

# Relative significance of environmental factors affecting hydrogen production from landfilled refuse samples

The relative significance of 11 environmental factors on the apparent steady-state concentration of hydrogen ( $ASSCH_2$ ) achieved during anaerobic degradation of refuse collected from landfills was evaluated by using multiple regression analysis. Simple correlation analysis revealed a significant negative association of  $ASSCH_2$  with newsprint moisture content (NMO) and pH of the sample. Application of five different variable selection procedures, which are commonly used in multiple regression analyses, showed that NMO, amylase (AMY), esterase (EST), cellulose to lignin ratio (CLR), volatile solids (VS), and nitrogen content (NIT) were significantly associated with  $ASSCH_2$  simultaneously. The other five factors did not show any significant effect on  $ASSCH_2$  in the presence of the six significant factors. Further analysis showed that the influence of AMY and EST on  $ASSCH_2$  was weak, hence they were not included in the regression model. CLR was also deleted from the final model because of the multicollinearity resulting from its high correlation with VS. The final model incorporated NMO,  $NMO^2$ ,  $VS^2$  and  $NIT^2$ ; it explained 95% of the total variability and predicted 98% of the observed  $ASSCH_2$ . An assessment of the relative significance of the independent variables indicated that NMO contributed the most, followed by  $NMO^2$  and  $VS^2$ , in that order, and the least by  $NIT^2$  towards  $ASSCH_2$ . The NMO and  $NIT^2$  showed an inhibitory effect on  $ASSCH_2$ . The results indicated that maintaining optimum moisture, along with optimum organic loading, and nitrogen content in landfills is necessary to achieve and maintain a low  $ASSCH_2$  and maximize refuse methanogenesis.

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**Keywords** – Anaerobic biodegradation; hydrogen; landfill; methane; multiple regression analysis; municipal solid waste; refuse; statistical modeling

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Received 16 July 1999, accepted in revised form 29 January 2000

## Introduction

Landfilling of municipal solid waste (MSW) is a widely practiced means of its disposal. The United States Environmental Protection Agency (USEPA) estimated that about

60% of the MSW produced in 1994 in the United States was disposed of in landfills (US EPA 1996). Approximately one-half of MSW is typically composed of cellulose and hemicellulose (Fairweather & Barlaz 1988; Barlaz *et al.* 1989), which are considered readily degradable in the

environment. It is generally assumed that this organic fraction of the refuse is degraded by indigenous anaerobic microorganisms, because oxygen, if present, is restricted to the uppermost regions in landfills. However, it has been shown that refuse components, which are otherwise considered readily degradable, can sometimes persist for surprisingly long periods of time in landfills owing to several environmental factors that limit microbial activity (Sufliya *et al.* 1992; Gurijala & Sufliya 1993). Landfills designed to overcome such limitations will certainly increase microbial metabolic activity and subsequently allow efficient refuse degradation. Hence, complete understanding of factors affecting microbial metabolism in landfills is of critical importance in designing new landfills.

Several metabolic groups of bacteria carry out a complex series of reactions during the anaerobic degradation of MSW, in which hydrogen and volatile fatty acids (VFA) are produced as transient intermediates (McInerney & Bryant 1980; Harper & Pohland 1986; Mormile *et al.* 1996). The smooth functioning of these ecosystems greatly relies upon the balance between the production and consumption of these intermediates (Kasper & Wuhrmann 1978; Labib *et al.* 1992; Kramer & Conrad 1993). Disturbances to such balance lead to hydrogen and VFA accumulation, which can negatively influence terminal electron accepting processes (TEAP). For instance, hydrogen partial pressures exceeding  $10^{-4}$  atm (80.7 nM of dissolved  $H_2$ ) cause VFA accumulation, while those below  $10^{-4}$  atm favor TEAP (Kasper & Wuhrmann 1978; McInerney & Bryant 1980; Harper & Pohland 1986; Zehnder & Stumm 1988; Krylova & Conrad 1998). VFA accumulation under a high hydrogen atmosphere, in turn, results in a pH decrease that is inhibitory to methanogenic bacteria and methanogenesis (Mosey 1982; Switzenbaum *et al.* 1990; Mormile *et al.* 1996).

