneity: Tau ² = 6.70; Chi ²	= 2.79,	df = 2(P = 0.2	5); I ² = 2	28%				
Test for overall effect: Z = 0.79 (I	P = 0.43)							
Total (95% CI)			160			161	100.0%	8.61 [3.69, 13.53]	•
Heterogeneity: Tau ² = 34.30; Ch	i ² = 21.7	70. df =	9 (P = 0	0.010); P	² = 59%			-	
Test for overall effect: Z = 3.43 (I	P = 0.00	06)	500						-20 -10 0 10 20
Test for subgroup differences: (chi ^z = 10	0.04, df	= 1 (P =	= 0.002)	, l [≈] = 90	.0%			Favours sharn Favours IDCS
GURE 6 Effects of therapy (conve	entional	and as	sisted)	combir	ned witl	n tDCS	on strok	e recovery as assessed	by the post-intervention data of the tDCS and
oups for the Barthel Index.									

MCID in the sham group shows the effectiveness of the associated therapy. tDCS benefits, estimated with group differences in effect sizes, were observed, despite the associated therapy. Our results agree with others (Elsner et al., 2020; Van Hoornweder et al., 2021) showing limited, but positive evidence of effect of tDCS on upper limb motor function (i.e., the FMA) and gross motor recovery (i.e., the BI). For muscle tone and the lower extremity function, the evidence is scarce.

Potential Factors Influencing the Transcranial Direct Current Stimulation Results on Motor Recovery in This Review

Randomized controlled trials assume the groups are similar at baseline, and statistical differences post-treatment are a consequence of the intervention. However, in stroke individuals, minimal differences in function between the sham and tDCS groups at the start of treatment could influence the overall interpretation considering that randomized controlled trials are typically underpowered to detect between group differences at baseline (Chhatbar et al., 2015). This is shown in our metaanalyses for the upper and lower extremity FMA. Specifically, tDCS had a significant effect on post-stroke recovery for the upper extremity FMA (25 studies included) when using change scores (i.e., mean difference between baseline and postintervention for each group), but not with the post-intervention data. For the lower extremity FMA (five studies), the postintervention data meta-analysis showed positive results of tDCS compared with sham, but the change score data, which is more sensitive to small changes from the intervention, showed that tDCS had no benefit compared with sham.

It was previously reported that the severity of stroke can influence the tDCS results, so less severe individuals have greater recovery using linear regressions or subgroup analysis (Baltar et al., 2020). Our subgroup analysis showed that the stroke severity did not influence the meta-analysis results for the upper and lower FMA and the MAS; however, for the BI assessment, the studies with less severe individuals had a trend for better recovery compared with studies using more severe individuals. Heterogeneity in the participant's characteristics across different assessments can potentially explain the results of the current review. Specifically, out of the 10 total possible studies included in the BI meta-analyses, the majority of the studies had subacute participants (i.e., between 1 week and 3 months) compared with chronic (i.e., >3 months) (approximately 60 vs. 10%, respectively), whereas 30% did not clearly state the duration of stroke. Out of the 25 total possible studies include in the upper extremity FMA meta-analyses, approximately 32% had subacute participants, approximately 52% had chronic participants, and approximately 16% did not clearly indicate stroke duration.

Differences in the assessment scales themselves should also be considered. The BI is an assessment for gross movements used for activities of daily living scored broadly, whereas the upper extremity FMA and the lower extremity FMA are assessments for specific movements related to motor function with highly detailed scoring. Lastly, the MAS assesses muscle tone with a small scoring range. There are ample opportunities for future studies to investigate the effects of tDCS using multiple assessments, as the effects of tDCS may be dependent on the assessment used and combined therapy. Reporting individual effect size from the sham and tDCS intervention for each assessment, as well as stroke severity, will also provide opportunities for more advanced analysis in future reviews.





