

INFLUENCE OF SPATIAL HETEROGENEITY ON ESTIMATES  
OF CONCENTRATION AND SPECIES DIVERSITY  
OF PELAGIC NET ZOOPLANKTON

By

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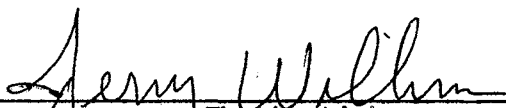
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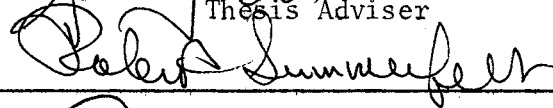
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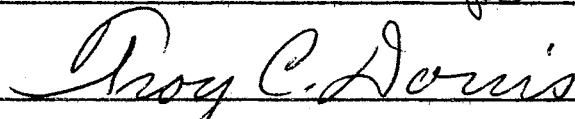
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Thesis Approved:



Thesis Adviser







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## PREFACE

The objective of this study was to determine if traditional sampling procedures adequately result in estimates of species diversity and concentration of pelagic net zooplankton. A mathematical model derived from information theory was used to estimate species diversity and an analysis of variance was done using the F statistic at the .05 level.

I thank my major adviser, Dr. Jerry Wilhm, and Drs. Robert C. Summerfelt and Troy C. Dorris, members of my advisory committee, for their advice and criticisms.

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## CHAPTER I

### INTRODUCTION

Investigators of spatial variations of zooplankton population parameters in lakes typically collect single or replicate samples from a single site at each station. Applegate and Mullan (1967) used oblique tows of a plankton net and a pump to determine zooplankton populations in Beaver and Bull Shoals lakes on the White River in Arkansas. They used a single site at each of their 14 stations. Zooplankton of Lewis and Clarke Lake on the Missouri River at the border of Nebraska and South Dakota were collected with oblique and horizontal hauls of plankton nets at single sites at each of ten stations (Cowell, 1970). Hammer and Sawchyn (1968) in a study of Diaptomus spp. in ponds in Canada sampled a single site and one depth in each of the ponds studied. Horizontal zooplankton distribution was studied in Finger Lakes of New York using plankton net hauls at single sites in the deeper part of the lakes (Hall and Waterman, 1967). Peterka (1970) studied relationships between water quality and plankton in Ashtabula Lake, North Dakota, using single site, single replicate samples for quantitative estimates of zooplankton. In all of these examples, it was assumed that the samples were representative of the aquatic environment at the station and the station represented the environment in the area.

Nonrandom spatial distribution or patchiness of zooplankton may bias estimates of population parameters (Wiebe, 1971; Wiebe, 1970; Wiebe and

Holland, 1968; Barnes and Marshall, 1951). Wiebe and Holland (1968) used a computer model to study the effect of changes in the size of patches, the distribution of the centers of patches, and net size on estimates of the mean number of individuals and variability about the mean. They found that size of patches affects both accuracy and precision of the estimates. Significant patchiness existed among marine zooplankton (Wiebe, 1970). Significant changes in coefficients of variability can occur as a function of patchiness size and distribution (Wiebe, 1971). Barnes and Marshall (1951) found variations in the type of dispersion among nauplii of marine copepods from underdispersed to overdispersed.

Species diversity indices have been used in recent years to characterize community structure. Patten (1962) used the following formula derived from information theory (Shannon and Weaver, 1963) to characterize plankton in Raritan Bay:

$$\bar{d} = \sum_{i=1}^s \left( \frac{n_i}{n} \right) \log_2 \left( \frac{n_i}{n} \right) ,$$

where  $s$  is the number of species,  $n_i$  is the number of individuals in the  $i$ th species, and  $n$  the total number of individuals. Since Patten's work, numerous studies of aquatic communities have included estimates of species diversity using the above formula (Wilhm, 1968; Kochsiek and Wilhm, 1970; Kochsiek, et al., 1971; Prather and Prophet, 1969; Cooper, 1972, and Staub, et al., 1970). This equation is dimensionless, independent of sample size, and expresses the relative importance of each species (Wilhm and Dorris, 1968).