Apparent steady-state concentrations of hydrogen ( $ASSCH_2$ ) were used to indicate imbalances of methanogenic fermentations (Strong & Cord-Ruwisch 1995; Mormile *et al.* 1996; Cord-Ruwisch *et al.* 1997) and to determine dominant TEAP in a variety of environments (Lovley *et al.* 1994; Jakobsen *et al.* 1998) because of its key role in anaerobic catabolism (see Background section below). Despite its importance, environmental factors influencing  $ASSCH_2$ , especially in landfill ecosystems, have been rather poorly understood. Available information, however, showed that organic overloading and presence of toxic material inhibit methanogenesis in landfills and digesters via hydrogen and VFA accumulation (McInerney & Bryant 1980; Mosey 1982; Harper & Pohland 1986; Switzenbaum

*et al.* 1990; Strong & Cord-Ruwisch 1995; Mormile *et al.* 1996; Cord-Ruwisch *et al.* 1997). Temperature was the only other environmental variable of which the effect on  $ASSCH_2$  was studied (Westermann 1994). In wetland sediment slurries, hydrogen production increased by a factor of 18 when the incubation temperature was raised from  $2^\circ C$  to  $37^\circ C$ . However, the syntrophic degradation reactions of VFA remained exergonic, once the corresponding hydrogen concentration reached a steady state at the temperatures tested.

Information on the relative significance of multiple environmental factors simultaneously influencing hydrogen production during refuse degradation is not available, although such information is necessary in effectively controlling accelerated hydrogen production. By controlling accelerated hydrogen production, the probability of uncoupling the metabolic reactions necessary for refuse methanogenesis is reduced and methane production can be optimized. Therefore, identification and evaluation of the relative significance of environmental determinants governing  $ASSCH_2$  during refuse degradation, presented in this study, will be extremely useful in designing landfills for improved methane recoveries.

## Background

The anaerobic biodegradation of MSW is carried out by several metabolic groups of bacteria (Zehnder *et al.* 1982; Barlaz 1997). The first group of bacteria hydrolyzes polymers such as cellulose and hemicellulose to their constituent monomers (Colberg 1988; McInerney 1988). These bacteria are the primary fermentative or hydrolytic bacteria (Schink 1997). The resulting monomers are further fermented to organic acids such as VFA and alcohols by the secondary fermentative bacteria or fatty acid oxidizing bacteria (FAOB). The FAOB convert these organic acids and alcohols to acetate, hydrogen and carbon dioxide (Dolfing 1988; Schink 1997).

The terminal group of bacteria complete the mineralization of refuse by converting the products of FAOB to methane and carbon dioxide. The terminal group of bacteria comprise methanogenic bacteria (Oremland 1988), sulfate-reducing bacteria (Widdel 1988), iron/manganese-reducing bacteria (Ghiorse 1988), and nitrate-reducing (Tiedje 1988) bacteria. Methanogenic bacteria can produce methane either from acetate, or from hydrogen and carbon dioxide (Zehnder *et al.* 1982; Thauer 1998; Conrad 1999; Wolfe 1999). All other terminal group bacteria degrade acetate and

other organic carbon at the expense of reducing appropriate electron acceptors. These electron-accepting processes carried out by the terminal group of bacteria are collectively known as TEAP.

Hydrogen is a critical though transient intermediate in refuse degradation. It is produced in the acetogenic dehydrogenation reactions and is consumed by terminal group of bacteria during TEAP. Acetogenic dehydrogenations of organic acids and alcohols in landfill environments are thermodynamically inhibited at hydrogen concentrations above 80.7 nM, which occurs when these anaerobic systems become overloaded with organic carbon. This leads to the accumulation of organic acids, which in turn reduces the pH to a level inhibitory to methanogenic bacteria and the methanogenesis (Mormile *et al.* 1996). For refuse degradation to occur smoothly, the hydrogen concentration must be maintained below 80.7 nM. A higher accumulation of this gas is indicative of fermentation imbalances resulting in the inhibition of methanogenesis from refuse.