Effect of Therapy Type Combined With Transcranial Direct Current Stimulation

Meta-analyses of post-intervention results between the tDCS and sham groups suggest that the type of therapy post-stroke patients receive in combination with tDCS may determine the overall effectiveness of recovery. Specifically, the BI postintervention and change scores results strengthen the evidence that assisted-type and conventional-type therapies have a positive effect on recovery for this assessment type. Additionally, the effect size of conventional therapy was larger than the assisted therapy for the BI, which mostly involves studies using more acute participants compared to the other assessments (**Supplementary Table 1**). For the upper extremity FMA, conventional therapy was beneficial compared with assisted therapy when change score data was analyzed. Miscellaneous-type therapy had a positive effect on upper extremity motor recovery using the upper extremity FMA only. However, evidence of using miscellaneous therapy in addition to tDCS is scarce and heterogeneous thus



FIGURE 8 | Effect size in the tDCS groups assessed by the upper extremity Fugl–Meyer Assessment relative to: (A) Number of sessions, (B) sessions per week, (C) session time, (D) total tDCS application time, (E) current, (F) electrode size, (G) current density, (H) charge, (I) charge density, (J) total charge, and (K) total charge density (TCD). An increase in Hedge's g effect size indicates a better score.

results should be interpreted cautiously. Likewise, there are a reduced number of studies in the comparison of therapy type when using the MAS and the lower extremity FMA assessments.

Dose Effect of Transcranial Direct Current Stimulation

The dosage effect was investigated using several parameters from tDCS and obtained effect size. The obtained effect size was not influenced by any metrics of dosage and there was also no influence of stimulation type (anodal vs. cathodal vs. hemispheric) or stroke duration. A previous review investigated the effects of dosage on motor recovery after stroke; however, only using the upper extremity FMA (Chhatbar et al., 2015). They showed that electrode size (cm²), charge density (mAh/cm²), and current density (mA/cm²) had significant dose-response relationships with upper extremity FMA and bihemispheric stimulation could be advantageous, which contrast with the current findings. Others also investigated the dosage effect by clustering studies in subgroups (Van Hoornweder et al., 2021) and found that the current, charge density, and stimulation duration influenced the obtained effect size. The previous review, however, included between 8 and 18 manuscripts, whereas the current review included 25 manuscripts (three of them conducting more than one study) for the upper extremity FMA and may explain the discrepancy between reviews. We also conduct analysis on the BI, which was not previously investigated. There are ample opportunities for mechanistic studies investigating dosage effect in stroke individuals.

Limitations and Future Directions

Precaution toward the results of the lower extremity FMA and the MAS should be given due to a low number of studies retrieved. The larger effect size of conventional therapy studies may be consequence of assisted studies using techniques that are in preliminary stages. Small difference in baseline function between the groups was found to influence the comparison between the tDCS and sham groups and should be considered when designing new randomized controlled trials. Future studies should provide clear details of participant's baseline function, stroke duration, and allocation concealment. Likewise, blinding the investigators applying the assessment scales from the participant group allocation is encouraged to minimize bias toward study's hypothesis. Given that tDCS effects may be dependent on the assessment used and stroke duration, future original studies should report multiple motor function aspects in the same participants such as muscle tone, specific and gross movements (as indicated by the MAS, the FMA, and the BI, respectively), as well as carefully balance the groups for stroke duration. Additionally, because placement of reference electrode is heterogeneous (Supplementary Table 1), mechanistic studies should investigate its effects in patients with stroke.

CONCLUSION

Evidence for the use of tDCS as a stand-alone therapy tDCS is weak. However, tDCS associated with other therapies had a positive effect when assessed by the BI but not by the MAS. The impact of tDCS is unclear when assessed by the upper or