Patchiness may result in considerable variation in total number of organisms, numbers within each species, and numbers of species. Since



numbers of organisms and species are basic data in species diversity indices, indices calculated from data taken from a single site may not be representative of the entire area. The objective of this study was to determine if spatial variation in species diversity and abundance of pelagic net zooplankton among samples at a particular sampling station caused statistically biased results.

The area selected for study was Eufaula Lake, part of the Arkansas River Basin Development Plan. The purposes of the lake and dam are flood control, power generation, and storage of water for conservation and silt deposition (Table I). The earth-filled dam is located on the Canadian River in McIntosh and Haskell counties, Oklahoma (Figure 1) at  $35^{\circ} 18' 20''$  N -  $95^{\circ} 21' 36''$  W. Shore line length is 965.61 km at power pool level with a mean depth of 7.07 m and shore development of 13.4. Final closure occurred in February, 1964.

TABLE I  
EUFAULA LAKE DATA

Feature	Elevation (m, sea level)	Area (Hectares)	Capacity (m <sup>3</sup> x 10 <sup>4</sup> )
Top of dam	186.54	----	----
Top of gates	181.97	57,870.67	474,650.80
Top of power pool	178.31	41,480.72	293,326.30
Spillway crest	172.21	19,425.12	11,064.50
Stream bed at dam	152.40	----	----

Calculated using data from U. S. Army, Corps of Engineers.