## Materials and methods

The data analyzed in this study and the methods used to obtain the data were previously published (Sufliata *et al.* 1992; Gurijala & Sufliata 1993; Mormile *et al.* 1996). Samples of MSW were collected from various sites and depths at the Fresh Kills Landfill, Staten Island, New York, USA, in October 1989. A bucket auger was used to drill down to the desired depth in order to obtain the samples. Fourteen boreholes were drilled in the landfill, and samples were collected at approximately 3-m intervals. The refuse was passed through a coarse sieve (5 × 5 cm) before being collected in plastic buckets containing an O-ring sealing lid. The headspace of the refuse-filled buckets was exchanged with oxygen-free nitrogen prior to their transport by overland courier to the laboratory, where they were stored at room temperature (Sufliata *et al.* 1992; Gurijala & Sufliata 1993).

Anaerobic refuse incubation vessels were constructed by joining a standard plumbing end cap (PVC plastic, 7.6 cm) to a reducing union (7.6 to 5 cm). Refuse material (200 to 300 g) was placed in the vessels while they were inside a portable anaerobic glovebag (AtmosBag, Aldrich Chemical Co., Milwaukee, Wisconsin, USA), which was constantly purged with nitrogen. The head space of each vessel was initially oxygen-free nitrogen. The vessels were incubated at room temperature and the headspace methane and hydrogen

were monitored at regular intervals (Gurijala & Sufliata 1993; Mormile *et al.* 1996).

All statistical analyses were performed on the main frame computer at the University of North Florida using MS Windows Network version of the SAS program (SAS Institute 1990). A multiple regression model that explained the vast majority of variation in and predicted most of the observed ASSCH<sub>2</sub> (dependent variable) was developed (Gurijala *et al.* 1997). The independent variables significantly correlated with ASSCH<sub>2</sub> from the refuse samples were selected by the application of forward selection, backward elimination, step-wise procedure, coefficient of determination ( $R^2$ ) and conceptual predictive (Cp-statistic) criteria, with  $\alpha$  values set at 0.1 to 0.5 (Mallows 1964; Neter *et al.* 1996). Examining  $p$ -values of the  $t$ -statistics refined the selection of variables.

Once the independent variables significantly associated with ASSCH<sub>2</sub> were selected, a second-order model was developed on the basis of partial regression plots and the residual plots. SAS procedure 'PROC UNIVARIATE', with 'NORMAL' option, Box-Cox transformations (Box & Cox 1964) and the residual plots revealed that the data did not need any transformation to satisfy the normality and equal variance assumptions. Each of the variables in the second-order model was then coded to reduce multicollinearity as explained by Neter *et al.* (1996).

All possible linear, square, and cross-product terms of the independent variables in the second-order model were analyzed again on the basis of forward selection, backward elimination, step-wise procedure,  $R^2$  criterion, and Cp statistic and by the examination of partial regression leverage plots and the  $p$ -values of  $t$ -statistics. Then the variables in the model were tested for possible outliers and multicollinearity by variance inflation factors (Cook 1979) and all the standard outlier-detection procedures provided by the SAS procedure 'PROC REG' with 'INFLUENCE' option (SAS Institute 1990).

The assumptions of the second-order model, namely the random errors ( $\epsilon_i$ ) are independently and normally distributed with mean zero [ $E(\epsilon_i) = 0$ ] and a constant variance [ $V(\epsilon_i) = \sigma^2$ ], were tested by using univariate procedure, residual plots, normal probability plots and Box-Cox transformations.

The model parameter estimates were computed by using the least-squares method (Masili-Libelli 1992; Saez & Rittmann 1992). Then the coded variables were transformed to their original scale to derive the final fitted model.

Comparing the coefficients of partial determinations

(*R*-values), *t*-statistics and *p*-values of the model parameter estimates (Neter *et al.* 1996; Gurijala *et al.* 1997) determined the relative contribution of each model variable towards ASSCH<sub>2</sub>. The *R*-value for each independent variable *X* was computed as follows (Neter *et al.* 1996):

$$R = \frac{[SSR_{(X|all\ the\ other\ variables)}]}{[SSE_{(all\ the\ other\ variables)}]} \quad [1]$$

where:

$[SSR_{(X|all\ the\ other\ variables)}]$  = the extra sum of squares of adding *X* to the model when all the other variables are already included and is given by  $[SSE_{(all\ the\ other\ variables)}] - [SSE_{(all\ the\ variables)}]$

$[SSE_{(all\ the\ other\ variables)}]$  = error sum of squares for the model with all the other independent variables except *X*

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## Results

The descriptive statistics and the definition of the variables were given in Table 1. The ASSCH<sub>2</sub> in refuse incubations varied greatly with a mean of 6.4 μM and a standard deviation of 14.0 μM. Pearson correlations indicated that pH and NMO are significantly, albeit negatively, associated with the ASSCH<sub>2</sub>. None of the other variables considered showed any significant relationship with the ASSCH<sub>2</sub> (Table 1). Because Pearson correlations are insufficient to determine simultaneous effects of multiple variables, landfill data was analyzed by multiple regression techniques.