REFERENCES

- Achacheluee, S. T., Rahnama, L., Karimi, N., Abdollahi, I., Arslan, S. A., and Jaberzadeh, S. (2018). The effect of unihemispheric concurrent dual-site transcranial direct current stimulation of primary motor and dorsolateral prefrontal cortices on motor function in patients with sub-acute stroke. *Front. Hum. Neurosci.* 12:441. doi: 10.3389/fnhum.2018.00441
- Adeyemo, B. O., Simis, M., Macea, D. D., and Fregni, F. (2012). Systematic review of parameters of stimulation, clinical trial design characteristics, and motor outcomes in non-invasive brain stimulation in stroke. *Front. Psychiatry* 3:88. doi: 10.3389/fpsyt.2012.00088
- Alisar, D. C., Ozen, S., and Sozay, S. (2020). Effects of bihemispheric transcranial direct current stimulation on upper extremity function in stroke patients: a randomized double-blind sham-controlled study. *J. Stroke Cerebrovasc. Dis.* 29:104454. doi: 10.1016/j.jstrokecerebrovasdis.2019. 104454
- Allman, C., Amadi, U., Winkler, A. M., Wilkins, L., Filippini, N., Kischka, U., et al. (2016). Ipsilesional anodal tDCS enhances the functional benefits of rehabilitation in patients after stroke. *Sci. Transl. Med.* 8:330re331. doi: 10.1126/ scitranslmed.aad5651

lower extremity FMA. Severity of stroke had minimal influence in these analyses and the effect of stroke duration is unclear. These findings combined suggest that tDCS could be beneficial for functionality and dependent on the assessment tool used. Dosage (e.g., sessions per week, duration, or charge) as well as stimulation type (anodal vs. cathodal) had no influence on the tDCS results, which may simplify the prescription of the technique. Large prospective controlled studies using different types of assessment should investigate the potential task-dependent benefit of tDCS in stroke individuals.

DATA AVAILABILITY STATEMENT

The raw data supporting the conclusions of this article will be made available upon reasonable request to the corresponding author.

AUTHOR CONTRIBUTIONS

A-MC, HP, HW, and JK designed the study. A-MC and JS reviewed the original studies and assessed the risk of bias. A-MC performed all the calculations and wrote the main draft. All authors approved the final version of the manuscript.

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SUPPLEMENTARY MATERIAL

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- Andrade, S. M., Batista, L. M., Nogueira, L. L., De Oliveira, E. A., De Carvalho, A. G., Lima, S. S., et al. (2017). Constraint-induced movement therapy combined with transcranial direct current stimulation over premotor cortex improves motor function in severe stroke: a pilot randomized controlled trial. *Rehabil. Res. Pract.* 2017:6842549. doi: 10.1155/2017/6842549
- Ang, K. K., Guan, C., Phua, K. S., Wang, C., Zhao, L., Teo, W. P., et al. (2015). Facilitating effects of transcranial direct current stimulation on motor imagery brain-computer interface with robotic feedback for stroke rehabilitation. Arch. Phys. Med. Rehabil. 96, S79–S87. doi: 10.1016/j.apmr.2014. 08.008
- Baltar, A., Piscitelli, D., Marques, D., Shirahige, L., and Monte-Silva, K. (2020). Baseline motor impairment predicts transcranial direct current stimulation combined with physical therapy-induced improvement in individuals with chronic stroke. *Neural Plast.* 2020:8859394. doi: 10.1155/2020/885 9394
- Beaulieu, L. D., Blanchette, A. K., Mercier, C., Bernard-Larocque, V., and Milot, M. H. (2019). Efficacy, safety, and tolerability of bilateral transcranial direct current stimulation combined to a resistance training program in chronic stroke survivors: a double-blind, randomized, placebo-controlled pilot study. *Restor. Neurol. Neurosci.* 37, 333–346. doi: 10.3233/RNN-190908