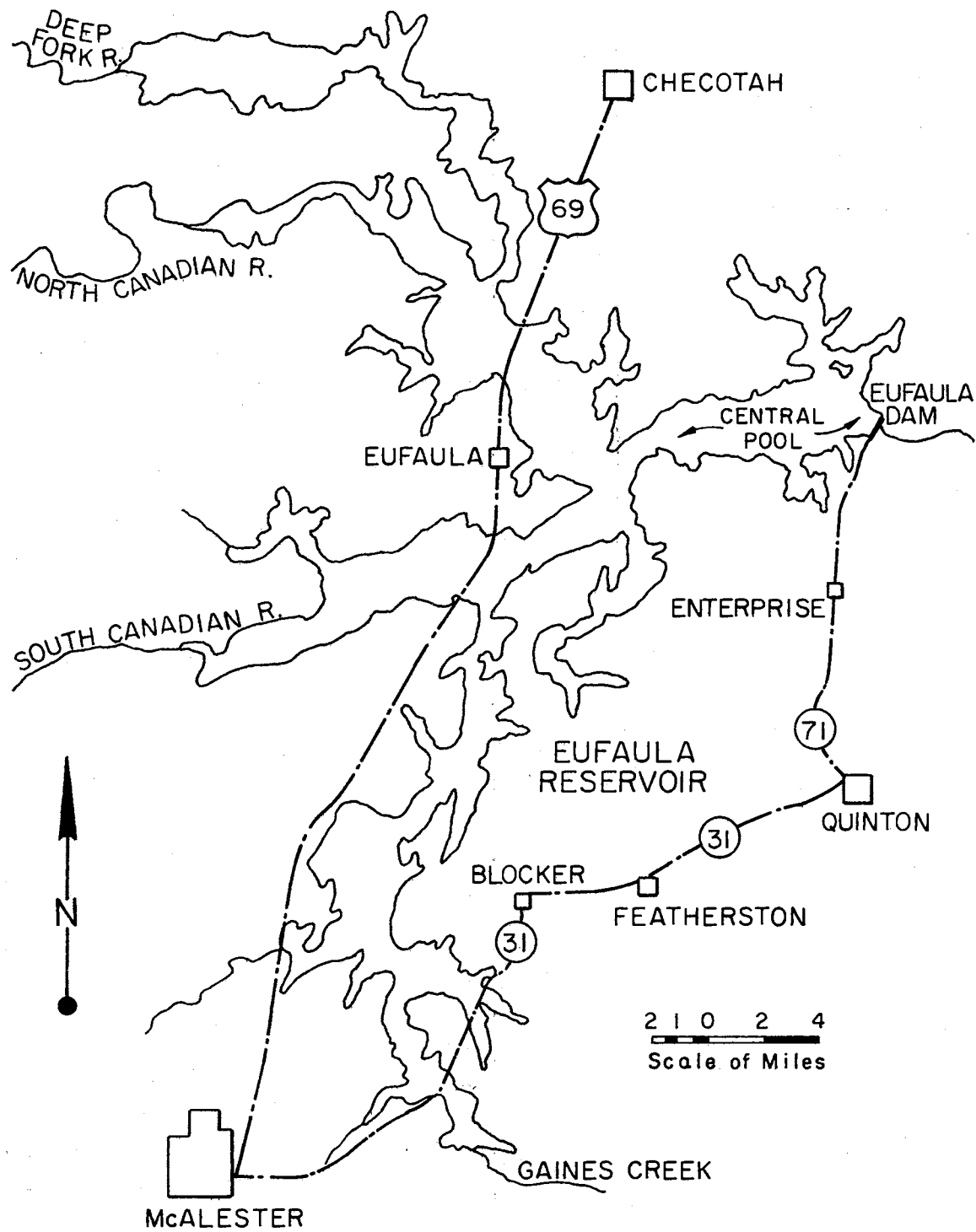


Figure 1. Area Location Map of Eufaula Lake, Oklahoma

## CHAPTER II

### METHODS

Four stations were established approximately 2 km apart in the central pool of Eufaula Lake (Figure 2), 4 - 6 August, 1969. All samples were taken between the hours of 1000 and 1500. At each station nine sites were specified in three rows and three columns (Figure 3). The center site was marked by a carboy bottle buoy. The other eight sites were located from the center site by using a magnetic compass and measured nylon line. Four vertical hauls were taken from the bottom to the surface at each site with a Wisconsin plankton sampler fitted with a size 20 net. Thus, a total of 36 samples were obtained from each station. The volume of water sampled was calculated by multiplying the depth of the water sampled by the area of the mouth of the net.

Three samples from each site were analyzed with the fourth held in reserve to be used if a sample were accidentally destroyed. Samples were diluted or concentrated to a specified volume and subsamples removed and analyzed in a counting chamber under a wide-field dissecting microscope (Ward, 1955; Cowell, 1967).

It has been shown that as samples are pooled,  $\bar{d}$  increases rapidly at first and then approaches an asymptote (Pielou, 1966; Wilhm and Dorris, 1968; and Kochsiek, et al., 1971). Kochsiek (1971) successively pooled counts of 100 zooplankters and found that for pelagic zooplankton, a sample size of 400 provided adequate precision in estimating  $\bar{d}$  with