In order to develop a multiple regression model from the landfill data, it was necessary to first identify significant variables simultaneously influencing ASSCH<sub>2</sub>. The significant variables out of 11 landfill characteristics (Table 1) were selected by using five different statistical methods. The forward selection procedure identified CLR, NMO, AMY, NIT and EST to be significantly associated with ASSCH<sub>2</sub> (Table 2). The backward elimination procedure selected NMO and CLR, whereas the step-wise procedure showed CLR, NMO and AMY to be significant factors associated with ASSCH<sub>2</sub>.

The *R*<sup>2</sup>-values and Cp statistics of models containing all possible combinations of independent variables further refined the variable selection (Table 3). Thus, models C, D and E showed the highest *R*<sup>2</sup> and lowest Cp. Further increase in the number of variables neither increased *R*<sup>2</sup> nor decreased Cp values substantially. Therefore, models C, D and E were selected to be adequate starting points in the development of the final model.

The Pearson correlation matrix revealed a strong linear relationship between CLR and VS (*r*<sup>2</sup> = 0.9388). In addition, the variance inflation factors (VIF) for VS (35.516) and CLR (39.146) far exceeded their tolerance levels (0.028 for VS and 0.026 for CLR) indicating multicollinearity between these two factors. As deletion of one of the two variables effectively eliminates multicollinearity, the models 1 and 2 (see below) containing CLR or VS were individually evaluated to determine which of the two variables should be deleted.

Partial regression plots and *t*-tests suggested that the effects of AMY and EST on ASSCH<sub>2</sub> in the presence of the

Table 1. Descriptive statistics and definition of the variables\*

Variable	Definitions of variables (units)	N	Mean	Standard deviation	Correlation coefficient†
ASSCH <sub>2</sub>	Apparent Steady State Concentration of Hydrogen (μM)	30	6.4	14.0	-
MR	Methane Production Rate (μmol dry kg <sup>-1</sup> d <sup>-1</sup> )	30	198.7	182.1	-0.30
NMO	Newsprint Moisture Content (%wt dry wt <sup>-1</sup> )	25	55.374	13.732	-0.38‡
MO	Moisture Content (% wt dry wt <sup>-1</sup> )	30	36.660	14.810	-0.31
SO <sub>4</sub>	Sulfate (mmol dry kg <sup>-1</sup> )	30	18.034	16.574	0.08
pH	pH of water extracts	30	6.977	0.535	-0.55‡
PRO	Protease (units h <sup>-1</sup> dry g <sup>-1</sup> )	27	0.039	0.153	-0.04
AMY	Amylase (nmol h <sup>-1</sup> dry g <sup>-1</sup> )	27	1.020	0.747	-0.03
EST	Esterase (nmol h <sup>-1</sup> dry g <sup>-1</sup> )	26	16.714	11.336	-0.14
VS	Volatile Solids (%wt dry wt <sup>-1</sup> )	22	37.768	19.202	0.13
NIT	Nitrogen (mg dry kg <sup>-1</sup> )	22	4456.9	1238.5	0.21
CLR	Cellulose to Lignin Ratio	22	1.594	0.806	-0.15

\*Data from Gurijala *et al.* (1997), Mormile *et al.* (1996), and Sufliita *et al.* (1992).

†Pearson correlation coefficient between ASSCH<sub>2</sub> and the variable considered.

‡Significant at α = 0.05.

Table 2. Selection of variables\* associating with ASSCH<sub>2</sub> by forward selection, backward elimination, and stepwise procedures

Selection method	Variable	p-value of F statistic
Forward Selection ( $\alpha = 0.5$ )	CLR	0.0053
	NMO	0.0025
	AMY	0.1437
	NIT	0.2615
	EST	0.3917
Backward elimination ( $\alpha = 0.1$ )	CLR	0.0014
	CLR	0.0053
	NMO	0.0025
Stepwise procedure ( $\alpha = 0.15$ )	AMY	0.1437

\*See Table 1 for the definition of variables.

other variables were insignificant. Thus, model C was selected and model D was reduced to model B in the further development of the final fitted model (Table 3).