- Bikson, M., Grossman, P., Thomas, C., Zannou, A. L., Jiang, J., Adnan, T., et al. (2016). Safety of transcranial direct current stimulation: evidence based update 2016. *Brain Stimul.* 9, 641–661.
- Bolognini, N., Russo, C., Souza Carneiro, M. I., Nicotra, A., Olgiati, E., Spandri, V., et al. (2020). Bi-hemispheric transcranial direct current stimulation for upperlimb hemiparesis in acute stroke: a randomized, double-blind, sham-controlled trial. *Eur. J. Neurol.* 27, 2473–2482. doi: 10.1111/ene.14451
- Bolognini, N., Vallar, G., Casati, C., Latif, L. A., El-Nazer, R., Williams, J., et al. (2011). Neurophysiological and behavioral effects of tDCS combined with constraint-induced movement therapy in poststroke patients. *Neurorehabil. Neural Repair* 25, 819–829. doi: 10.1177/1545968311411056
- Bornheim, S., Croisier, J. L., Maquet, P., and Kaux, J. F. (2020). Transcranial direct current stimulation associated with physical-therapy in acute stroke patients – a randomized, triple blind, sham-controlled study. *Brain Stimul.* 13, 329–336. doi: 10.1016/j.brs.2019.10.019
- Butler, A. J., Shuster, M., O'hara, E., Hurley, K., Middlebrooks, D., and Guilkey, K. (2013). A meta-analysis of the efficacy of anodal transcranial direct current stimulation for upper limb motor recovery in stroke survivors. *J. Hand Ther.* 26, 162–170; quiz171. doi: 10.1016/j.jht.2012.07.002
- CDC (2021). *Stroke Facts [Online]*. Available: https://www.cdc.gov/stroke/facts. httm (accessed March 23, 2021)
- Chang, M. C., Kim, D. Y., and Park, D. H. (2015). Enhancement of cortical excitability and lower limb motor function in patients with stroke by transcranial direct current stimulation. *Brain Stimul.* 8, 561–566. doi: 10.1016/ j.brs.2015.01.411
- Chen, C. L., Chen, C. Y., Chen, H. C., Wu, C. Y., Lin, K. C., Hsieh, Y. W., et al. (2019). Responsiveness and minimal clinically important difference of Modified Ashworth Scale in patients with stroke. *Eur. J. Phys. Rehabil. Med.* 55, 754–760. doi: 10.23736/S1973-9087.19.05545-X
- Cheng, H. J., Ng, K. K., Qian, X., Ji, F., Lu, Z. K., Teo, W. P., et al. (2021). Task-related brain functional network reconfigurations relate to motor recovery in chronic subcortical stroke. *Sci. Rep.* 11:8442. doi: 10.1038/s41598-021-87789-5
- Chhatbar, P. Y., Ramakrishnan, V., Kautz, S., George, M. S., Adams, R. J., and Feng, W. (2015). Transcranial direct current stimulation post-stroke upper extremity motor recovery studies exhibit a dose-response relationship. *Brain Stimul.* 9, 16–26. doi: 10.1016/j.brs.2015.09.002
- Collin, C., Wade, D., Davies, S., and Horne, V. (1988). The barthel ADL index: a reliability study. *Int. Disabil. Stud.* 10, 61–63. doi: 10.3109/09638288809164103
- Di Pino, G., Pellegrino, G., Assenza, G., Capone, F., Ferreri, F., Formica, D., et al. (2014). Modulation of brain plasticity in stroke: a novel model for neurorehabilitation. *Nat. Rev. Neurol.* 10, 597–608. doi: 10.1038/nrneurol.2014. 162
- Edwards, D. J., Cortes, M., Rykman-Peltz, A., Chang, J., Elder, J., Thickbroom, G., et al. (2019). Clinical improvement with intensive robot-assisted arm training in chronic stroke is unchanged by supplementary tDCS. *Restor. Neurol. Neurosci.* 37, 167–180. doi: 10.3233/RNN-180869
- Elsner, B., Kugler, J., Pohl, M., and Mehrholz, J. (2020). Transcranial direct current stimulation (tDCS) for improving activities of daily living, and physical and cognitive functioning, in people after stroke. *Cochrane Database Syst. Rev.* 11:CD009645.
- Fusco, A., Assenza, F., Iosa, M., Izzo, S., Altavilla, R., Paolucci, S., et al. (2014). The ineffective role of cathodal tDCS in enhancing the functional motor outcomes in early phase of stroke rehabilitation: an experimental trial. *Biomed. Res. Int.* 2014:547290. doi: 10.1155/2014/547290
- Hatem, S. M., Saussez, G., Della Faille, M., Prist, V., Zhang, X., Dispa, D., et al. (2016). Rehabilitation of motor function after stroke: a multiple systematic review focused on techniques to stimulate upper extremity recovery. *Front. Hum. Neurosci.* 10:442. doi: 10.3389/fnhum.2016. 00442
- Hedges, L. V. (1981). Distribution theory for Glass's estimator of effect size and related estimators. J. Educ. Stat. 6, 107–128. doi: 10.2307/1164588
- Hesse, S., Waldner, A., Mehrholz, J., Tomelleri, C., Pohl, M., and Werner, C. (2011). Combined transcranial direct current stimulation and robot-assisted arm training in subacute stroke patients: an exploratory, randomized multicenter trial. *Neurorehabil. Neural Repair* 25, 838–846. doi: 10.1177/154596831141 3906