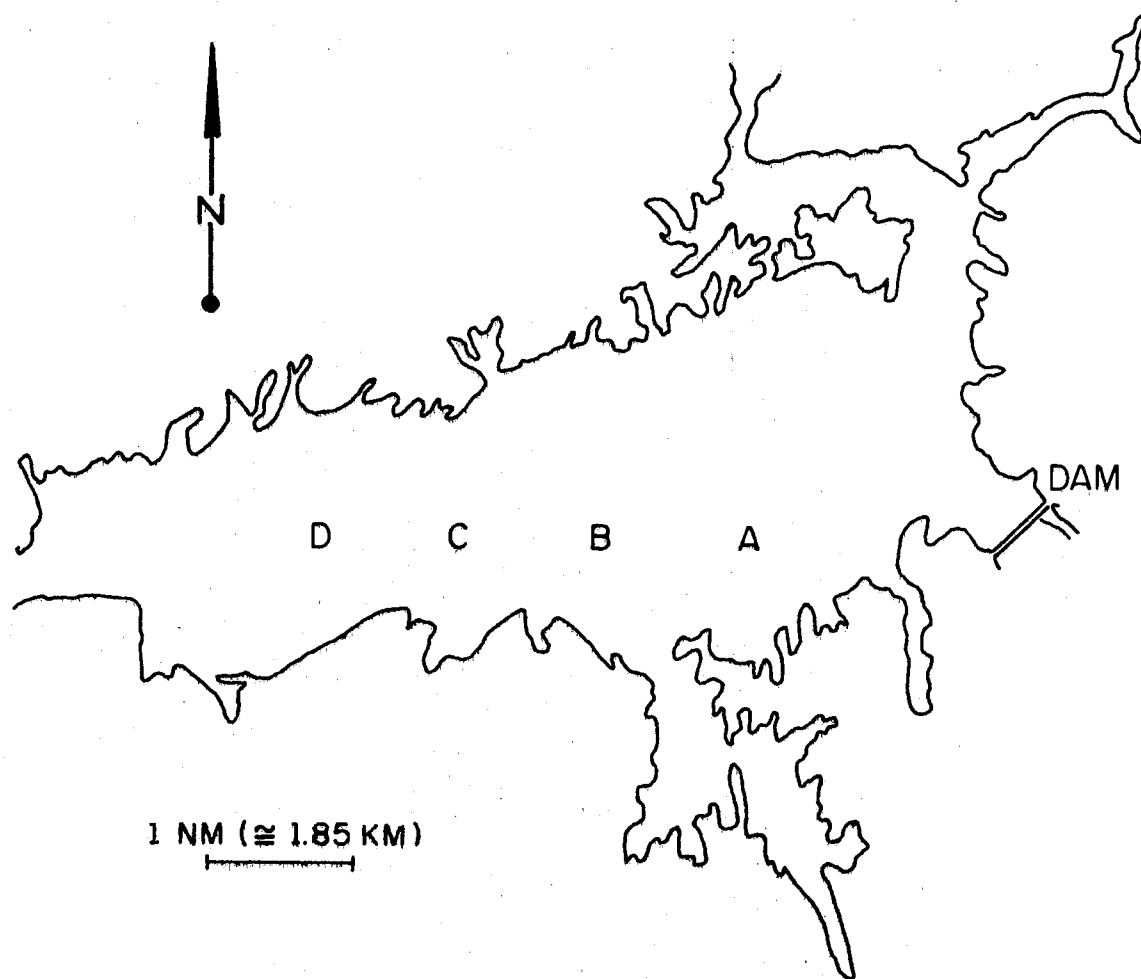


Figure 2. Shoreline Map of the Main Pool of Eufaula Lake  
Showing Location of Sampling Stations