The residual plots and the partial regression plots suggested the addition of higher-order (square, cross-product, or both) terms of the significant independent variables to the model to better account for the variation observed in ASSCH<sub>2</sub>. Therefore, the higher-order terms of the significant independent variables were tested for inclusion in the model and were confirmed by the examination of residual and partial regression plots. The variables thus selected were included in the models 1 and 2 after being coded by subtracting the mean and dividing by the standard deviation. The coding effectively eliminated multicollinearity among the higher power terms as indicated by the Pearson correlation matrix and VIF.

$$\text{ASSCH}_2 = \beta_0 + \beta_1 \text{TNMO} + \beta_2 \text{TNMO}^2 + \beta_3 \text{TVS}^2 + \beta_4 \text{TNIT}^2 + \varepsilon \quad [2]$$

$$\text{ASSCH}_2 = \beta_0 + \beta_1 \text{TNMO} + \beta_2 \text{TNMO}^2 + \beta_3 \text{TCLR} + \beta_4 \text{TCLR}^2 + \varepsilon \quad [3]$$

where:

TVS<sup>2</sup>, TNMO<sup>2</sup> and TNIT<sup>2</sup> are coded square terms of VS, NMO, and NIT, respectively

TNMO is coded NMO

$\beta_0$  to  $\beta_4$  are the parameters

$\varepsilon$  is the random error of the model.

The assumptions involved in model 1, which contained VS, were reasonable. The assumption that the observations are independent is true because the plots of observed ASSCH<sub>2</sub> production vs. the residuals did not show significant patterns (Fig. 1). Analysis using SAS UNIVARIATE procedure with the NORMAL option revealed that the assumption of observations being normally distributed was also reasonable. The assumption that the mean of the random error is zero [ $E(\varepsilon_i) = 0$ ] cannot be tested, because the sum of residuals always equals zero. However, this assumption was considered reasonable because all the other assumptions were true (Neter *et al.* 1996). The equal variance assumption was tested by separating the observations into two groups by level of independent variables, and by comparing the sample variances of each half by an *F*-test. The *F*-test showed that there is no significant difference between the variances of the two halves of the data set indicating that the equal variance assumption is valid.

Partial regression plots of model 1 showed significant association between the independent variables and ASSCH<sub>2</sub> (Fig. 2). This model also explained 95% of the variation in observed ASSCH<sub>2</sub> as indicated by its *R*<sup>2</sup> value (0.9549). In addition, a scatter plot of observed vs. predicted values of

Table 3. Selection by *R*<sup>2</sup> criterion and *C*<sub>p</sub> statistic of variables\* associating with ASSCH<sub>2</sub>

Model	Variable(s)	(P)†	<i>R</i> <sup>2</sup>	<i>C</i> <sub>p</sub>
A	CLR	2	35.796	4.82
B	NMO, CLR	3	63.097	-2.03
C	NMO, VS, NIT	4	69.399	-2.07
D	NMO, AMY, CLR	4	67.847	-1.57
E	NMO, AMY, VS, NIT	5	71.643	-0.802
F	NMO, pH, VS, NIT	5	71.247	-0.674
G	NMO, pH, AMY, VS, NIT	6	72.813	0.818
H	NMO, pH, AMY, EST, VS, NIT	7	73.252	2.675
I	MR, NMO, MO, SO <sub>4</sub> , pH, VS, NIT	8	74.315	4.331
J	MR, NMO, MO, SO <sub>4</sub> , pH, AMY, NIT, CLR	9	75.098	6.077
K	MR, NMO, MO, SO <sub>4</sub> , pH, AMY, VS, NIT, CLR	10	75.330	8.002
L	MR, NMO, MO, SO <sub>4</sub> , pH, PRO, AMY, VS, NIT, CLR	11	75.334	10.000
M	All 11 variables	12	75.336	12.000

\*See Table 1 for the definition of variables.