- Hesse, S., Werner, C., Schonhardt, E. M., Bardeleben, A., Jenrich, W., and Kirker, S. G. (2007). Combined transcranial direct current stimulation and robotassisted arm training in subacute stroke patients: a pilot study. *Restor. Neurol. Neurosci.* 25, 9–15.
- Hsieh, Y. W., Wang, C. H., Wu, S. C., Chen, P. C., Sheu, C. F., and Hsieh, C. L. (2007). Establishing the minimal clinically important difference of the Barthel Index in stroke patients. *Neurorehabil. Neural Repair* 21, 233–238. doi: 10.1177/ 1545968306294729
- Hummel, F. C., and Cohen, L. G. (2006). Non-invasive brain stimulation: a new strategy to improve neurorehabilitation after stroke? *Lancet Neurol.* 5, 708–712. doi: 10.1016/S1474-4422(06)70525-7
- Ilic, N. V., Dubljanin-Raspopovic, E., Nedeljkovic, U., Tomanovic-Vujadinovic, S., Milanovic, S. D., Petronic-Markovic, I., et al. (2016). Effects of anodal tDCS and occupational therapy on fine motor skill deficits in patients with chronic stroke. *Restor. Neurol. Neurosci.* 34, 935–945. doi: 10.3233/RNN-160668
- Jin, M., Zhang, Z., Bai, Z., and Fong, K. N. K. (2019). Timing-dependent interaction effects of tDCS with mirror therapy on upper extremity motor recovery in patients with chronic stroke: a randomized controlled pilot study. *J. Neurol. Sci.* 405:116436. doi: 10.1016/j.jns.2019.116436
- Khedr, E. M., Shawky, O. A., El-Hammady, D. H., Rothwell, J. C., Darwish, E. S., Mostafa, O. M., et al. (2013). Effect of anodal versus cathodal transcranial direct current stimulation on stroke rehabilitation: a pilot randomized controlled trial. *Neurorehabil. Neural Repair* 27, 592–601. doi: 10.1177/1545968313484808
- Kim, D. Y., Lim, J. Y., Kang, E. K., You, D. S., Oh, M. K., Oh, B. M., et al. (2010). Effect of transcranial direct current stimulation on motor recovery in patients with subacute stroke. *Am. J. Phys. Med. Rehabil.* 89, 879–886. doi: 10.1097/PHM.0b013e3181f70aa7
- Kim, K., Kim, Y. M., and Kim, E. K. (2014). Correlation between the activities of daily living of stroke patients in a community setting and their quality of life. *J. Phys. Ther. Sci.* 26, 417–419. doi: 10.1589/jpts.26.417
- Kim, K. U., Kim, S. H., and An, T. G. (2016). Effect of transcranial direct current stimulation on visual perception function and performance capability of activities of daily living in stroke patients. *J. Phys. Ther. Sci.* 28, 2572–2575. doi: 10.1589/jpts.28.2572