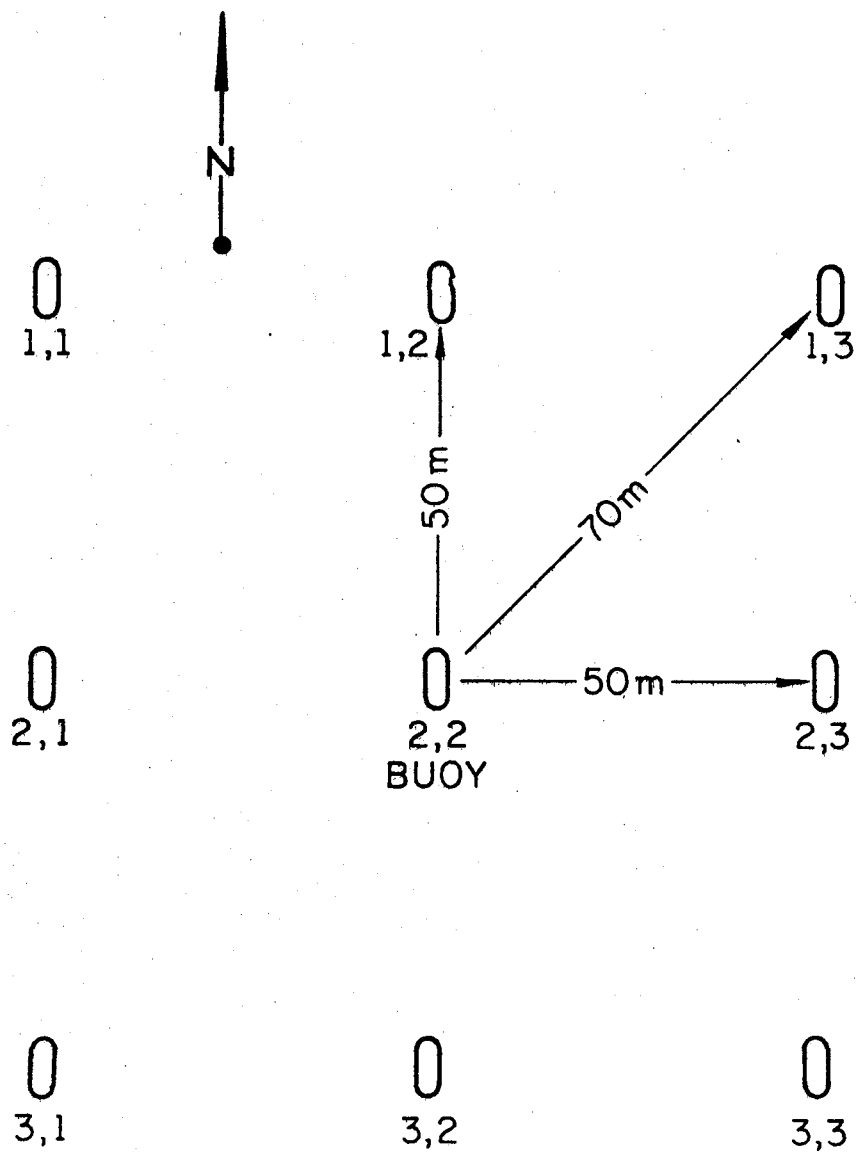


Figure 3. Diagram Showing Arrangement of Sites at Each Sampling Station

insignificant gains above 400. Cooper (1972) pooled counts of 100 to a total of 1000 and found 100 to be adequate for estimating periphyton diversity in a stream. Prather and Prophet (1968) obtained similar results on the basis of pooling 1 ml subsamples and found one subsample sufficient to estimate  $\bar{d}$ . In the present study, 400 zooplankters were identified to species and analyses made of number of species, number of individuals per species, and total number of individuals. Appropriate calculations were made to determine concentration and the equation derived from information theory was used to calculate species diversity.

Variation within stations was analyzed using the analysis of variance (Steel and Torrie, 1960). The F statistic was used to test the null hypothesis that  $\bar{d}$  and concentration estimates are not affected by spatial heterogeneity (patchiness) of pelagic net zooplankton. The alternative was that significant variations would occur, indicating patchiness affects.

## CHAPTER III

### RESULTS

A total of 23 species of zooplankton were found (Table II). Occurrence of species indicates numerical dominance shared by only a few species. Nine species and immature copepoda composed approximately 99%, and five species and immature copepoda composed approximately 80%, of the total number of individuals counted. Pennak (1946) concluded, after examining data from 24 Colorado lakes and 22 lakes from other parts of the world, that most limnetic zooplankton communities are characterized at any one time by one numerically dominant species of Cladocera, Copepoda, and Rotifera.

Mean concentration estimates (Table III) of the nine most numerous species and immature copepoda ranged from less than  $0.1 \text{ liter}^{-1}$  for Asplanchna sp. at station B to 11.3 for immature copepoda at station A. The most abundant species were Ceriodaphnia lacustris with a concentration of  $7.5 \text{ liter}^{-1}$  at station D. C. lacustris and Keratella valga increased and Polyarthra vulgaris decreased slightly in abundance from east to west. Other species showed less or no longitudinal variation.

Mean zooplankton concentrations in Keystone Lake in August were two to four times greater than obtained in this study (Kochsiek, 1971). Rotifers were more abundant in Keystone than in Eufaula, perhaps attributable to more stream-like conditions existing in eight of the nine stations in Keystone Lake as compared to the expansive open water



TABLE II

SPECIES LIST AND FREQUENCY OF OCCURRENCE IN THE 108  
 SAMPLES OF PELAGIC NET ZOOPLANKTON, MAIN POOL,  
 EUFAULA LAKE, AUGUST, 1969

	Occurrence (max. = 108)
Cladocera	
<u>Daphnia ambigua</u>	1
<u>Daphnia galeata</u>	29
<u>Daphnia parvula</u>	30
<u>Daphnia pulex</u>	8
<u>Diaphanasoma leuchtenbergianum</u>	108
<u>Ceriodaphnia lacustris</u>	108
<u>Bosmina longirostris</u>	108
Copepoda	
Immature copepoda	108
Calanoida	
<u>Diaptomus clavipes</u>	108
Cyclopoida	
<u>Mesocyclops edax</u>	108
<u>Ectocyclops phaleratus</u>	1
Rotifera	
<u>Keratella valga</u>	108
<u>Keratella longispina</u>	1
<u>Filinia longiseta</u>	9
<u>Trichocerca multicrinus</u>	1
<u>Trichocerca similis</u>	24
<u>Trichocerca longiseta</u>	1
<u>Hexarthra sp.</u>	108
<u>Asplanchna sp.</u>	108
<u>Platylabus patulus</u>	9
<u>Polyarthra vulgaris</u>	108
<u>Kurzia latissima</u>	1
<u>Monostyla quadridentata</u>	1
Coelenterata	
<u>Craspedacusta sowerbyi</u>	1