†Number of parameters include the intercept

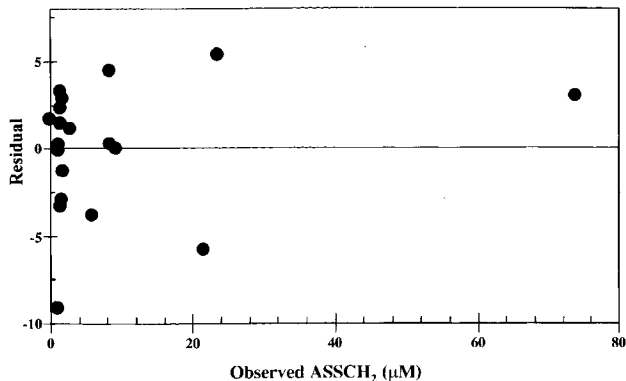


Fig. 1. Plot of residuals vs. observed apparent steady-state concentrations of hydrogen (ASSCH<sub>2</sub>) showing the variance is independent and constant.

ASSCH<sub>2</sub> showed that model 1 predicted 98% ( $r^2 = 0.9795$ ) of the observed ASSCH<sub>2</sub> (Fig. 3).

On the other hand, the inherent assumptions for model 2, which included CLR, were found to be invalid. In addition, model 2 was less efficient than model 1 in explaining total variation in and predicting observed ASSCH<sub>2</sub>. Hence, model 1 with VS was used in the evaluation of relative significance of significant variables associated with ASSCH<sub>2</sub>.

The relative importance of the significant factors affecting ASSCH<sub>2</sub> during refuse degradation was evaluated from model 1. The coded variables were decoded to their original scale before calculating least-square estimates of the parameters. The *t*-statistics and *p*-values of the parameter estimates, and the *R*-values, indicated the relative contribu-

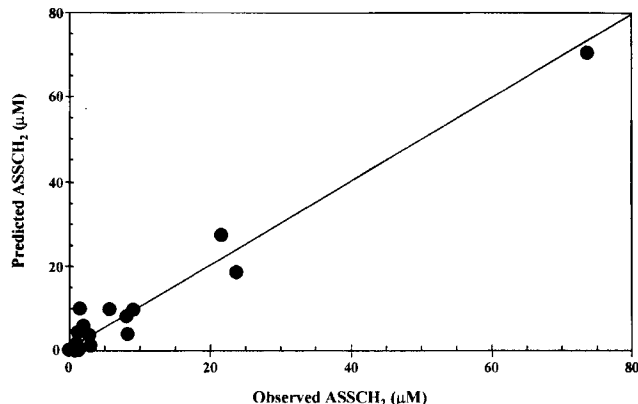


Fig. 3. The relationship between predicted and observed apparent steady-state concentrations of hydrogen (ASSCH<sub>2</sub>). Line of equality shown.

tion of the independent variables present in model 1 (Table 4). The higher the  $|t|$  value, the higher the *R*-value, and the lower the *p*-value the more the contribution of that independent variable. Thus, NMO contributed the most, followed by NMO<sup>2</sup> and VS<sup>2</sup> and the least by NIT<sup>2</sup> towards ASSCH<sub>2</sub>. The parameter estimates indicated that the influence of VS<sup>2</sup> and NMO<sup>2</sup> was positive, while that of NMO and NIT<sup>2</sup> was negative on ASSCH<sub>2</sub> (Table 4).

### Discussion

Environmental factors contributing towards the accumulation of hydrogen during anaerobic refuse degradation will