- Koo, W. R., Jang, B. H., and Kim, C. R. (2018). Effects of anodal transcranial direct current stimulation on somatosensory recovery after stroke: a randomized controlled trial. Am. J. Phys. Med. Rehabil. 97, 507–513. doi: 10.1097/PHM. 000000000000910
- Kwong, P. W. H., and Ng, S. S. M. (2019). Cutoff score of the lower-extremity motor subscale of fugl-meyer assessment in chronic stroke survivors: a cross-sectional study. Arch. Phys. Med. Rehabil. 100, 1782–1787. doi: 10.1016/j.apmr.2019.01. 027
- Liao, W. W., Chiang, W. C., Lin, K. C., Wu, C. Y., Liu, C. T., Hsieh, Y. W., et al. (2020). Timing-dependent effects of transcranial direct current stimulation with mirror therapy on daily function and motor control in chronic stroke: a randomized controlled pilot study. *J. Neuroeng. Rehabil.* 17:101. doi: 10.1186/ s12984-020-00722-1
- Lindenberg, R., Renga, V., Zhu, L. L., Nair, D., and Schlaug, G. (2010). Bihemispheric brain stimulation facilitates motor recovery in chronic stroke patients. *Neurology* 75, 2176–2184. doi: 10.1212/WNL.0b013e318202013a
- Mazzoleni, S., Tran, V. D., Iardella, L., Dario, P., and Posteraro, F. (2017). Randomized, sham-controlled trial based on transcranial direct current stimulation and wrist robot-assisted integrated treatment on subacute stroke patients: intermediate results. *IEEE Int. Conf. Rehabil. Robot.* 2017, 555–560. doi: 10.1109/ICORR.2017.8009306
- Nair, D. G., Renga, V., Lindenberg, R., Zhu, L., and Schlaug, G. (2011). Optimizing recovery potential through simultaneous occupational therapy and noninvasive brain-stimulation using tDCS. *Restor. Neurol. Neurosci.* 29, 411–420. doi: 10.3233/RNN-2011-0612
- Nitsche, M. A., and Paulus, W. (2000). Excitability changes induced in the human motor cortex by weak transcranial direct current stimulation. *J. Physiol.* 527(Pt 3), 633–639. doi: 10.1111/j.1469-7793.2000.t01-1-00633.x
- Orrù, G., Conversano, C., Hitchcott, P. K., and Gemignani, A. (2019). Motor stroke recovery after tDCS: a systematic review. *Rev. Neurosci.* 31, 201–218. doi: 10.1515/revneuro-2019-0047
- Oveisgharan, S., Organji, H., and Ghorbani, A. (2018). Enhancement of motor recovery through left dorsolateral prefrontal cortex stimulation after acute