TABLE III  
 MEAN CONCENTRATIONS IN NUMBER LITER<sup>-1</sup> OF PELAGIC  
 NET ZOOPLANKTON, MAIN POOL, EUFAULA LAKE,  
 AUGUST, 1969

	Station			
	A	B	C	D
<u>Diaphanasoma leuchtenbergianum</u>	2.8	2.4	3.3	2.6
<u>Ceriodaphnia lacustris</u>	2.8	4.7	6.0	7.5
<u>Bosmina longirostris</u>	0.4	0.5	0.6	0.4
Immature copepoda	11.3	6.1	11.0	10.4
<u>Diaptomus clavipes</u>	5.3	3.7	3.5	4.2
<u>Mesocyclops edax</u>	1.1	0.9	1.5	1.3
<u>Keratella valga</u>	0.7	0.1	2.0	2.1
<u>Hexarthra</u> sp.	0.2	0.1	0.3	0.4
<u>Asplanchna</u> sp.	0.3	< 0.1	0.3	0.2
<u>Polyarthra vulgaris</u>	2.6	0.5	0.5	0.2
Others	0.1	> 0.6	1.2	0.2

environment existing in the main pool of Eufaula Lake.

Estimates of  $\bar{d}$  ranged from 2.16 to 2.96, within station D. Kochsiek, et al. (1971) found  $\bar{d}$  values around 2.9 for zooplankton in August, 1967, in Keystone Lake, and Prather and Prophet (1969) obtained values ranging from 1.83 to 3.06 for zooplankton found during the summer of 1968 in three small lakes in Kansas (Table IV).

Variation among stations of concentration was apparent, but little variation existed among stations in  $\bar{d}$  and number of species. Mean concentration ranged from 19.7 at station B to 30.2 at station C, while mean number of species ranged from 10.6 at stations A, B, and D to 11.1 at station C. Mean  $\bar{d}$  ranged from 2.50 at station A to 2.54 at stations B and D. Variation of concentration within a station was greatest at station C and least at station B.

Significant variation existed at station A for  $\bar{d}$  and station B for concentration (Table V). The coefficient of variation (CV) of concentration estimates ranged from 13.55% at station D to 33.54% at station A. CV for  $\bar{d}$  ranged from 3.47% at station B to 11.10% at station D. The greater CV at D may have been caused by emergent tree areas. Changes in abundance of the most common species have an effect on  $\bar{d}$  estimates and lowest  $\bar{d}$  occurred within station D. Where variance may vary in proportion to the mean, as with plankton data, logarithmic transformations have a stabilizing influence on the variance (Cassie, 1963). The logarithmic transformation used in the equation to calculate  $\bar{d}$  probably lowers the CV.

By comparing the CV for  $\bar{d}$  to those of concentration it appears that more confidence may be placed on the  $\bar{d}$  estimates derived from traditional sampling programs than from total concentration estimates. However,

TABLE IV

NUMBER OF SPECIES ( $s$ ), TOTAL CONCENTRATION IN NUMBER LITER<sup>-1</sup>  
 ( $c$ ), AND  $\bar{d}$  ( $N = 400$ ) AND STANDARD DEVIATIONS  
 OF PELAGIC NET ZOOPLANKTON IN MAIN POOL  
 OF EUFAULA LAKE, AUGUST, 1969