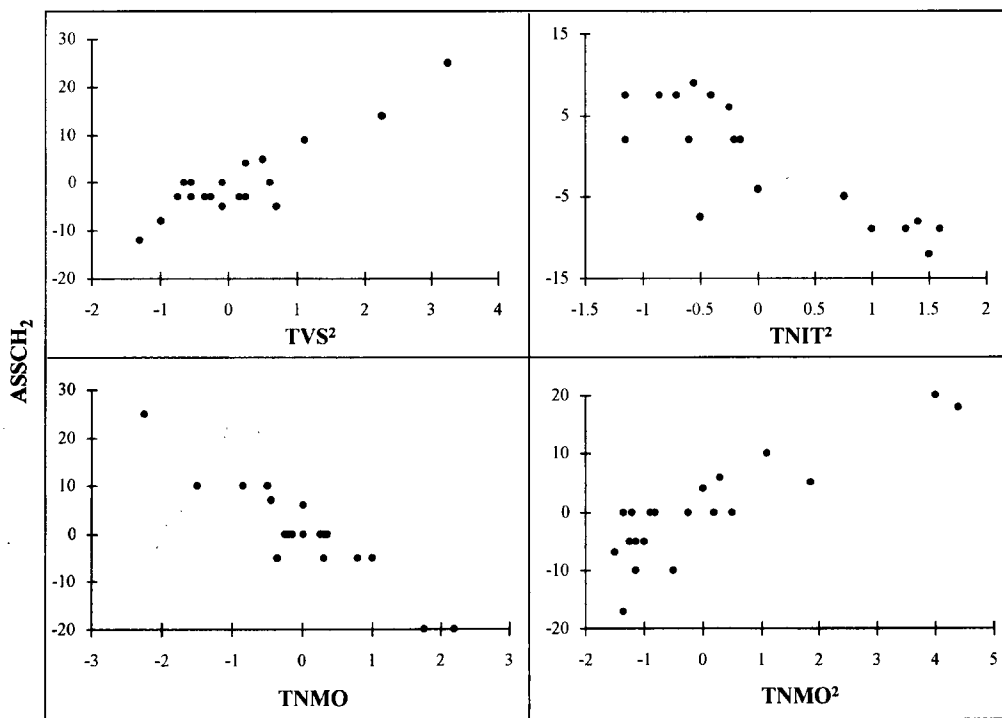


Fig. 2. Partial regression plots showing the relationship between apparent steady-state concentrations of hydrogen (ASSCH<sub>2</sub>) and independent variables in the final fitted model. See Table 1 for the description and units of x-axis labels.

Table 4. Parameter estimates and coefficients of partial determination of independent variables in the final model

Variable	Parameter	Parameter estimate*	††	p‡	R-Value§
Intercept	$\beta_0$	4.394	2.584	0.0207	–
NMO	$\beta_1$	– 8.906	– 9.748	0.0001	0.7832
NMO <sup>2</sup>	$\beta_2$	4.207	7.758	0.0001	0.7396
VS <sup>2</sup>	$\beta_3$	5.835	6.359	0.0001	0.6158
NIT <sup>2</sup>	$\beta_4$	– 6.483	– 6.392	0.0001	0.5093

\*Parameters were calculated by using least-squares method.

†The *t*-statistic.

‡The *p*-value arising from a *t*-test of the null hypothesis that  $\beta = 0$ .

§The coefficient of partial determination.

also inhibit methanogenesis, as the accumulation of the former clearly indicates fermentation imbalances leading to the inhibition of the latter in a variety of ecosystems (Strong & Cord-Ruwisch 1995; Mormile *et al.* 1996; Cord-Ruwisch *et al.* 1997). Pearson correlation analysis was performed on landfill data to identify significant environmental factors affecting ASSCH<sub>2</sub>. This analysis revealed that pH had significant negative influence on ASSCH<sub>2</sub> indicating that the samples producing hydrogen at high concentrations are often acidic. Mormile *et al.* (1996) have shown that acidic refuse samples accumulated more hydrogen than those with neutral or alkaline pH and that this acidity is due to the accumulation of VFA. The negative relationship between ASSCH<sub>2</sub> and pH is due to VFA accumulation, which is caused by a hydrogen accumulation in excess of 10<sup>–4</sup> atm (Thauer *et al.* 1977; McInerney & Bryant 1980; Zehnder & Stumm 1988; Krylova & Conrad 1998).

However, Pearson correlations are insufficient to assess the relative significance of multiple factors simultaneously affecting ASSCH<sub>2</sub>. Therefore, multiple regression analysis, which essentially involved development of a second-order regression model, was used to evaluate the relative significance of 11 landfill variables. The regression model incorporated significant linear, square, and/or cross-product terms of NMO, NIT, and VS, and successfully captured the trends of the data set as shown by its ability to account for 95% of the total variation in and predicted 98% of the observed ASSCH<sub>2</sub>. Mathematical and kinetic models have been developed from data obtained in experiments, which involved artificially manipulated variables, mainly for the purpose of understanding the physiological basis of organic carbon degradation (Belevi & Baccini 1989; El-Fadel *et al.* 1989; Young 1989). Statistical models constructed from the field data, such as the one developed in this study, overcome limitations inherent to mathematical and kinetic models in the evaluation of relative importance of environmental factors. In fact, multiple regression models developed from

landfill data have been successfully used to determine the relative importance of environmental factors in refuse methanogenesis (Gurijala *et al.* 1997) and to evaluate simultaneous effects of pH and temperature to predict acetate and butyrate production from cheese-processing wastewater (Hwang & Hansen 1997).