ischemic stroke. J. Stroke Cerebrovasc. Dis. 27, 185–191. doi: 10.1016/j. jstrokecerebrovasdis.2017.08.026

- Page, M. J., Mckenzie, J. E., Bossuyt, P. M., Boutron, I., Hoffmann, T. C., Mulrow, C. D., et al. (2021). The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. *BMJ* 372:n71.
- Page, S. J., Fulk, G. D., and Boyne, P. (2012). Clinically important differences for the upper-extremity Fugl-Meyer Scale in people with minimal to moderate impairment due to chronic stroke. *Phys. Ther.* 92, 791–798. doi: 10.2522/ptj. 20110009
- Pandian, S., Arya, K. N., and Kumar, D. (2016). Minimal clinically important difference of the lower-extremity fugl-meyer assessment in chronic-stroke. *Top. Stroke Rehabil.* 23, 233–239. doi: 10.1179/1945511915Y.0000000003
- Pinto, E. F., Gupta, A., Kulkarni, G. B., and Andrade, C. (2021). A randomized, double-blind, sham-controlled study of transcranial direct current stimulation as an augmentation intervention for the attenuation of motor deficits in patients with stroke. J. ECT 37, 281–290. doi: 10.1097/YCT.000000000000769
- Prathum, T., Piriyaprasarth, P., Aneksan, B., Hiengkaew, V., Pankhaew, T., Vachalathiti, R., et al. (2021). Effects of home-based dual-hemispheric transcranial direct current stimulation combined with exercise on upper and lower limb motor performance in patients with chronic stroke. *Disabil. Rehabil.* 1–12. doi: 10.1080/09638288.2021.1891464 [Epub ahead of print].
- Rocha, S., Silva, E., Foerster, A., Wiesiolek, C., Chagas, A. P., Machado, G., et al. (2016). The impact of transcranial direct current stimulation (tDCS) combined with modified constraint-induced movement therapy (mCIMT) on upper limb function in chronic stroke: a double-blind randomized controlled trial. *Disabil. Rehabil.* 38, 653–660. doi: 10.3109/09638288.2015.1055382
- Rossi, C., Sallustio, F., Di Legge, S., Stanzione, P., and Koch, G. (2013). Transcranial direct current stimulation of the affected hemisphere does not accelerate recovery of acute stroke patients. *Eur. J. Neurol.* 20, 202–204. doi: 10.1111/j. 1468-1331.2012.03703.x
- Seo, H. G., Lee, W. H., Lee, S. H., Yi, Y., Kim, K. D., and Oh, B. M. (2017). Roboticassisted gait training combined with transcranial direct current stimulation in chronic stroke patients: a pilot double-blind, randomized controlled trial. *Restor. Neurol. Neurosci.* 35, 527–536. doi: 10.3233/RNN-170745
- Straudi, S., Fregni, F., Martinuzzi, C., Pavarelli, C., Salvioli, S., and Basaglia, N. (2016). tDCS and robotics on upper limb stroke rehabilitation: effect modification by stroke duration and type of stroke. *Biomed. Res. Int.* 2016:5068127. doi: 10.1155/2016/5068127
- Triccas, L. T., Burridge, J. H., Hughes, A., Verheyden, G., Desikan, M., and Rothwell, J. (2015). A double-blinded randomised controlled trial exploring the effect of anodal transcranial direct current stimulation and uni-lateral robot therapy for the impaired upper limb in sub-acute

and chronic stroke. NeuroRehabilitation 37, 181-191. doi: 10.3233/NRE-15 1251

- Van Hoornweder, S., Vanderzande, L., Bloemers, E., Verstraelen, S., Depestele, S., Cuypers, K., et al. (2021). The effects of transcranial direct current stimulation on upper-limb function post-stroke: a meta-analysis of multiplesession studies. *Clin. Neurophysiol.* 132, 1897–1918. doi: 10.1016/j.clinph.2021. 05.015
- Viana, R. T., Laurentino, G. E., Souza, R. J., Fonseca, J. B., Silva Filho, E. M., Dias, S. N., et al. (2014). Effects of the addition of transcranial direct current stimulation to virtual reality therapy after stroke: a pilot randomized controlled trial. *NeuroRehabilitation* 34, 437–446. doi: 10.3233/NRE-141065
- Xu, J., Branscheidt, M., Schambra, H., Steiner, L., Widmer, M., Diedrichsen, J., et al. (2019). Rethinking interhemispheric imbalance as a target for stroke neurorehabilitation. Ann. Neurol. 85, 502–513. doi: 10.1002/ana.25452
- Yao, X., Cui, L., Wang, J., Feng, W., Bao, Y., and Xie, Q. (2020). Effects of transcranial direct current stimulation with virtual reality on upper limb function in patients with ischemic stroke: a randomized controlled trial. *J. Neuroeng. Rehabil.* 17:73. doi: 10.1186/s12984-020-00699-x
- Yi, Y. G., Chun, M. H., Do, K. H., Sung, E. J., Kwon, Y. G., and Kim, D. Y. (2016). The effect of transcranial direct current stimulation on neglect syndrome in stroke patients. *Ann. Rehabil. Med.* 40, 223–229. doi: 10.5535/arm.2016.40.2. 223

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