Station A				Station B			
Site	s	c	$\bar{d}$	Site	s	c	$\bar{d}$
1,1	11.0	28.6	2.53	1,1	10.6	24.0	2.47
1,2	10.0	35.1	2.31	1,2	11.3	21.5	2.55
1,3	9.7	22.0	2.41	1,3	11.0	15.3	2.51
2,1	11.0	20.4	2.48	2,1	10.3	25.5	2.60
2,2	11.0	24.8	2.68	2,2	10.3	20.1	2.48
2,3	11.3	28.9	2.58	2,3	10.0	17.2	2.57
3,1	10.3	29.1	2.54	3,1	10.6	18.1	2.57
3,2	10.3	36.0	2.53	3,2	11.3	18.2	2.51
3,3	11.3	24.0	2.43	3,3	10.6	17.3	2.66
$\bar{X}$	10.6	27.6	2.50	$\bar{X}$	10.6	19.7	2.54
s		9.3	0.10	s		3.0	0.09
Station C				Station D			
Site	s	c	$\bar{d}$	Site	s	c	$\bar{d}$
1,1	10.6	22.1	2.60	1,1	10.7	27.5	2.91
1,2	10.6	24.5	2.45	1,2	11.3	25.3	2.96
1,3	11.7	31.2	2.65	1,3	11.3	33.8	2.71
2,1	12.0	27.1	2.52	2,1	10.7	31.7	2.42
2,2	11.0	37.7	2.47	2,2	10.7	27.1	2.16
2,3	11.3	32.8	2.57	2,3	10.0	29.1	2.47
3,1	11.0	29.8	2.48	3,1	10.0	31.2	2.43
3,2	11.3	33.4	2.50	3,2	10.3	32.3	2.39
3,3	10.6	33.4	2.52	3,3	10.6	27.7	2.49
$\bar{X}$	11.1	30.2	2.52	$\bar{X}$	10.6	29.5	2.54
s		6.9	0.01	s		4.0	0.28

TABLE V  
ANALYSIS OF VARIANCE OF  $\bar{d}$  AND CONCENTRATION (c) ESTIMATES OF PELAGIC NET  
ZOOPLANKTON, MAIN POOL, EUFAULA LAKE, AUGUST, 1969  
( $\alpha = .05$ )

Source of Variation	Degrees of Freedom	$F_{tab}$		$F_{cal}$			
				Station A	Station B	Station C	Station D
Total	26						
Treatments	8	2.51	$\bar{d}$ c	2.52 1.03	1.42 2.60	0.95 1.53	1.79 1.54
Rows	2	3.55	$\bar{d}$ c	4.24 0.72	1.28 1.78	2.00 2.65	3.27 0.36
Columns	2	3.55	$\bar{d}$ c	0.33 1.48	1.33 6.04*	0.92 2.18	2.68 0.69
Rows x Columns	4	2.93	$\bar{d}$ c	2.76 0.95	1.54 1.30	0.48 0.64	0.61 2.54
Error	18	CV	$\bar{d}$ c	4.63% 33.54%	3.47% 18.51%	4.61% 22.77%	11.10% 13.55%
		SE	$\bar{d}$ c	.01 1.78	.02 .71	.01 1.32	.05 .77

\* $F_{tab}$  ( $\alpha = .01$ ) = 6.01.

total concentration estimates seem to be derived adequately from traditional sampling programs.

Concentration of zooplankton estimates indicate fish food potential and lower trophic level abundance and  $\bar{d}$  indicating the stability and relationship to perturbations within the aquatic system. Margalef (1968) relates stability in ecosystems to diversity of species, with an increase in diversity generally accompanying succession toward a more stable, climax community. After analyzing populations on benthic invertebrates it was concluded that  $\bar{d}$  values of less than 1 are usually obtained in areas of heavy pollution, values of 1 to 3 in moderate pollution areas, and values exceeding 3 in clean water areas (Wilhm and Dorris, 1968). Staub, et al. (1970) results with plankton agreed with Wilhm and Dorris except they made an added division where  $\bar{d}$  values between one and two were found in areas of moderate pollution and 2 to 3 in areas of light pollution. Values of  $\bar{d}$  in this study ranged from 2.16 to 2.91 with a mean of 2.53.

## CHAPTER IV

### CONCLUSIONS

The null hypothesis, that  $\bar{d}$  and concentration estimates are not affected by spatial heterogeneity (patchiness) of pelagic net zooplankton was not rejected in three of the four tests. This means that traditional sampling programs where a single site represents a station and the water mass in that area are adequate when estimating  $\bar{d}$  and concentration of pelagic net zooplankton using the sampling procedures used in this study.

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