The evaluation of relative contributions of significant variables in the regression model revealed that ASSCH<sub>2</sub> was greatly influenced by NMO followed by NMO<sup>2</sup>, VS<sup>2</sup>, and the least by NIT<sup>2</sup>. The influence exerted by NMO, and NIT<sup>2</sup> was negative, while that of NMO<sup>2</sup> and VS<sup>2</sup> was positive on ASSCH<sub>2</sub>. The factors negatively influencing ASSCH<sub>2</sub> should also enhance refuse methanogenesis and vice versa, because non-methanogenic samples accumulated high amounts of hydrogen (Mormile *et al.* 1996). Thus, NMO and NIT<sup>2</sup> should favor, while NMO<sup>2</sup> and VS<sup>2</sup> should inhibit refuse methanogenesis. However, moisture content of the total refuse sample, rather than that associated with only newsprint (the NMO) affects refuse methanogenesis more significantly than any other variable tested (Gurijala *et al.* 1997). This suggested that while total moisture content of the sample enhanced the activity of the methanogens, the NMO enabled methanogens as well as those involved in other TEAP to remain active hydrogen scavengers. In either case it is not surprising that moisture positively affected hydrogen scavenging as well as methanogenic activities, inasmuch as moisture is an essential requirement for landfilled refuse fermentations (Barlaz *et al.* 1990; Senior *et al.* 1990; Gurijala & Suflita 1993; Gurijala *et al.* 1997). The negative contribution of NIT<sup>2</sup> towards ASSCH<sub>2</sub> indicated that anaerobic decomposition of nitrogen-rich refuse samples maintains low ASSCH<sub>2</sub> and high methane production. Again, this is not unexpected as nitrogen is an essential nutrient for efficient biodegradation and bioassimilation of organic carbon (Stevenson 1986; Burton & Watson-Craik 1996).

The positive influence of VS<sup>2</sup> and NMO<sup>2</sup> showed that

high concentration of VS and NMO enhanced hydrogen accumulation and so inhibited methanogenesis. This inhibitory effect of high VS content on refuse methanogenesis was also observed previously (Gurijala *et al.* 1997). Moreover, the inhibition of methanogenesis in anaerobic environments overloaded with organic carbon as evidenced by hydrogen accumulation is widely known (Mosey 1982; Harper & Pohland 1986; Senior *et al.* 1990; Switzenbaum *et al.* 1990; Mormile *et al.* 1996; Gurijala *et al.* 1997). In fact, hydrogen partial pressures exceeding  $10^{-4}$  atm thermodynamically inhibit VFA oxidation (Thauer *et al.* 1977; Zehnder & Stumm 1988; Krylova & Conrad 1998), which in turn inhibit methanogenesis by decreasing pH below the tolerance levels of methanogenic bacteria (Senior *et al.* 1990; Mormile *et al.* 1996). The high NMO content may also have contributed towards organic overloading by exposing cellulose and starch contained in the newsprint to landfill bacteria (Smook 1989). The results presented here provide statistical evidence for the hydrogen and VFA involvement in the inhibition of refuse methanogenesis in landfills.

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As with any statistical model, the regression model presented here is specific for Fresh Kills landfill. However, the model clearly identifies moisture levels as measured by newsprint moisture content, volatile solids, and nitrogen content as significant impactors of landfill ASSCH<sub>2</sub>. The statistical model developed here shows that the combination of optimum moisture levels, avoidance of organic overloading, and sufficient nitrogen content are required to control the overproduction of hydrogen, and ultimately increase methane yields from refuse degradation in landfills.

## Acknowledgments

We thank J. M. Sufita, F. Concannon, R. K. Ham, M. R. Norman, P. R. Fritsche, A. C. Palmisano, B. S. Schwab, D. A. Maruscik, W. L. Rathje and W. W. Hughes for their contributions to the original landfill data. We also thank M. A. Barlaz, M. A. Hamilton and H. A. Cash for reviewing the manuscript.



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